

# Package ‘PoolBal’

January 20, 2025

**Version** 0.1-0

**Encoding** UTF-8

**Title** Balancing Central and Marginal Rejection of Pooled p-Values

**Description** When using pooled p-values to adjust for multiple testing, there is an inherent balance that must be struck between rejection based on weak evidence spread among many tests and strong evidence in a few, explored in Salahub and Oldford (2023) <[arXiv:2310.16600](https://arxiv.org/abs/2310.16600)>. This package provides functionality to compute marginal and central rejection levels and the centrality quotient for p-value pooling functions and provides implementations of the chi-squared quantile pooled p-value (described in Salahub and Oldford (2023)) and a proposal from Heard and Rubin-Delanchy (2018) <[doi:10.1093/biomet/asx076](https://doi.org/10.1093/biomet/asx076)> to control the quotient's value.

**Author** Chris Salahub [aut, cre]

**Maintainer** Chris Salahub <[chris.salahub@uwaterloo.ca](mailto:chris.salahub@uwaterloo.ca)>

**Depends** R (>= 4.3.0)

**Imports** methods

**License** GPL (>= 3)

**NeedsCompilation** no

**Repository** CRAN

**RoxygenNote** 7.2.3

**Date/Publication** 2023-11-22 10:10:02 UTC

## Contents

altFrequencyMat . . . . .	2
betaDiv . . . . .	3
chiKappa . . . . .	4
chiPc . . . . .	5
chiPool . . . . .	6
chiPr . . . . .	7
chiQ . . . . .	8
convertGeneticSigma . . . . .	9
estimatePc . . . . .	10

estimatePrb . . . . .	11
estimateQ . . . . .	12
findA . . . . .	13
hrPc . . . . .	14
hrPool . . . . .	15
hrPr . . . . .	16
hrQ . . . . .	17
hrStat . . . . .	18
klDiv . . . . .	19
marHistHeatMap . . . . .	20
rBetaH4 . . . . .	21
satterApproxP . . . . .	22
satterChiPool . . . . .	23

<b>Index</b>	<b>24</b>
--------------	-----------

---

altFrequencyMat	<i>Identify a region of plausible alternative hypotheses in the proportion, strength of non-null evidence space</i>
-----------------	---

---

## Description

This function provides a convenient way to interact with simulations performed over a grid of possible alternatives spanning the proportion ( $\eta$ ) and strength (KL divergence) of evidence against the null hypothesis under beta alternatives.

## Usage

```
altFrequencyMat(logKappaRange, logW = FALSE)
```

## Arguments

logKappaRange	pair of numeric values
logW	logical, should the log scale simulation be used?

## Details

The simulation this function summarized used a range of  $\eta$ ,  $w$ , and KL divergence values to generate thousands of potential alternative distributions. The power of each chi-squared pooled p-value for 161 kappa values ranging from  $\exp(-8)$  to  $\exp(8)$  selected uniformly on the log scale was then computed for each alternative using 10,000 simulated examples. Every choice of kappa was compared to the maximum power across all kappas for each setting using a binomial test of differences. This same simulation was repeated twice: once for  $w$  values selected uniformly from 0 to 1 and another where selection was uniform on the log scale. The internal data summarizes the results by reporting the count of instances in  $w$  (or  $\log w$ ) where a given kappa value was most powerful for a given  $\eta$  and KL divergence.

Though the simulation data is not exported to users and so cannot be accessed directly, this function allows a user to query the data with a range of kappa values (corresponding to those where a given

sample seems most powerful) and returns the count of cases in  $w$  where a kappa in the corresponding kappa range was most powerful given the eta, KL-divergence combination with beta alternatives. The simulations only spanned kappa values from  $\exp(-8)$  to  $\exp(8)$ , so providing values outside this range will give very inaccurate results.

**Value**

An 81 by 81 matrix giving summarized counts of cases.

**Author(s)**

Chris Salahub

**Examples**

```
altFrequencyMat(c(-1, 1), logW = FALSE)
altFrequencyMat(c(-1, 1), logW = TRUE)
```

---

betaDiv	<i>Compute the Kullback-Leibler divergence between the beta and uniform distributions</i>
---------	---

---

**Description**

Computes the Kullback-Leibler divergence for the special case of the uniform density against the beta density.

**Usage**

```
betaDiv(a, w = (1 - a)/(b - a), b = 1/w + a * (1 - 1/w))
```

**Arguments**

a	first shape parameter between 0 and infinity
w	UMP parameter between 0 and 1
b	second shape parameter between 0 and infinity

**Details**

This function accepts either the  $a/b$  parameterization (equivalent to shape1/shape2 in R), or the  $a/w$  parameterization which links the divergence to the UMP test.

**Value**

A real value.

**Author(s)**

Chris Salahub

**Examples**

```
betaDiv(a = 0.5, w = 0.5)
betaDiv(a = 0.1, b = 1)
```

---

chiKappa

*Chi-squared kappa for a given centrality quotient*


---

**Description**

Computes the kappa (degrees of freedom) required to obtain a given centrality quotient using the chi-square pooled p-value.

**Usage**

```
chiKappa(
  cq,
  M,
  alpha = 0.05,
  interval = c(0, 100),
  tol = .Machine$double.eps^0.5
)
```

**Arguments**

cq	numeric between 0 and 1
M	integer sample size greater than 0
alpha	numeric between 0 and 1
interval	numeric of length 2, where should roots be sought?
tol	numeric, how close do values need to be for equality?

**Details**

This function is essentially a wrapper for uniroot which finds where chiCentQuot gives an output equal to the given centrality quotient to provide an approximate kappa giving that quotient.

**Value**

A numeric within interval.

**Author(s)**

Chris Salahub

**Examples**

```
chiKappa(0.5, 10, 0.05)
chiKappa(0.5, 20, 0.05)
chiKappa(0.5, 100, 0.05, interval = c(0, 10))
```

---

chiPc	<i>Chi-squared central rejection level</i>
-------	--

---

**Description**

Computes the central rejection level for the chi-squared pooled p-value.

**Usage**

```
chiPc(kappa, M, alpha = 0.05)
```

**Arguments**

kappa	numeric between 0 and infinity
M	integer sample size greater than 0
alpha	numeric between 0 and 1

**Details**

The central rejection level is the maximum p-value shared among all tests which still results in rejection of the null using a pooled p-value. For the chi-squared pooled p-value, this is an upper tail probability of the chi-squared distribution. This function computes the upper tail probability for a given sample size M, degrees of freedom kappa, and rejection level alpha.

**Value**

A numeric between 0 and 1.

**Author(s)**

Chris Salahub

**Examples**

```
chiPc(2, 10, 0.05)  
chiPc(2, 20, 0.05) # increases in sample size
```

---

`chiPool`*Chi-squared p-value pooling*

---

**Description**

This implements the chi-squared pooled p-value which can be used to control the centrality quotient when pooling p-values.

**Usage**

```
chiPool(p, kappa)
```

**Arguments**

<code>p</code>	numeric vector of p-values between 0 and 1
<code>kappa</code>	numeric value between 0 and infinity

**Details**

The chi-squared pooled p-value is a quantile transformation pooled p-value based on the chi-squared distribution with degrees of freedom  $kappa$ . By setting  $kappa$  between 0 and infinity, smooth interpolation is achieved between Tippett's minimum pooled p-value and Stouffer's normal quantile pooled p-value respectively. Choosing a  $kappa$  value of 2, Fisher's pooling function is obtained. Tippett's pooled p-value is maximally non-central and Stouffer's is maximally central, while Fisher's presents a balance between marginal and central rejection.

**Value**

A pooled p-value between 0 and 1.

**Author(s)**

Chris Salahub

**Examples**

```
p <- c(0.1, 0.5, 0.9)
chiPool(p, exp(-4))
chiPool(p, 2)
chiPool(p, exp(4))
```

---

chiPr	<i>Chi-squared marginal rejection level</i>
-------	---

---

**Description**

Computes the marginal rejection level for the chi-squared pooled p-value.

**Usage**

```
chiPr(kappa, M, alpha = 0.05)
```

**Arguments**

kappa	numeric between 0 and infinity
M	integer sample size greater than 0
alpha	numeric between 0 and 1

**Details**

The marginal rejection level is the maximum p-value in a single test which results in rejection when all other tests produce p-values of one. For the chi-squared pooled p-value, this is an upper tail probability of the chi-squared distribution. This function computes the upper tail probability for a given sample size M, degrees of freedom kappa, and rejection level alpha.

**Value**

A numeric between 0 and 1.

**Author(s)**

Chris Salahub

**Examples**

```
chiPr(2, 10, 0.05)  
chiPr(2, 20, 0.05)
```

---

chiQ	<i>Chi-squared centrality quotient</i>
------	--

---

**Description**

Computes the centrality quotient of the chi-square pooled p-value.

**Usage**

```
chiQ(kappa, M, alpha = 0.05)
```

**Arguments**

kappa	numeric between 0 and infinity
M	integer sample size greater than 0
alpha	numeric between 0 and 1

**Details**

The centrality quotient of a pooled p-value measures the relative preference it gives to p-values all sharing the same level of evidence over a single test with strong evidence relative to others. For the chi-square pooled p-value, this is a conditional probability which this function computes.

**Value**

A numeric between 0 and 1.

**Author(s)**

Chris Salahub

**Examples**

```
chiQ(2, 10, 0.05)
chiQ(2, 20, 0.05)
chiQ(0.5, 20, 0.05)
```



---

convertGeneticSigma     *Convert p-value correlation to chi-squared covariance*

---

### Description

Convert a matrix of correlations between p-values to a matrix of covariances between their chi-squared transforms.

### Usage

```
convertGeneticSigma(sigma, kappa, models = chiCorMods)
```

### Arguments

sigma	M by M correlation matrix between markers
kappa	numeric degrees of freedom
models	model object with a predict method

### Details

This function uses models fit to large simulated data sets to convert a matrix of correlations between genetic markers the covariance matrix of chi-squared random variables gained from transforming p-values on these markers. The simulations used to create data for these models assume the p-values for each marker arise from tests of association with a common, normally distributed trait independent of all markers. As a result, this conversion function should be used only in analogous settings.

Models were fit for degrees of freedom at increments of 0.1 between -8 and 8 on the log scale, and interpolation is applied if the degrees of freedom given to the function does not fall exactly on this grid (with a warning provided to the user).

If a user wants to generalize this setting, the option to provide a custom list of models which predict based on a named argument 'zcor' is supported. Each model must have a name in the list that can be converted to a numeric, and these are assumed to be on the natural log scale.

### Value

M by M matrix of chi-squared covariances

### Author(s)

Chirs Salahub

---

`estimatePc`*Compute the central rejection level*

---

**Description**

Estimates the central rejection level for an arbitrary pooled p-value function.

**Usage**

```
estimatePc(  
  poolFun,  
  alpha = 0.05,  
  M = 2,  
  interval = c(0, 1),  
  poolArgs = list(),  
  ...  
)
```

**Arguments**

<code>poolFun</code>	function accepting a vector of p-values
<code>alpha</code>	numeric between 0 and 1
<code>M</code>	integer, how many p-values are there?
<code>interval</code>	two numerics giving the bounds of root-searching
<code>poolArgs</code>	(optional) additional named arguments for <code>poolFun</code>
<code>...</code>	additional arguments to <code>uniroot</code>

**Details**

The central rejection level is the maximum p-value shared among all tests which still results in rejection of the null using a pooled p-value.

This function is essentially a wrapper for `uniroot`, and accepts a pooling function which takes a numeric vector as its first argument and potentially other arguments given in `poolArgs` and returns a single value. Using this pooling function, a specified dimension `M` and a rejection level `alpha`, `uniroot` searches for the root to `poolFun - alpha` along the line where all p-values are equal.

**Value**

The `uniroot` output.

**Author(s)**

Chris Salahub

**Examples**

```
tippool <- function(p) 1 - (1 - min(p))^(length(p))
estimatePc(tippool, 0.05, M = 10, interval = c(0, 1))
```

---

estimatePrb	<i>Compute the marginal rejection level</i>
-------------	---

---

**Description**

Estimates the marginal rejection level for an arbitrary pooled p-value function.

**Usage**

```
estimatePrb(
  poolFun,
  alpha = 0.05,
  b = 1,
  M = 2,
  interval = c(0, b),
  poolArgs = list(),
  ...
)
```

**Arguments**

poolFun	function accepting a vector of p-values
alpha	numeric between 0 and 1
b	numeric, the value of the M - 1 repeated p-values
M	integer, how many p-values are there?
interval	two numerics giving the bounds of root-searching
poolArgs	(optional) additional named arguments for poolFun
...	additional arguments to uniroot

**Details**

The marginal rejection level is the maximum p-value in a single test less than b which still results in rejection of the null when all other tests have a p-value of b.

This function is essentially a wrapper for uniroot, and accepts a pooling function which takes a numeric vector as its first argument and potentially other arguments given in poolArgs and returns a single value. Using this pooling function, a specified dimension M and a rejection level alpha, uniroot searches for the root to poolFun - alpha along one margin when all other p-values are equal to b.

**Value**

The uniroot output.

**Author(s)**

Chris Salahub

**Examples**

```
stopool <- function(p) pnorm(sum(qnorm(p, lower.tail = FALSE))/ sqrt(length(p)), lower.tail = FALSE)
estimatePrb(stopool, 0.05, M = 10, interval = c(.Machine$double.eps, 1))
estimatePrb(stopool, 0.05, M = 10, b = 0.5, interval = c(.Machine$double.eps, 1))
```

---

 estimateQ

---

*Compute the centrality quotient*


---

**Description**

Estimates the centrality quotient for an arbitrary pooled p-value function.

**Usage**

```
estimateQ(
  poolFun,
  alpha = 0.05,
  M = 2,
  interval = c(0, 1),
  poolArgs = list(),
  ...
)
```

**Arguments**

poolFun	function accepting a vector of p-values
alpha	numeric between 0 and 1
M	integer, how many p-values are there?
interval	two numerics giving the bounds of root-searching
poolArgs	(optional) additional named arguments for poolFun
...	additional arguments to uniroot

**Details**

The centrality quotient communicates the tendency for a test to favour evidence shared among all tests over strong evidence in a single test.

This function uses the individual estimation functions for central and marginal rejection levels to compute the centrality quotient for an arbitrary pooled p-value function. The option to specify b for marginal rejection is included in case the pooled p-value has strange behaviour when p-values are equal to 1.

**Value**

The uniroot output.

**Author(s)**

Chris Salahub

**Examples**

```
estimateQ(chiPool, alpha = 0.05, M = 10, poolArgs = list(kappa = 10))
```

---

findA

*Estimate parameter for a given beta KL divergence and UMP test*

---

**Description**

Computes the first parameter value for a given KL divergence and UMP test.

**Usage**

```
findA(w, logd = 0, ...)
```

**Arguments**

w	UMP parameter between 0 and 1
logd	numeric value, the log KL divergence
...	additional arguments to uniroot

**Details**

This function uses uniroot to invert the beta divergence for a given w and return the a value which gives that beta divergence given the UMP parameter w. The search interval is specified internally, so should not be passed in using additional argument.

**Value**

A real value.

**Author(s)**

Chris Salahub

**Examples**

```
findA(0.5, logd = 0)
```

---

`hrPc`*Empirical UMP beta central rejection level*

---

**Description**

Uses simulation to estimate the central rejection level for the UMP pooled p-value of a restricted beta family

**Usage**

```
hrPc(w, alpha = 0.05, M = 2, nsim = 1e+05)
```

**Arguments**

<code>w</code>	numeric between 0 and 1
<code>alpha</code>	numeric between 0 and 1
<code>M</code>	integer sample size greater than 0
<code>nsim</code>	integer, the number of simulated null cases generated

**Details**

The central rejection level is the maximum p-value shared among all tests which still results in rejection of the null using a pooled p-value.

To test the null hypotheses that all p-values are uniform against a restricted beta family  $0 < a \leq 1 \leq b$ , the most powerful pooled p-value linearly combines upper and lower tail probabilities of the chi-squared distribution with two degrees of freedom with weights  $w$  and  $(1 - w)$  where  $w = (1 - a)/(b - a)$ .

This function estimates the central rejection level empirically by simulating a specified number of null cases to give an empirical pooled p-value for the rejection level  $\alpha$ .

**Value**

A numeric between 0 and 1.

**Author(s)**

Chris Salahub

**Examples**

```
hrPc(w = 0.5, alpha = 0.05, M = 10)
hrPc(w = 0.5, alpha = 0.05, M = 20)
```

---

hrPool	<i>Empirical UMP beta pooled p-value</i>
--------	--

---

**Description**

Uses simulation under the null to approximate the UMP pooled p-value for a restricted beta family.

**Usage**

```
hrPool(w = 1, M = 10, nsim = 1e+05)
```

**Arguments**

w	numeric value between 0 and 1
M	integer, the number of tests to pool
nsim	integer, the number of simulated null cases generated

**Details**

To test the null hypotheses that all p-values are uniform against a restricted beta family  $0 < a \leq 1 \leq b$ , the most powerful pooled p-value linearly combines upper and lower tail probabilities of the chi-squared distribution with two degrees of freedom with weights  $w$  and  $(1 - w)$  where  $w = (1 - a)/(b - a)$ .

This function computes the statistic given by this combination for a collection of p-values, and then simulates a specified number of null cases to give an empirical pooled p-value. It produces a closure so that the time-intensive simulation step doesn't need to be repeated.

**Value**

A closure which accepts a vector of values between 0 and 1 and returns a single numeric between 0 and 1

**Author(s)**

Chris Salahub

**Examples**

```
p <- c(0.1, 0.5, 0.9)
hr2 <- hrPool(w = 0.2, M = 3)
hr2(p)
hr5 <- hrPool(w = 0.5, M = 3, nsim = 100)
hr5(p)
```

---

`hrPr`*Empirical UMP beta marginal rejection level*

---

**Description**

Uses simulation to estimate the marginal rejection level for the UMP pooled p-value of a restricted beta family

**Usage**

```
hrPr(w, alpha = 0.05, M = 2, nsim = 1e+05)
```

**Arguments**

<code>w</code>	numeric between 0 and 1
<code>alpha</code>	numeric between 0 and 1
<code>M</code>	integer sample size greater than 0
<code>nsim</code>	integer, the number of simulated null cases generated

**Details**

The marginal rejection level is the maximum p-value in a single tests which still results in rejection of the null when all other tests have a p-value of 1.

To test the null hypotheses that all p-values are uniform against a restricted beta family  $0 < a \leq 1 \leq b$ , the most powerful pooled p-value linearly combines upper and lower tail probabilities of the chi-squared distribution with two degrees of freedom with weights  $w$  and  $(1 - w)$  where  $w = (1 - a)/(b - a)$ .

This function estimates the marginal rejection level empirically by simulating a specified number of null cases to give an empirical pooled p-value for the rejection level  $\alpha$ .

**Value**

A numeric between 0 and 1.

**Author(s)**

Chris Salahub

**Examples**

```
hrPr(w = 0.5, alpha = 0.05, M = 10)
hrPr(w = 0.5, alpha = 0.05, M = 10) # decreases in sample size
```



---

hrQ                                      *Empirical UMP beta centrality quotient*

---

**Description**

Estimates the centrality quotient for the UMP pooled p-value of a restricted beta family.

**Usage**

```
hrQ(w, alpha = 0.05, M = 2, nsim = 1e+05)
```

**Arguments**

w	numeric between 0 and 1
alpha	numeric between 0 and 1
M	integer sample size greater than 0
nsim	integer, the number of simulated null cases generated

**Details**

The centrality quotient communicates the tendency for a test to favour evidence shared among all tests over strong evidence in a single test.

To test the null hypotheses that all p-values are uniform against a restricted beta family  $0 < a \leq 1 \leq b$ , the most powerful pooled p-value linearly combines upper and lower tail probabilities of the chi-squared distribution with two degrees of freedom with weights  $w$  and  $(1 - w)$  where  $w = (1 - a)/(b - a)$ .

This function uses the individual estimation functions for central and marginal rejection levels to compute the centrality quotient for the UMP pooled p-value.

**Value**

An empirical estimate of the centrality quotient.

**Author(s)**

Chris Salahub

**Examples**

```
hrQ(0.8, alpha = 0.05, M = 10)
```

---

hrStat	<i>UMP beta p-value pooled statistic</i>
--------	--

---

**Description**

Computes the UMP p-value pooling statistic for a restricted beta family.

**Usage**

```
hrStat(p, w = 1)
```

**Arguments**

p	numeric vector of p-values between 0 and 1
w	numeric value between 0 and 1

**Details**

To test the null hypotheses that all p-values are uniform against a restricted beta family  $0 < a \leq 1 \leq b$ , the most powerful pooled p-value linearly combines upper and lower tail probabilities of the chi-squared distribution with two degrees of freedom with weights  $w$  and  $(1 - w)$  where  $w = (1 - a)/(b - a)$ .

This function computes the statistic given by this combination for a collection of p-values, simulation or approximation is required to convert this to a p-value.

**Value**

A numeric value giving the pooled statistic.

**Author(s)**

Chris Salahub

**Examples**

```
p <- c(0.1, 0.5, 0.9)
hrStat(p, 0.2)
hrStat(p, 0.5)
hrStat(p, 0.9)
```

---

kIDiv *Compute the Kullback-Leibler divergence*

---

**Description**

Computes the Kullback-Leibler divergence for two arbitrary densities f1 and f2.

**Usage**

```
kIDiv(f1, f2, lower = 0, upper = 1)
```

**Arguments**

f1	density function of a real-valued random variable
f2	density function of a real-values random variable
lower	real value, the lower bound of integration
upper	real value, the upper bound of integration

**Details**

Given lower and upper bounds, this function integrates the expression for the Kullback-Leibler divergence  $KL(f1|f2)$ .

**Value**

A real value.

**Author(s)**

Chris Salahub

**Examples**

```
kIDiv(dunif, function(x) dbeta(x, 0.5, 1))
```

---

marHistHeatMap      *Heatmap with marginal histograms*

---

### Description

Display a matrix using a heatmap with marginal histograms.

### Usage

```
marHistHeatMap(  
  mat,  
  main = "",  
  ylab = expression(eta),  
  xlab = "lnD(a,w)",  
  pal = NULL,  
  histFill = adjustcolor("firebrick", 0.5),  
  ...  
)
```

### Arguments

mat	numeric matrix to be plotted
main	title
ylab	y axis label
xlab	x axis label
pal	palette for heatmap
histFill	colour to fill histogram bars
...	additional arguments to image

### Details

This function accepts a matrix of values and plots the matrix with saturation/hue determined by a provided palette argument generated by `colorRampPalette`, for example. Marginal histograms summarizing the relative frequencies along both dimensions are also plotted to give a complete sense of the individual distributions alongside their joint distribution. This was designed to summarize the alternative distribution space summarized by `altFrequencyMat`, and the defaults reflect this.

### Value

Plot the data using a heatmap and marginal histograms and return nothing.

### Author(s)

Chris Salahub

**Examples**

```
marHistHeatMap(altFrequencyMat(c(0, 2)))
```

---

rBetaH4

*Generate realizations of beta alternative distributions*


---

**Description**

These functions can be used to generate samples of p-values all following a beta distribution (H4) or following either uniform or beta distributions according to proportion eta (H3).

**Usage**

```
rBetaH4(a, b = 1/w + a * (1 - 1/w), w = (1 - a)/(b - a), M = 2, N = 10)
```

```
rBetaH3(
  a,
  b = 1/w + a * (1 - 1/w),
  w = (1 - a)/(b - a),
  eta = 0.5,
  M = 2,
  N = 10
)
```

**Arguments**

a	first beta parameter, numeric between 0 and infinity
b	second beta parameter, numeric between 0 and infinity
w	UMP parameter between 0 and 1
M	number of p-values per realization
N	number of realizations
eta	numeric between 0 and 1, proportion of non-null tests per sample

**Details****Alternatives**

These functions are provided as a convenience, and support a/b (shape1/shape2) or a/w specification of beta parameters.

**Value**

An N by M matrix of simulated p-values.

**Functions**

- rBetaH4(): iid Beta(a,w) p-values
- rBetaH3(): M\*eta iid Beta(a,w) p-values, others uniform

**Author(s)**

Chris Salahub

**Examples**

```
rBetaH4(a = 0.5, b = 1.5, M = 10, N = 100)
rBetaH3(a = 0.5, b = 1.5, eta = 0.5, M = 10, N = 100)
```

---

satterApproxP

*Satterthwaite p-values*

---

**Description**

p-value of the sum of dependent chi-squared using the Satterthwaite approximation for the degrees of freedom.

**Usage**

```
satterApproxP(qs, covmat, kappa)
```

**Arguments**

qs	M numeric values (observed chi-squared values)
covmat	M by M covariance matrix of qs
kappa	degrees of freedom of qs

**Details**

Computes the p-value of an observed vector of chi-squared variables using the Satterthwaite approximation. This approximates the sum of dependent chi-squared variables with a scaled chi-squared distribution with degrees of freedom chosen to match the first two moments of the dependent sum.

**Value**

a numeric in [0,1], the p-value of the sum

**Author(s)**

Chris Salahub

---

`satterChiPool`*Pool p-values using the Satterthwaite approximation*

---

**Description**

Compute the pooled p-value of dependent p-values based on the dependence present when they are all converted to chi-squared random variables by the same chi-squared quantile function.

**Usage**

```
satterChiPool(ps, covmat, kappa)
```

**Arguments**

<code>ps</code>	numeric vector of M p-values
<code>covmat</code>	M by M covariance matrix of chi-squared random variables arising from quantile transformations of ps
<code>kappa</code>	numeric degrees of freedom

**Details**

Care must be taken in the arguments for this function, as the `covmat` argument accepts the covariance of the transformed variables rather than the covariance of the p-values, and so passes the argument `covmat` directly to the function that computes the Satterthwaite approximation. For the case of genetic markers, the `convertGeneticSigma` function provides the appropriate matrix given a genetic correlation matrix.

**Value**

A pooled p-value between 0 and 1.

**Author(s)**

Chris Salahub

# Index

altFrequencyMat, 2

betaDiv, 3

chiKappa, 4

chiPc, 5

chiPool, 6

chiPr, 7

chiQ, 8

convertGeneticSigma, 9

estimatePc, 10

estimatePrb, 11

estimateQ, 12

findA, 13

hrPc, 14

hrPool, 15

hrPr, 16

hrQ, 17

hrStat, 18

kIDiv, 19

marHistHeatMap, 20

rBetaH3 (rBetaH4), 21

rBetaH4, 21

satterApproxP, 22

satterChiPool, 23