

# Package ‘ReIns’

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**Title** Functions from ``Reinsurance: Actuarial and Statistical Aspects''

**Description** Functions from the book ``Reinsurance: Actuarial and Statistical Aspects'' (2017) by Hansjoerg Albrecher, Jan Beirlant and Jef Teugels <<https://www.wiley.com/en-us/Reinsurance%3A+Actuarial+and+Statistical+Aspects-p-9780470772683>>.

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Burr

*The Burr distribution*

### Description

Density, distribution function, quantile function and random generation for the Burr distribution (type XII).

### Usage

```
dburr(x, alpha, rho, eta = 1, log = FALSE)
pburr(x, alpha, rho, eta = 1, lower.tail = TRUE, log.p = FALSE)
qburrr(p, alpha, rho, eta = 1, lower.tail = TRUE, log.p = FALSE)
rburr(n, alpha, rho, eta = 1)
```

### Arguments

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
alpha	The $\alpha$ parameter of the Burr distribution, a strictly positive number.
rho	The $\rho$ parameter of the Burr distribution, a strictly negative number.

<code>eta</code>	The $\eta$ parameter of the Burr distribution, a strictly positive number. The default value is 1.
<code>log</code>	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
<code>lower.tail</code>	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
<code>log.p</code>	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

### Details

The Cumulative Distribution Function (CDF) of the Burr distribution is equal to  $F(x) = 1 - ((\eta + x^{-\rho \times \alpha})/\eta)^{1/\rho}$  for all  $x \geq 0$  and  $F(x) = 0$  otherwise. We need that  $\alpha > 0$ ,  $\rho < 0$  and  $\eta > 0$ .

Beirlant et al. (2004) uses parameters  $\eta, \tau, \lambda$  which correspond to  $\eta, \tau = -\rho \times \alpha$  and  $\lambda = -1/\rho$ .

### Value

`dburr` gives the density function evaluated in  $x$ , `pburr` the CDF evaluated in  $x$  and `qburr` the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

`rburr` returns a random sample of length  $n$ .

### Author(s)

Tom Reynkens.

### References

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

### See Also

[tBurr, Distributions](#)

### Examples

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dburr(x, alpha=2, rho=-1), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, pburr(x, alpha=2, rho=-1), xlab="x", ylab="CDF", type="l")
```

cEPD

*EPD estimator for right censored data***Description**

Computes the EPD estimates adapted for right censored data.

**Usage**

```
cEPD(data, censored, rho = -1, beta = NULL, logk = FALSE,
      plot = FALSE, add = FALSE, main = "EPD estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
rho	A parameter for the $\rho$ -estimator of Fraga Alves et al. (2003) when strictly positive or choice(s) for $\rho$ if negative. Default is -1.
beta	Parameter for EPD ( $\beta = -\rho/\gamma$ ). If NULL (default), beta is estimated by $-\rho/H_{k,n}$ with $H_{k,n}$ the Hill estimator.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "EPD estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The function EPD uses  $\tau$  which is equal to  $-\beta$ .

This estimator is only suitable for right censored data.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma1	Vector of the corresponding estimates for the $\gamma$ parameter of the EPD.
kappa1	Vector of the corresponding MLE estimates for the $\kappa$ parameter of the EPD.
beta	Vector of estimates for (or values of) the $\beta$ parameter of the EPD.
Delta	Difference between gamma1 and the Hill estimator for censored data.

**Author(s)**

Tom Reynkens based on R code from Anastasios Bardoutsos.

**References**

Beirlant, J., Bardoutsos, A., de Wet, T. and Gijbels, I. (2016). "Bias Reduced Tail Estimation for Censored Pareto Type Distributions." *Statistics & Probability Letters*, 109, 78–88.

Fraga Alves, M.I. , Gomes, M.I. and de Haan, L. (2003). "A New Class of Semi-parametric Estimators of the Second Order Parameter." *Portugaliae Mathematica*, 60, 193–214.

**See Also**

[EPD](#), [cProbEPD](#), [cGPDmle](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# EPD estimator adapted for right censoring
cepd <- cEPD(Z, censored=censored, plot=TRUE)
```

---

cExpQQ

*Exponential quantile plot for right censored data*

---

**Description**

Exponential QQ-plot adapted for right censored data.

**Usage**

```
cExpQQ(data, censored, plot = TRUE, main = "Exponential QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
plot	Logical indicating if the quantiles should be plotted in an exponential QQ-plot, default is TRUE.
main	Title for the plot, default is "Exponential QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The exponential QQ-plot adapted for right censoring is given by

$$(-\log(1 - F_{km}(Z_{j,n})), Z_{j,n})$$

for  $j = 1, \dots, n - 1$ , with  $Z_{i,n}$  the  $i$ -th order statistic of the data and  $F_{km}$  the Kaplan-Meier estimator for the CDF. Hence, it has the same empirical quantiles as an ordinary exponential QQ-plot but replaces the theoretical quantiles  $-\log(1 - j/(n + 1))$  by  $-\log(1 - F_{km}(Z_{j,n}))$ .

This QQ-plot is only suitable for right censored data.

In Beirlant et al. (2007), only a Pareto QQ-plot adapted for right-censored data is proposed. This QQ-plot is constructed using the same ideas, but is not described in the paper.

**Value**

A list with following components:

eqq.the	Vector of the theoretical quantiles, see Details.
eqq.emp	Vector of the empirical quantiles from the data.

**Author(s)**

Tom Reynkens

**References**

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

**See Also**

[ExpQQ](#), [cLognormalQQ](#), [cParetoQQ](#), [cWeibullQQ](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)
```



```

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Exponential QQ-plot adapted for right censoring
cExpQQ(Z, censored=censored)

```

---

cgenHill

*Generalised Hill estimator for right censored data*


---

## Description

Computes the generalised Hill estimates adapted for right censored data.

## Usage

```

cgenHill(data, censored, logk = FALSE, plot = FALSE, add = FALSE,
         main = "Generalised Hill estimates of the EVI", ...)

```

## Arguments

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma_1$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma_1$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Generalised Hill estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

## Details

The generalised Hill estimator adapted for right censored data is equal to the ordinary generalised Hill estimator divided by the proportion of the  $k$  largest observations that is non-censored.

This estimator is only suitable for right censored data.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma1	Vector of the corresponding generalised Hill estimates.

**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[genHill](#), [cHill](#), [cProbGH](#), [cQuantGH](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Generalised Hill estimator adapted for right censoring
cghill <- cgenHill(Z, censored=censored, plot=TRUE)
```

---

cGPDmle

*GPD-ML estimator for right censored data*

---

**Description**

Computes ML estimates of fitting GPD to peaks over a threshold adapted for right censoring.

**Usage**

```
cGPDmle(data, censored, start = c(0.1,1), warnings = FALSE, logk = FALSE,
         plot = FALSE, add = FALSE, main = "POT estimates of the EVI", ...)
```

```
cPOT(data, censored, start = c(0.1,1), warnings = FALSE, logk = FALSE,
      plot = FALSE, add = FALSE, main = "POT estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
start	Vector of length 2 containing the starting values for the optimisation. The first element is the starting value for the estimator of $\gamma_1$ and the second element is the starting value for the estimator of $\sigma_1$ . Default is <code>c(0.1, 1)</code> .
warnings	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma_1$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma_1$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "POT estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The GPD-MLE estimator for the EVI adapted for right censored data is equal to the ordinary GPD-MLE estimator for the EVI divided by the proportion of the  $k$  largest observations that is non-censored. The estimates for  $\sigma$  are the ordinary GPD-MLE estimates for  $\sigma$ .

This estimator is only suitable for right censored data.

cPOT is the same function but with a different name for compatibility with [POT](#).

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma1	Vector of the corresponding MLE estimates for the $\gamma_1$ parameter of the GPD.
sigma1	Vector of the corresponding MLE estimates for the $\sigma_1$ parameter of the GPD.

**Author(s)**

Tom Reynkens

## References

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

## See Also

[GPDmle](#), [cProbGPD](#), [cQuantGPD](#), [cEPD](#)

## Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# GPD-ML estimator adapted for right censoring
cpot <- cGPDmle(Z, censored=censored, plot=TRUE)
```

---

cHill

*Hill estimator for right censored data*

---

## Description

Computes the Hill estimator for positive extreme value indices, adapted for right censoring, as a function of the tail parameter  $k$  (Beirlant et al., 2007). Optionally, these estimates are plotted as a function of  $k$ .

## Usage

```
cHill(data, censored, logk = FALSE, plot = FALSE, add = FALSE,
      main = "Hill estimates of the EVI", ...)
```

## Arguments

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.

plot	Logical indicating if the estimates of $\gamma_1$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma_1$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Hill estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The Hill estimator adapted for right censored data is equal to the ordinary Hill estimator  $H_{k,n}$  divided by the proportion of the  $k$  largest observations that is non-censored.

This estimator is only suitable for right censored data, use [icHill](#) for interval censored data.

See Section 4.3.2 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma1	Vector of the corresponding Hill estimates.

### Author(s)

Tom Reynkens

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

### See Also

[Hill](#), [icHill](#), [cParetoQQ](#), [cProb](#), [cQuant](#)

### Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)
```

```
# Censoring indicator
censored <- (X>Y)

# Hill estimator adapted for right censoring
chill <- cHill(Z, censored=censored, plot=TRUE)
```

---

cLognormalQQ

*Log-normal quantile plot for right censored data*


---

### Description

Log-normal QQ-plot adapted for right censored data.

### Usage

```
cLognormalQQ(data, censored, plot = TRUE, main = "Log-normal QQ-plot", ...)
```

### Arguments

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
plot	Logical indicating if the quantiles should be plotted in a log-normal QQ-plot, default is TRUE.
main	Title for the plot, default is "Log-normal QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The log-normal QQ-plot adapted for right censoring is given by

$$(\Phi^{-1}(F_{km}(Z_{j,n})), \log(Z_{j,n}))$$

for  $j = 1, \dots, n - 1$ , with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\Phi^{-1}$  the quantile function of the standard normal distribution and  $F_{km}$  the Kaplan-Meier estimator for the CDF. Hence, it has the same empirical quantiles as an ordinary log-normal QQ-plot but replaces the theoretical quantiles  $\Phi^{-1}(j/(n + 1))$  by  $\Phi^{-1}(F_{km}(Z_{j,n}))$ .

This QQ-plot is only suitable for right censored data.

In Beirlant et al. (2007), only a Pareto QQ-plot adapted for right-censored data is proposed. This QQ-plot is constructed using the same ideas, but is not described in the paper.

### Value

A list with following components:

lqq.the	Vector of the theoretical quantiles, see Details.
lqq.emp	Vector of the empirical quantiles from the log-transformed data.

**Author(s)**

Tom Reynkens

**References**

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

**See Also**

[LognormalQQ](#), [cExpQQ](#), [cParetoQQ](#), [cWeibullQQ](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Log-normal QQ-plot adapted for right censoring
cLognormalQQ(Z, censored=censored)
```

---

cMoment

*MOM estimator for right censored data*

---

**Description**

Computes the Method of Moment estimates adapted for right censored data.

**Usage**

```
cMoment(data, censored, logk = FALSE, plot = FALSE, add = FALSE,
        main = "Moment estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma_1$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma_1$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Moment estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The moment estimator adapted for right censored data is equal to the ordinary moment estimator divided by the proportion of the  $k$  largest observations that is non-censored.

This estimator is only suitable for right censored data.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma1	Vector of the corresponding moment estimates.

**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[Moment](#), [cProbMOM](#), [cQuantMOM](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)
```



```

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Moment estimator adapted for right censoring
cmom <- cMoment(Z, censored=censored, plot=TRUE)

```

---

cParetoQQ

*Pareto quantile plot for right censored data*


---

### Description

Pareto QQ-plot adapted for right censored data.

### Usage

```
cParetoQQ(data, censored, plot = TRUE, main = "Pareto QQ-plot", ...)
```

### Arguments

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
plot	Logical indicating if the quantiles should be plotted in a Pareto QQ-plot, default is TRUE.
main	Title for the plot, default is "Pareto QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The Pareto QQ-plot adapted for right censoring is given by

$$(-\log(1 - F_{km}(Z_{j,n})), \log Z_{j,n})$$

for  $j = 1, \dots, n - 1$ , with  $Z_{i,n}$  the  $i$ -th order statistic of the data and  $F_{km}$  the Kaplan-Meier estimator for the CDF. Hence, it has the same empirical quantiles as an ordinary Pareto QQ-plot but replaces the theoretical quantiles  $-\log(1 - j/(n + 1))$  by  $-\log(1 - F_{km}(Z_{j,n}))$ .

This QQ-plot is only suitable for right censored data, use [icParetoQQ](#) for interval censored data.

### Value

A list with following components:

pqq.the	Vector of the theoretical quantiles, see Details.
pqq.emp	Vector of the empirical quantiles from the log-transformed data.

**Author(s)**

Tom Reynkens

**References**

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

**See Also**

[ParetoQQ](#), [icParetoQQ](#), [cExpQQ](#), [cLognormalQQ](#), [cWeibullQQ](#), [cHill](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Pareto QQ-plot adapted for right censoring
cParetoQQ(Z, censored=censored)
```

---

cProb

*Estimator of small exceedance probabilities and large return periods using censored Hill*

---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the estimates for the EVI obtained from the Hill estimator adapted for right censoring.

**Usage**

```
cProb(data, censored, gamma1, q, plot = FALSE, add = FALSE,
      main = "Estimates of small exceedance probability", ...)

cReturn(data, censored, gamma1, q, plot = FALSE, add = FALSE,
      main = "Estimates of large return period", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">cHill</a> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for cProb and "Estimates of large return period" for cReturn.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The probability is estimated as

$$\hat{P}(X > q) = (1 - km) \times (q/Z_{n-k,n})^{-1/H_{k,n}^c}$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $H_{k,n}^c$  the Hill estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ .

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for cProb.
R	Vector of the corresponding estimates for the return period, only returned for cReturn.
q	The used large quantile.

**Author(s)**

Tom Reynkens

**References**

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

**See Also**

[cHill](#), [cQuant](#), [Prob](#), [KaplanMeier](#)

**Examples**

```

# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Hill estimator adapted for right censoring
chill <- cHill(Z, censored=censored, plot=TRUE)

# Small exceedance probability
q <- 10
cProb(Z, censored=censored, gamma1=chill$gamma1, q=q, plot=TRUE)

# Return period
cReturn(Z, censored=censored, gamma1=chill$gamma1, q=q, plot=TRUE)

```

---

cProbEPD

*Estimator of small exceedance probabilities and large return periods  
using censored EPD*


---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the parameters from the EPD fit adapted for right censoring.

**Usage**

```
cProbEPD(data, censored, gamma1, kappa1, beta, q, plot = FALSE, add = FALSE,
          main = "Estimates of small exceedance probability", ...)
```

```
cReturnEPD(data, censored, gamma1, kappa1, beta, q, plot = FALSE, add = FALSE,
            main = "Estimates of large return period", ...)
```

**Arguments**

**data**                Vector of  $n$  observations.

**censored**            A logical vector of length  $n$  indicating if an observation is censored.

**gamma1**              Vector of  $n - 1$  estimates for the EVI obtained from [cEPD](#).

kappa1	Vector of $n - 1$ estimates for $\kappa_1$ obtained from <a href="#">cEPD</a> .
beta	Vector of $n - 1$ estimates for $\beta$ obtained from <a href="#">cEPD</a> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for cProbEPD and "Estimates of large return period" for cReturnEPD.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The probability is estimated as

$$\hat{P}(X > q) = (1 - km) \times (1 - F(q))$$

with  $F$  the CDF of the EPD with estimated parameters  $\hat{\gamma}_1$ ,  $\hat{\kappa}_1$  and  $\hat{\tau} = -\hat{\beta}$  and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$  (the  $(k + 1)$ -th largest data point).

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for cProbEPD.
R	Vector of the corresponding estimates for the return period, only returned for cReturnEPD.
q	The used large quantile.

### Author(s)

Tom Reynkens.

### References

Beirlant, J., Bardoutsos, A., de Wet, T. and Gijbels, I. (2016). "Bias Reduced Tail Estimation for Censored Pareto Type Distributions." *Statistics & Probability Letters*, 109, 78–88.

### See Also

[cEPD](#), [ProbEPD](#), [Prob](#), [KaplanMeier](#)

**Examples**

```

# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# EPD estimator adapted for right censoring
cepd <- cEPD(Z, censored=censored, plot=TRUE)

# Small exceedance probability
q <- 10
cProbEPD(Z, censored=censored, gamma1=cepd$gamma1,
         kappa1=cepd$kappa1, beta=cepd$beta, q=q, plot=TRUE)

# Return period
cReturnEPD(Z, censored=censored, gamma1=cepd$gamma1,
          kappa1=cepd$kappa1, beta=cepd$beta, q=q, plot=TRUE)

```

---

cProbGH

*Estimator of small exceedance probabilities and large return periods  
using censored generalised Hill*

---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the estimates for the EVI obtained from the generalised Hill estimator adapted for right censoring.

**Usage**

```

cProbGH(data, censored, gamma1, q, plot = FALSE, add = FALSE,
        main = "Estimates of small exceedance probability", ...)

cReturnGH(data, censored, gamma1, q, plot = FALSE, add = FALSE,
         main = "Estimates of large return period", ...)

```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">cgenHill</a> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for cProbGH and "Estimates of large return period" for cReturnGH.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The probability is estimated as

$$\hat{P}(X > q) = (1 - km) \times (1 + \hat{\gamma}_1/a_{k,n} \times (q - Z_{n-k,n}))^{-1/\hat{\gamma}_1}$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\hat{\gamma}_1$  the generalised Hill estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ . The value  $a$  is defined as

$$a_{k,n} = Z_{n-k,n} H_{k,n} (1 - \min(\hat{\gamma}_1, 0)) / \hat{p}_k$$

with  $H_{k,n}$  the ordinary Hill estimator and  $\hat{p}_k$  the proportion of the  $k$  largest observations that is non-censored.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for cProbGH.
R	Vector of the corresponding estimates for the return period, only returned for cReturnGH.
q	The used large quantile.

**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[cQuantGH](#), [cgenHill](#), [ProbGH](#), [cProbMOM](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Generalised Hill estimator adapted for right censoring
cghill <- cgenHill(Z, censored=censored, plot=TRUE)

# Small exceedance probability
q <- 10
cProbGH(Z, censored=censored, gamma1=cghill$gamma1, q=q, plot=TRUE)

# Return period
cReturnGH(Z, censored=censored, gamma1=cghill$gamma1, q=q, plot=TRUE)
```

---

cProbGPD

*Estimator of small exceedance probabilities and large return periods  
using censored GPD-MLE*

---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the GPD-ML estimator adapted for right censoring.

**Usage**

```
cProbGPD(data, censored, gamma1, sigma1, q, plot = FALSE, add = FALSE,
          main = "Estimates of small exceedance probability", ...)

cReturnGPD(data, censored, gamma1, sigma1, q, plot = FALSE, add = FALSE,
            main = "Estimates of large return period", ...)
```



**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <code>cGPDmle</code> .
sigma1	Vector of $n - 1$ estimates for $\sigma_1$ obtained from <code>cGPDmle</code> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for <code>cProbGPD</code> and "Estimates of large return period" for <code>cReturnGPD</code> .
...	Additional arguments for the plot function, see <code>plot</code> for more details.

**Details**

The probability is estimated as

$$\hat{P}(X > q) = (1 - km) \times (1 + \hat{\gamma}_1/a_{k,n} \times (q - Z_{n-k,n}))^{-1/\hat{\gamma}_1}$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\hat{\gamma}_1$  the generalised Hill estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ . The value  $a$  is defined as

$$a_{k,n} = \hat{\sigma}_1/\hat{p}_k$$

with  $\hat{\sigma}_1$  the ML estimate for  $\sigma_1$  and  $\hat{p}_k$  the proportion of the  $k$  largest observations that is non-censored.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for <code>cProbGPD</code> .
R	Vector of the corresponding estimates for the return period, only returned for <code>cReturnGPD</code> .
q	The used large quantile.

**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[cQuantGPD](#), [cGPDmle](#), [ProbGPD](#), [Prob](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# GPD-MLE estimator adapted for right censoring
cpot <- cGPDmle(Z, censored=censored, plot=TRUE)

# Exceedance probability
q <- 10
cProbGPD(Z, gamma1=cpot$gamma1, sigma1=cpot$sigma1,
          censored=censored, q=q, plot=TRUE)

# Return period
cReturnGPD(Z, gamma1=cpot$gamma1, sigma1=cpot$sigma1,
            censored=censored, q=q, plot=TRUE)
```

---

cProbMOM

*Estimator of small exceedance probabilities and large return periods  
using censored MOM*

---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the Method of Moments estimates for the EVI adapted for right censoring.

**Usage**

```
cProbMOM(data, censored, gamma1, q, plot = FALSE, add = FALSE,
          main = "Estimates of small exceedance probability", ...)

cReturnMOM(data, censored, gamma1, q, plot = FALSE, add = FALSE,
            main = "Estimates of large return period", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <code>cMoment</code> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for <code>cProbMOM</code> and "Estimates of large return period" for <code>cReturnMOM</code> .
...	Additional arguments for the plot function, see <code>plot</code> for more details.

**Details**

The probability is estimated as

$$\hat{P}(X > q) = (1 - km) \times (1 + \hat{\gamma}_1/a_{k,n} \times (q - Z_{n-k,n}))^{-1/\hat{\gamma}_1}$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\hat{\gamma}_1$  the MOM estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ . The value  $a$  is defined as

$$a_{k,n} = Z_{n-k,n} H_{k,n} (1 - \min(\hat{\gamma}_1, 0)) / \hat{p}_k$$

with  $H_{k,n}$  the ordinary Hill estimator and  $\hat{p}_k$  the proportion of the  $k$  largest observations that is non-censored.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for <code>cProbMOM</code> .
R	Vector of the corresponding estimates for the return period, only returned for <code>cReturnMOM</code> .
q	The used large quantile.

**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[cQuantMOM](#), [cMoment](#), [ProbMOM](#), [Prob](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Moment estimator adapted for right censoring
cmom <- cMoment(Z, censored=censored, plot=TRUE)

# Small exceedance probability
q <- 10
cProbMOM(Z, censored=censored, gamma1=cmom$gamma1, q=q, plot=TRUE)

# Return period
cReturnMOM(Z, censored=censored, gamma1=cmom$gamma1, q=q, plot=TRUE)
```

---

cQuant

*Estimator of large quantiles using censored Hill*


---

**Description**

Computes estimates of large quantiles  $Q(1 - p)$  using the estimates for the EVI obtained from the Hill estimator adapted for right censoring.

**Usage**

```
cQuant(data, censored, gamma1, p, plot = FALSE, add = FALSE,
       main = "Estimates of extreme quantile", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">cHill</a> .

p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The quantile is estimated as

$$\hat{Q}(1 - p) = Z_{n-k,n} \times ((1 - km)/p)^{H_{k,n}^c}$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $H_{k,n}^c$  the Hill estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ .

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

### Author(s)

Tom Reynkens.

### References

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

### See Also

[cHill](#), [cProb](#), [Quant](#), [KaplanMeier](#)

### Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
```

```

Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Hill estimator adapted for right censoring
chill <- cHill(Z, censored=censored, plot=TRUE)

# Large quantile
p <- 10^(-4)
cQuant(Z, gamma1=chill$gamma, censored=censored, p=p, plot=TRUE)

```

---

cQuantGH

*Estimator of large quantiles using censored Hill*


---

### Description

Computes estimates of large quantiles  $Q(1 - p)$  using the estimates for the EVI obtained from the generalised Hill estimator adapted for right censoring.

### Usage

```

cQuantGH(data, censored, gamma1, p, plot = FALSE, add = FALSE,
          main = "Estimates of extreme quantile", ...)

```

### Arguments

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">cgenHill</a> .
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The quantile is estimated as

$$\hat{Q}(1 - p) = Z_{n-k,n} + a_{k,n}(((1 - km)/p)^{\hat{\gamma}_1} - 1)/\hat{\gamma}_1$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\hat{\gamma}_1$  the generalised Hill estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ . The value  $a$  is defined as

$$a_{k,n} = Z_{n-k,n} H_{k,n}(1 - S_{Z,k,n})/\hat{p}_k$$

with  $H_{k,n}$  the ordinary Hill estimator and  $\hat{p}_k$  the proportion of the  $k$  largest observations that is non-censored, and

$$S_{Z,k,n} = 1 - (1 - M_1^2/M_2)^{(1 - 1/k)/2}$$

with

$$M_l = 1/k \sum_{j=1}^k (\log X_{n-j+1,n} - \log X_{n-k,n})^l.$$

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

### Author(s)

Tom Reynkens

### References

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

### See Also

[cProbGH](#), [cgenHill](#), [QuantGH](#), [Quant](#), [KaplanMeier](#)

### Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Generalised Hill estimator adapted for right censoring
cghill <- cgenHill(Z, censored=censored, plot=TRUE)

# Large quantile
p <- 10^(-4)
cQuantGH(Z, gamma=cghill$gamma, censored=censored, p=p, plot=TRUE)
```

---

cQuantGPD

*Estimator of large quantiles using censored GPD-MLE*


---

### Description

Computes estimates of large quantiles  $Q(1 - p)$  using the estimates for the EVI obtained from the GPD-ML estimator adapted for right censoring.

### Usage

```
cQuantGPD(data, censored, gamma1, sigma1, p, plot = FALSE, add = FALSE,
           main = "Estimates of extreme quantile", ...)
```

### Arguments

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">cGPDmle</a> .
sigma1	Vector of $n - 1$ estimates for $\sigma_1$ obtained from <a href="#">cGPDmle</a> .
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The quantile is estimated as

$$\hat{Q}(1 - p) = Z_{n-k,n} + a_{k,n}(((1 - km)/p)^{\hat{\gamma}_1} - 1)/\hat{\gamma}_1$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\hat{\gamma}_1$  the generalised Hill estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ . The value  $a$  is defined as

$$a_{k,n} = \hat{\sigma}_1 / \hat{p}_k$$

with  $\hat{\sigma}_1$  the ML estimate for  $\sigma_1$  and  $\hat{p}_k$  the proportion of the  $k$  largest observations that is non-censored.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.



**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[cProbGPD](#), [cGPDmle](#), [QuantGPD](#), [Quant](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# GPD-MLE estimator adapted for right censoring
cpot <- cGPDmle(Z, censored=censored, plot=TRUE)

# Large quantile
p <- 10^(-4)
cQuantGPD(Z, gamma1=cpot$gamma1, sigma1=cpot$sigma1,
          censored=censored, p=p, plot=TRUE)
```

---

cQuantMOM

*Estimator of large quantiles using censored MOM*


---

**Description**

Computes estimates of large quantiles  $Q(1 - p)$  using the estimates for the EVI obtained from the MOM estimator adapted for right censoring.

**Usage**

```
cQuantMOM(data, censored, gamma1, p, plot = FALSE, add = FALSE,
          main = "Estimates of extreme quantile", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
gamma1	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">cMoment</a> .
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The quantile is estimated as

$$\hat{Q}(1 - p) = Z_{n-k,n} + a_{k,n}(((1 - km)/p)^{\hat{\gamma}_1} - 1)/\hat{\gamma}_1$$

with  $Z_{i,n}$  the  $i$ -th order statistic of the data,  $\hat{\gamma}_1$  the MOM estimator adapted for right censoring and  $km$  the Kaplan-Meier estimator for the CDF evaluated in  $Z_{n-k,n}$ . The value  $a$  is defined as

$$a_{k,n} = Z_{n-k,n} H_{k,n} (1 - \min(\hat{\gamma}_1, 0)) / \hat{p}_k$$

with  $H_{k,n}$  the ordinary Hill estimator and  $\hat{p}_k$  the proportion of the  $k$  largest observations that is non-censored.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

**Author(s)**

Tom Reynkens

**References**

Einmahl, J.H.J., Fils-Villetard, A. and Guillou, A. (2008). "Statistics of Extremes Under Random Censoring." *Bernoulli*, 14, 207–227.

**See Also**

[cProbMOM](#), [cMoment](#), [QuantMOM](#), [Quant](#), [KaplanMeier](#)

**Examples**

```

# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Moment estimator adapted for right censoring
cmom <- cMoment(Z, censored=censored, plot=TRUE)

# Large quantile
p <- 10^(-4)
cQuantMOM(Z, censored=censored, gamma1=cmom$gamma1, p=p, plot=TRUE)

```

---

crHill

*Hill-type estimator for the conditional EVI*


---

**Description**

Hill-type estimator for the conditional Extreme Value Index (EVI) adapted for censored data.

**Usage**

```

crHill(x, Xtilde, Ytilde, censored, h,
       kernel = c("biweight", "normal", "uniform", "triangular", "epanechnikov"),
       logk = FALSE, plot = FALSE, add = FALSE, main = "", ...)

```

**Arguments**

x	Value of the conditioning variable $X$ to estimate the EVI at.
Xtilde	Vector of length $n$ containing the censored sample of the conditioning variable $X$ .
Ytilde	Vector of length $n$ containing the censored sample of the variable $Y$ .
censored	A logical vector of length $n$ indicating if an observation is censored.
h	Bandwidth of the non-parametric estimator.
kernel	Kernel of the non-parametric estimator. One of "biweight" (default), "normal", "uniform", "triangular" and "epanechnikov".

logk	Logical indicating if the Hill-type estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "" (no title).
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

This is a Hill-type estimator of the EVI of  $Y$  given  $X = x$ . The estimator uses the censored sample  $(\tilde{X}_i, \tilde{Y}_i)$ , for  $i = 1, \dots, n$ , where  $X$  and  $Y$  are censored at the same time. We assume that  $Y$  and the censoring variable are conditionally independent given  $X$ .

See Section 4.4.3 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding Hill-type estimates.

### Author(s)

Tom Reynkens

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

### See Also

[crParetoQQ](#), [crSurv](#), [cHill](#)

### Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
Y <- rpareto(200, shape=2)

# Censoring variable
C <- rpareto(200, shape=1)

# Observed (censored) sample of variable Y
Ytilde <- pmin(Y, C)
```

```

# Censoring indicator
censored <- (Y>C)

# Conditioning variable
X <- seq(1, 10, length.out=length(Y))

# Observed (censored) sample of conditioning variable
Xtilde <- X
Xtilde[censored] <- X[censored] - runif(sum(censored), 0, 1)

# Conditional Pareto QQ-plot
crParetoQQ(x=1, Xtilde=Xtilde, Ytilde=Ytilde, censored=censored, h=2)

# Plot Hill-type estimates
crHill(x=1, Xtilde, Ytilde, censored, h=2, plot=TRUE)

```

---

crParetoQQ

*Conditional Pareto quantile plot for right censored data*


---

### Description

Conditional Pareto QQ-plot adapted for right censored data.

### Usage

```

crParetoQQ(x, Xtilde, Ytilde, censored, h,
           kernel = c("biweight", "normal", "uniform", "triangular", "epanechnikov"),
           plot = TRUE, add = FALSE, main = "Pareto QQ-plot", type = "p", ...)

```

### Arguments

x	Value of the conditioning variable $X$ at which to make the conditional Pareto QQ-plot.
Xtilde	Vector of length $n$ containing the censored sample of the conditioning variable $X$ .
Ytilde	Vector of length $n$ containing the censored sample of the variable $Y$ .
censored	A logical vector of length $n$ indicating if an observation is censored.
h	Bandwidth of the non-parametric estimator for the conditional survival function ( <a href="#">crSurv</a> ).
kernel	Kernel of the non-parametric estimator for the conditional survival function ( <a href="#">crSurv</a> ). One of "biweight" (default), "normal", "uniform", "triangular" and "epanechnikov".
plot	Logical indicating if the quantiles should be plotted in a Pareto QQ-plot, default is TRUE.

add	Logical indicating if the quantiles should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Pareto QQ-plot".
type	Type of the plot, default is "p" meaning points are plotted, see <a href="#">plot</a> for more details.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

We construct a Pareto QQ-plot for  $Y$  conditional on  $X = x$  using the censored sample  $(\tilde{X}_i, \tilde{Y}_i)$ , for  $i = 1, \dots, n$ , where  $X$  and  $Y$  are censored at the same time. We assume that  $Y$  and the censoring variable are conditionally independent given  $X$ .

The conditional Pareto QQ-plot adapted for right censoring is given by

$$(-\log(1 - \hat{F}_{Y|X}(\tilde{Y}_{j,n}|x)), \log \tilde{Y}_{j,n})$$

for  $j = 1, \dots, n - 1$ , with  $\tilde{Y}_{i,n}$  the  $i$ -th order statistic of the censored data and  $\hat{F}_{Y|X}(y|x)$  the non-parametric estimator for the conditional CDF of Akritas and Van Keilegom (2003), see [crSurv](#).

See Section 4.4.3 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

pqq.the	Vector of the theoretical quantiles, see Details.
pqq.emp	Vector of the empirical quantiles from the log-transformed $Y$ data.

### Author(s)

Tom Reynkens

### References

Akritas, M.G. and Van Keilegom, I. (2003). "Estimation of Bivariate and Marginal Distributions With Censored Data." *Journal of the Royal Statistical Society: Series B*, 65, 457–471.

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

### See Also

[crSurv](#), [crHill](#), [cParetoQQ](#)

### Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
Y <- rpareto(200, shape=2)
```

```

# Censoring variable
C <- rpareto(200, shape=1)

# Observed (censored) sample of variable Y
Ytilde <- pmin(Y, C)

# Censoring indicator
censored <- (Y>C)

# Conditioning variable
X <- seq(1, 10, length.out=length(Y))

# Observed (censored) sample of conditioning variable
Xtilde <- X
Xtilde[censored] <- X[censored] - runif(sum(censored), 0, 1)

# Conditional Pareto QQ-plot
crParetoQQ(x=1, Xtilde=Xtilde, Ytilde=Ytilde, censored=censored, h=2)

# Plot Hill-type estimates
crHill(x=1, Xtilde, Ytilde, censored, h=2, plot=TRUE)

```

---

crSurv

*Non-parametric estimator of conditional survival function*


---

### Description

Non-parametric estimator of the conditional survival function of  $Y$  given  $X$  for censored data, see Akritas and Van Keilegom (2003).

### Usage

```
crSurv(x, y, Xtilde, Ytilde, censored, h,
       kernel = c("biweight", "normal", "uniform", "triangular", "epanechnikov"))
```

### Arguments

<code>x</code>	The value of the conditioning variable $X$ to evaluate the survival function at. <code>x</code> needs to be a single number or a vector with the same length as <code>y</code> .
<code>y</code>	The value(s) of the variable $Y$ to evaluate the survival function at.
<code>Xtilde</code>	Vector of length $n$ containing the censored sample of the conditioning variable $X$ .
<code>Ytilde</code>	Vector of length $n$ containing the censored sample of the variable $Y$ .
<code>censored</code>	A logical vector of length $n$ indicating if an observation is censored.
<code>h</code>	Bandwidth of the non-parametric estimator.
<code>kernel</code>	Kernel of the non-parametric estimator. One of "biweight" (default), "normal", "uniform", "triangular" and "epanechnikov".

## Details

We estimate the conditional survival function

$$1 - F_{Y|X}(y|x)$$

using the censored sample  $(\tilde{X}_i, \tilde{Y}_i)$ , for  $i = 1, \dots, n$ , where  $X$  and  $Y$  are censored at the same time. We assume that  $Y$  and the censoring variable are conditionally independent given  $X$ .

The estimator is given by

$$1 - \hat{F}_{Y|X}(y|x) = \prod_{\tilde{Y}_i \leq y} (1 - W_{n,i}(x; h_n) / (\sum_{j=1}^n W_{n,j}(x; h_n) I\{\tilde{Y}_j \geq \tilde{Y}_i\}))^{\Delta_i}$$

where  $\Delta_i = 1$  when  $(\tilde{X}_i, \tilde{Y}_i)$  is censored and 0 otherwise. The weights are given by

$$W_{n,i}(x; h_n) = K((x - \tilde{X}_i)/h_n) / \sum_{\Delta_j=1} K((x - \tilde{X}_j)/h_n)$$

when  $\Delta_i = 1$  and 0 otherwise.

See Section 4.4.3 in Albrecher et al. (2017) for more details.

## Value

Estimates for  $1 - F_{Y|X}(y|x)$  as described above.

## Author(s)

Tom Reynkens

## References

- Akritis, M.G. and Van Keilegom, I. (2003). "Estimation of Bivariate and Marginal Distributions With Censored Data." *Journal of the Royal Statistical Society: Series B*, 65, 457–471.
- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

## See Also

[crParetoQQ](#), [crHill](#)

## Examples

```
# Set seed
set.seed(29072016)

# Pareto random sample
Y <- rpareto(200, shape=2)

# Censoring variable
C <- rpareto(200, shape=1)
```



```

# Observed (censored) sample of variable Y
Ytilde <- pmin(Y, C)

# Censoring indicator
censored <- (Y>C)

# Conditioning variable
X <- seq(1, 10, length.out=length(Y))

# Observed (censored) sample of conditioning variable
Xtilde <- X
Xtilde[censored] <- X[censored] - runif(sum(censored), 0, 1)

# Plot estimates of the conditional survival function
x <- 5
y <- seq(0, 5, 1/100)
plot(y, crSurv(x, y, Xtilde=Xtilde, Ytilde=Ytilde, censored=censored, h=5), type="l",
      xlab="y", ylab="Conditional survival function")

```

CTE

*Conditional Tail Expectation***Description**

Compute Conditional Tail Expectation (CTE)  $CTE_{1-p}$  of the fitted spliced distribution.

**Usage**

```
CTE(p, splicefit)
```

```
ES(p, splicefit)
```

**Arguments**

p	The probability associated with the CTE (we estimate $CTE_{1-p}$ ).
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> , <a href="#">SpliceFiticPareto</a> or <a href="#">SpliceFitGPD</a> .

**Details**

The Conditional Tail Expectation is defined as

$$CTE_{1-p} = E(X|X > Q(1-p)) = E(X|X > VaR_{1-p}) = VaR_{1-p} + \Pi(VaR_{1-p})/p,$$

where  $\Pi(u) = E((X - u)_+)$  is the premium of the excess-loss insurance with retention  $u$ .

If the CDF is continuous in  $p$ , we have  $CTE_{1-p} = TVaR_{1-p} = 1/p \int_0^p VaR_{1-s} ds$  with  $TVaR$  the Tail Value-at-Risk.

See Reynkens et al. (2017) and Section 4.6 of Albrecher et al. (2017) for more details.

The ES function is the same function as CTE but is deprecated.

**Value**

Vector with the CTE corresponding to each element of  $p$ .

**Author(s)**

Tom Reynkens with R code from Roel Verbelen for the mixed Erlang quantiles.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[qSplice](#), [ExcessSplice](#), [SpliceFit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

p <- seq(0.01, 0.99, 0.01)
# Plot of CTE
plot(p, CTE(p, splicefit), type="l", xlab="p", ylab=bquote(CTE[1-p]))

## End(Not run)
```

---

cWeibullQQ

*Weibull quantile plot for right censored data*


---

**Description**

Weibull QQ-plot adapted for right censored data.

**Usage**

```
cWeibullQQ(data, censored, plot = TRUE, main = "Weibull QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
censored	A logical vector of length $n$ indicating if an observation is censored.
plot	Logical indicating if the quantiles should be plotted in a Weibull QQ-plot, default is TRUE.
main	Title for the plot, default is "Weibull QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The Weibull QQ-plot adapted for right censoring is given by

$$(\log(-\log(1 - F_{km}(Z_{j,n}))), \log(Z_{j,n}))$$

for  $j = 1, \dots, n - 1$ , with  $Z_{i,n}$  the  $i$ -th order statistic of the data and  $F_{km}$  the Kaplan-Meier estimator for the CDF. Hence, it has the same empirical quantiles as an ordinary Weibull QQ-plot but replaces the theoretical quantiles  $\log(-\log(1 - j/(n + 1)))$  by  $\log(-\log(1 - F_{km}(Z_{j,n})))$ .

This QQ-plot is only suitable for right censored data.

In Beirlant et al. (2007), only a Pareto QQ-plot adapted for right-censored data is proposed. This QQ-plot is constructed using the same ideas, but is not described in the paper.

**Value**

A list with following components:

wqq.the	Vector of the theoretical quantiles, see Details.
wqq.emp	Vector of the empirical quantiles from the log-transformed data.

**Author(s)**

Tom Reynkens

**References**

Beirlant, J., Guillou, A., Dierckx, G. and Fils-Villetard, A. (2007). "Estimation of the Extreme Value Index and Extreme Quantiles Under Random Censoring." *Extremes*, 10, 151–174.

**See Also**

[WeibullQQ](#), [cExpQQ](#), [cLognormalQQ](#), [cParetoQQ](#), [KaplanMeier](#)

**Examples**

```
# Set seed
set.seed(29072016)

# Pareto random sample
X <- rpareto(500, shape=2)
```

```

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Weibull QQ-plot adapted for right censoring
cWeibullQQ(Z, censored=censored)

```

---

EPD

*EPD estimator*


---

### Description

Fit the Extended Pareto Distribution (GPD) to the exceedances (peaks) over a threshold. Optionally, these estimates are plotted as a function of  $k$ .

### Usage

```

EPD(data, rho = -1, start = NULL, direct = FALSE, warnings = FALSE,
     logk = FALSE, plot = FALSE, add = FALSE, main = "EPD estimates of the EVI", ...)

```

### Arguments

<code>data</code>	Vector of $n$ observations.
<code>rho</code>	A parameter for the $\rho$ -estimator of Fraga Alves et al. (2003) when strictly positive or choice(s) for $\rho$ if negative. Default is -1.
<code>start</code>	Vector of length 2 containing the starting values for the optimisation. The first element is the starting value for the estimator of $\gamma$ and the second element is the starting value for the estimator of $\kappa$ . This argument is only used when <code>direct=TRUE</code> . Default is NULL meaning the initial value for $\gamma$ is the Hill estimator and the initial value for $\kappa$ is 0.
<code>direct</code>	Logical indicating if the parameters are obtained by directly maximising the log-likelihood function, see Details. Default is FALSE.
<code>warnings</code>	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.
<code>logk</code>	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is FALSE.
<code>plot</code>	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
<code>add</code>	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
<code>main</code>	Title for the plot, default is "EPD estimates of the EVI".
<code>...</code>	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

## Details

We fit the Extended Pareto distribution to the relative excesses over a threshold ( $X/u$ ). The EPD has distribution function  $F(x) = 1 - (x(1 + \kappa - \kappa x^\tau))^{-1/\gamma}$  with  $\tau = \rho/\gamma < 0 < \gamma$  and  $\kappa > \max(-1, 1/\tau)$ .

The parameters are determined using MLE and there are two possible approaches: maximise the log-likelihood directly (`direct=TRUE`) or follow the approach detailed in Beirlant et al. (2009) (`direct=FALSE`). The latter approach uses the score functions of the log-likelihood.

See Section 4.2.1 of Albrecher et al. (2017) for more details.

## Value

A list with following components:

<code>k</code>	Vector of the values of the tail parameter $k$ .
<code>gamma</code>	Vector of the corresponding estimates for the $\gamma$ parameter of the EPD.
<code>kappa</code>	Vector of the corresponding MLE estimates for the $\kappa$ parameter of the EPD.
<code>tau</code>	Vector of the corresponding estimates for the $\tau$ parameter of the EPD using Hill estimates and values for $\rho$ .

## Author(s)

Tom Reynkens

## References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant, J., Joossens, E. and Segers, J. (2009). "Second-Order Refined Peaks-Over-Threshold Modelling for Heavy-Tailed Distributions." *Journal of Statistical Planning and Inference*, 139, 2800–2815.
- Fraga Alves, M.I. , Gomes, M.I. and de Haan, L. (2003). "A New Class of Semi-parametric Estimators of the Second Order Parameter." *Portugaliae Mathematica*, 60, 193–214.

## See Also

[GPDmle](#), [ProbEPD](#)

## Examples

```
data(secura)

# EPD estimates for the EVI
epd <- EPD(secura$size, plot=TRUE)

# Compute return periods
ReturnEPD(secura$size, 10^10, gamma=epd$gamma, kappa=epd$kappa,
          tau=epd$tau, plot=TRUE)
```

---

 EPDfit

*Fit EPD using MLE*


---

**Description**

Fit the Extended Pareto Distribution (EPD) to data using Maximum Likelihood Estimation (MLE).

**Usage**

```
EPDfit(data, tau, start = c(0.1, 1), warnings = FALSE)
```

**Arguments**

data	Vector of $n$ observations.
tau	Value for the $\tau$ parameter of the EPD.
start	Vector of length 2 containing the starting values for the optimisation. The first element is the starting value for the estimator of $\gamma$ and the second element is the starting value for the estimator of $\kappa$ . Default is $c(0.1, 1)$ .
warnings	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.

**Details**

See Section 4.2.1 of Albrecher et al. (2017) for more details.

**Value**

A vector with the MLE estimate for the  $\gamma$  parameter of the EPD as the first component and the MLE estimate for the  $\kappa$  parameter of the EPD as the second component.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Joossens, E. and Segers, J. (2009). "Second-Order Refined Peaks-Over-Threshold Modelling for Heavy-Tailed Distributions." *Journal of Statistical Planning and Inference*, 139, 2800–2815.

**See Also**

[EPD](#), [GPDfit](#)

**Examples**

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Fit EPD to last 500 observations
res <- EPDfit(SOAdata/sort(soa$size)[500], tau=-1)
```

---

EVTfit

*EVT fit*

---

**Description**

Create an S3 object using an EVT (Extreme Value Theory) fit.

**Usage**

```
EVTfit(gamma, endpoint = NULL, sigma = NULL)
```

**Arguments**

gamma	Vector of estimates for $\gamma$ .
endpoint	Vector of endpoints (with the same length as gamma). When NULL (default), a vector containing Inf for each value of gamma will be used.
sigma	Vector of scale estimates for the GPD (with the same length as gamma). When NULL (default), not included in the object.

**Details**

See Reynkens et al. (2017) and Section 4.3 of Albrecher et al. (2017) for details.

**Value**

An S3 object containing the above input arguments.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

**See Also**

[SpliceFit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#)

**Examples**

```
# Create Mefit object
mefit <- Mefit(p=c(0.65,0.35), shape=c(39,58), theta=16.19, M=2)

# Create EVTfit object
evtfit <- EVTfit(gamma=c(0.76,0.64), endpoint=c(39096, Inf))

# Create SpliceFit object
splicefit <- SpliceFit(const=c(0.5,0.996), trunclower=0, t=c(1020,39096), type=c("ME","TPa","Pa"),
                      Mefit=mefit, EVTfit=evtfit)

# Show summary
summary(splicefit)
```

---

ExcessEPD

*Estimates for excess-loss premiums using EPD estimates*

---

**Description**

Estimate premiums of excess-loss reinsurance with retention  $R$  and limit  $L$  using EPD estimates.

**Usage**

```
ExcessEPD(data, gamma, kappa, tau, R, L = Inf, warnings = TRUE, plot = TRUE, add = FALSE,
          main = "Estimates for premium of excess-loss insurance", ...)
```

**Arguments**

<code>data</code>	Vector of $n$ observations.
<code>gamma</code>	Vector of $n - 1$ estimates for the EVI, obtained from <a href="#">EPD</a> .
<code>kappa</code>	Vector of $n - 1$ estimates for $\kappa$ , obtained from <a href="#">EPD</a> .
<code>tau</code>	Vector of $n - 1$ estimates for $\tau$ , obtained from <a href="#">EPD</a> .
<code>R</code>	The retention level of the (re-)insurance.
<code>L</code>	The limit of the (re-)insurance, default is <code>Inf</code> .
<code>warnings</code>	Logical indicating if warnings are displayed, default is <code>TRUE</code> .
<code>plot</code>	Logical indicating if the estimates should be plotted as a function of $k$ , default is <code>FALSE</code> .
<code>add</code>	Logical indicating if the estimates should be added to an existing plot, default is <code>FALSE</code> .
<code>main</code>	Title for the plot, default is "Estimates for premium of excess-loss insurance".
<code>...</code>	Additional arguments for the plot function, see <a href="#">plot</a> for more details.



### Details

We need that  $u \geq X_{n-k,n}$ , the  $(k+1)$ -th largest observation. If this is not the case, we return NA for the premium. A warning will be issued in that case if `warnings=TRUE`.

The premium for the excess-loss insurance with retention  $R$  and limit  $L$  is given by

$$E(\min(X - R)_+, L) = \Pi(R) - \Pi(R + L)$$

where  $\Pi(u) = E((X - u)_+) = \int_u^\infty (1 - F(z)) dz$  is the premium of the excess-loss insurance with retention  $u$ . When  $L = \infty$ , the premium is equal to  $\Pi(R)$ .

We estimate  $\Pi$  by

$$\hat{\Pi}(u) = (k+1)/(n+1) \times (X_{n-k,n})^{1/\hat{\gamma}} \times ((1 - \hat{\kappa}/\hat{\gamma})(1/\hat{\gamma} - 1)^{-1} u^{1-1/\hat{\gamma}} + \hat{\kappa}/(\hat{\gamma} X_{n-k,n}^{\hat{\tau}})(1/\hat{\gamma} - \hat{\tau} - 1)^{-1} u^{1+\hat{\tau}-1/\hat{\gamma}})$$

with  $\hat{\gamma}$ ,  $\hat{\kappa}$  and  $\hat{\tau}$  the estimates for the parameters of the EPD.

See Section 4.6 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
premium	The corresponding estimates for the premium.
R	The retention level of the (re-)insurance.
L	The limit of the (re-)insurance.

### Author(s)

Tom Reynkens

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

### See Also

[EPD](#), [ExcessHill](#), [ExcessGPD](#)

### Examples

```
data(secura)

# EPD estimator
epd <- EPD(secura$size)

# Premium of excess-loss insurance with retention R
R <- 10^7
ExcessEPD(secura$size, gamma=epd$gamma, kappa=epd$kappa, tau=epd$tau, R=R, ylim=c(0,2*10^4))
```

ExcessGPD

*Estimates for excess-loss premiums using GPD-MLE estimates***Description**

Estimate premiums of excess-loss reinsurance with retention  $R$  and limit  $L$  using GPD-MLE estimates.

**Usage**

```
ExcessGPD(data, gamma, sigma, R, L = Inf, warnings = TRUE, plot = TRUE, add = FALSE,
          main = "Estimates for premium of excess-loss insurance", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">GPDmle</a> .
sigma	Vector of $n - 1$ estimates for $\sigma$ obtained from <a href="#">GPDmle</a> .
R	The retention level of the (re-)insurance.
L	The limit of the (re-)insurance, default is Inf.
warnings	Logical indicating if warnings are displayed, default is TRUE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates for premium of excess-loss insurance".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

We need that  $u \geq X_{n-k,n}$ , the  $(k+1)$ -th largest observation. If this is not the case, we return NA for the premium. A warning will be issued in that case if `warnings=TRUE`. One should then use global fits: [ExcessSplice](#).

The premium for the excess-loss insurance with retention  $R$  and limit  $L$  is given by

$$E(\min(X - R)_+, L) = \Pi(R) - \Pi(R + L)$$

where  $\Pi(u) = E((X - u)_+) = \int_u^\infty (1 - F(z)) dz$  is the premium of the excess-loss insurance with retention  $u$ . When  $L = \infty$ , the premium is equal to  $\Pi(R)$ .

We estimate  $\Pi$  by

$$\hat{\Pi}(u) = (k + 1)/(n + 1) \times \hat{\sigma}_k / (1 - \hat{\gamma}_k) \times (1 + \hat{\gamma}_k / \hat{\sigma}_k (u - X_{n-k,n}))^{1-1/\hat{\gamma}_k},$$

with  $\hat{\gamma}_k$  and  $\hat{\sigma}_k$  the estimates for the parameters of the GPD.

See Section 4.6 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
premium	The corresponding estimates for the premium.
R	The retention level of the (re-)insurance.
L	The limit of the (re-)insurance.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[GPDmle](#), [ExcessHill](#), [ExcessEPD](#)

**Examples**

```
data(secura)

# GPDmle estimator
mle <- GPDmle(secura$size)

# Premium of excess-loss insurance with retention R
R <- 10^7
ExcessGPD(secura$size, gamma=mle$gamma, sigma=mle$sigma, R=R, ylim=c(0,2*10^4))
```

---

ExcessPareto

*Estimates for excess-loss premiums using a Pareto model*

---

**Description**

Estimate premiums of excess-loss reinsurance with retention  $R$  and limit  $L$  using a (truncated) Pareto model.

**Usage**

```
ExcessPareto(data, gamma, R, L = Inf, endpoint = Inf, warnings = TRUE, plot = TRUE,
             add = FALSE, main = "Estimates for premium of excess-loss insurance", ...)

ExcessHill(data, gamma, R, L = Inf, endpoint = Inf, warnings = TRUE, plot = TRUE,
           add = FALSE, main = "Estimates for premium of excess-loss insurance", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI, obtained from <a href="#">Hill</a> or <a href="#">trHill</a> .
R	The retention level of the (re-)insurance.
L	The limit of the (re-)insurance, default is Inf.
endpoint	Endpoint for the truncated Pareto distribution. When Inf, the default, the ordinary Pareto model is used.
warnings	Logical indicating if warnings are displayed, default is TRUE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates for premium of excess-loss insurance".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

We need that  $u \geq X_{n-k,n}$ , the  $(k+1)$ -th largest observation. If this is not the case, we return NA for the premium. A warning will be issued in that case if `warnings=TRUE`. One should then use global fits: [ExcessSplice](#).

The premium for the excess-loss insurance with retention  $R$  and limit  $L$  is given by

$$E(\min(X - R)_+, L) = \Pi(R) - \Pi(R + L)$$

where  $\Pi(u) = E((X - u)_+) = \int_u^\infty (1 - F(z)) dz$  is the premium of the excess-loss insurance with retention  $u$ . When  $L = \infty$ , the premium is equal to  $\Pi(R)$ .

We estimate  $\Pi$  (for the untruncated Pareto distribution) by

$$\hat{\Pi}(u) = (k+1)/(n+1)/(1/H_{k,n} - 1) \times (X_{n-k,n}^{1/H_{k,n}} u^{1-1/H_{k,n}}),$$

with  $H_{k,n}$  the Hill estimator.

The `ExcessHill` function is the same function but with a different name for compatibility with old versions of the package.

See Section 4.6 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
premium	The corresponding estimates for the premium.
R	The retention level of the (re-)insurance.
L	The limit of the (re-)insurance.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[Hill](#), [ExcessEPD](#), [ExcessGPD](#), [ExcessSplice](#)

**Examples**

```
data(secura)

# Hill estimator
H <- Hill(secura$size)

# Premium of excess-loss insurance with retention R
R <- 10^7
ExcessPareto(secura$size, H$gamma, R=R)
```

---

 ExcessSplice

---

*Estimates for excess-loss premiums using splicing*


---

**Description**

Estimate premiums of excess-loss reinsurance with retention  $R$  and limit  $L$  using fitted spliced distribution.

**Usage**

```
ExcessSplice(R, L=Inf, splicefit)
```

**Arguments**

R	The retention level of the (re-)insurance or a vector of retention levels for the (re-)insurance.
L	The limit for the (re-)insurance or a vector of limits for the (re-)insurance, default is Inf.
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> , <a href="#">SpliceFiticPareto</a> or <a href="#">SpliceFitGPD</a> .

**Details**

The premium for the excess-loss insurance with retention  $R$  and limit  $L$  is given by

$$E(\min(X - R)_+, L) = \Pi(R) - \Pi(R + L)$$

where  $\Pi(u) = E((X - u)_+) = \int_u^\infty (1 - F(z))dz$  is the premium of the excess-loss insurance with retention  $u$ . When  $L = \infty$ , the premium is equal to  $\Pi(R)$ .

See Reynkens et al. (2017) and Section 4.6 of Albrecher et al. (2017) for more details.

**Value**

An estimate for the premium is returned (for every value of  $R$ ).

**Author(s)**

Tom Reynkens with R code from Roel Verbelen for the estimates for the excess-loss premiums using the mixed Erlang distribution.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[SpliceFit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.8)

# Excess-loss premium
ExcessSplice(R=2, splicefit=splicefit)

## End(Not run)
```

---

ExpQQ                      *Exponential quantile plot*

---

### Description

Computes the empirical quantiles of a data vector and the theoretical quantiles of the standard exponential distribution. These quantiles are then plotted in an exponential QQ-plot with the theoretical quantiles on the  $x$ -axis and the empirical quantiles on the  $y$ -axis.

### Usage

```
ExpQQ(data, plot = TRUE, main = "Exponential QQ-plot", ...)
```

### Arguments

data	Vector of $n$ observations.
plot	Logical indicating if the quantiles should be plotted in an Exponential QQ-plot, default is TRUE.
main	Title for the plot, default is "Exponential QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The exponential QQ-plot is defined as

$$(-\log(1 - i/(n + 1)), X_{i,n})$$

for  $i = 1, \dots, n$ , with  $X_{i,n}$  the  $i$ -th order statistic of the data.

Note that the mean excess plot is the derivative plot of the Exponential QQ-plot.

See Section 4.1 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

eqq.the	Vector of the theoretical quantiles from a standard exponential distribution.
eqq.emp	Vector of the empirical quantiles from the data.

### Author(s)

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[MeanExcess](#), [LognormalQQ](#), [ParetoQQ](#), [WeibullQQ](#)

**Examples**

```
data(norwegianfire)

# Exponential QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ExpQQ(norwegianfire$size[norwegianfire$year==76])

# Pareto QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ParetoQQ(norwegianfire$size[norwegianfire$year==76])
```

---

Extended Pareto

*The Extended Pareto Distribution*

---

**Description**

Density, distribution function, quantile function and random generation for the Extended Pareto Distribution (EPD).

**Usage**

```
depd(x, gamma, kappa, tau = -1, log = FALSE)
pepd(x, gamma, kappa, tau = -1, lower.tail = TRUE, log.p = FALSE)
qepd(p, gamma, kappa, tau = -1, lower.tail = TRUE, log.p = FALSE)
repd(n, gamma, kappa, tau = -1)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
gamma	The $\gamma$ parameter of the EPD, a strictly positive number.
kappa	The $\kappa$ parameter of the EPD. It should be larger than $\max\{-1, 1/\tau\}$ .
tau	The $\tau$ parameter of the EPD, a strictly negative number. Default is -1.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the EPD is equal to  $F(x) = 1 - (x(1 + \kappa - \kappa x^\tau))^{-1/\gamma}$  for all  $x > 1$  and  $F(x) = 0$  otherwise.

Note that an EPD random variable with  $\tau = -1$  and  $\kappa = \gamma/\sigma - 1$  is GPD distributed with  $\mu = 1$ ,  $\gamma$  and  $\sigma$ .



**Value**

depd gives the density function evaluated in  $x$ , pepd the CDF evaluated in  $x$  and qepd the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

repd returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**References**

Beirlant, J., Joossens, E. and Segers, J. (2009). "Second-Order Refined Peaks-Over-Threshold Modelling for Heavy-Tailed Distributions." *Journal of Statistical Planning and Inference*, 139, 2800–2815.

**See Also**

[Pareto, GPD, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, depd(x, gamma=1/2, kappa=1, tau=-1), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, pepd(x, gamma=1/2, kappa=1, tau=-1), xlab="x", ylab="CDF", type="l")
```

---

Frechet

*The Frechet distribution*

---

**Description**

Density, distribution function, quantile function and random generation for the Fréchet distribution (inverse Weibull distribution).

**Usage**

```
dfrechet(x, shape, loc = 0, scale = 1, log = FALSE)
pfrechet(x, shape, loc = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
qfrechet(p, shape, loc = 0, scale = 1, lower.tail = TRUE, log.p = FALSE)
rfrechet(n, shape, loc = 0, scale = 1)
```

**Arguments**

<code>x</code>	Vector of quantiles.
<code>p</code>	Vector of probabilities.
<code>n</code>	Number of observations.
<code>shape</code>	Shape parameter of the Fréchet distribution.
<code>loc</code>	Location parameter of the Fréchet distribution, default is 0.
<code>scale</code>	Scale parameter of the Fréchet distribution, default is 1.
<code>log</code>	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
<code>lower.tail</code>	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
<code>log.p</code>	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the Fréchet distribution is equal to  $F(x) = \exp(-((x - loc)/scale)^{-shape})$  for all  $x \geq loc$  and  $F(x) = 0$  otherwise. Both shape and scale need to be strictly positive.

**Value**

`dfrechet` gives the density function evaluated in  $x$ , `pfrechet` the CDF evaluated in  $x$  and `qfrechet` the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

`rfrechet` returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**See Also**

[tFréchet, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(1,10,0.01)
plot(x, dfrechet(x, shape=2), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(1,10,0.01)
plot(x, pfrechet(x, shape=2), xlab="x", ylab="CDF", type="l")
```

genHill

*Generalised Hill estimator***Description**

Computes the generalised Hill estimator for real extreme value indices as a function of the tail parameter  $k$ . Optionally, these estimates are plotted as a function of  $k$ .

**Usage**

```
genHill(data, gamma, logk = FALSE, plot = FALSE, add = FALSE,
        main = "Generalised Hill estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI, typically Hill estimates are used.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Generalised Hill estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The generalised Hill estimator is an estimator for the slope of the  $k$  last points of the generalised QQ-plot:

$$\hat{\gamma}_{k,n}^{GH} = 1/k \sum_{j=1}^k \log UH_{j,n} - \log UH_{k+1,n}$$

with  $UH_{j,n} = X_{n-j,n} H_{j,n}$  the UH scores and  $H_{j,n}$  the Hill estimates. This is analogous to the (ordinary) Hill estimator which is the estimator of the slope of the  $k$  last points of the Pareto QQ-plot when using constrained least squares.

See Section 4.2.2 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding generalised Hill estimates.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

Beirlant, J., Vynckier, P. and Teugels, J.L. (1996). "Excess Function and Estimation of the Extreme-value Index". *Bernoulli*, 2, 293–318.

**See Also**

[Hill](#), [genQQ](#), [Moment](#)

**Examples**

```
data(soa)

# Hill estimator
H <- Hill(soa$size, plot=FALSE)
# Moment estimator
M <- Moment(soa$size)
# Generalised Hill estimator
gH <- genHill(soa$size, gamma=H$gamma)

# Plot estimates
plot(H$k[1:5000], M$gamma[1:5000], xlab="k", ylab=expression(gamma), type="l", ylim=c(0.2,0.5))
lines(H$k[1:5000], gH$gamma[1:5000], lty=2)
legend("topright", c("Moment", "Generalised Hill"), lty=1:2)
```

---

genQQ

*Generalised quantile plot*

---

**Description**

Computes the empirical quantiles of the UH scores of a data vector and the theoretical quantiles of the standard exponential distribution. These quantiles are then plotted in a generalised QQ-plot with the theoretical quantiles on the  $x$ -axis and the empirical quantiles on the  $y$ -axis.

**Usage**

```
genQQ(data, gamma, plot = TRUE, main = "Generalised QQ-plot", ...)
```

```
generalizedQQ(data, gamma, plot = TRUE, main = "Generalised QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI, typically Hill estimates are used.
plot	Logical indicating if the quantiles should be plotted in a generalised QQ-plot, default is TRUE.
main	Title for the plot, default is "Generalised QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The generalizedQQ function is the same function but with a different name for compatibility with the old S-Plus code.

The UH scores are defined as  $UH_{j,n} = X_{n-j,n}H_{j,n}$  with  $H_{j,n}$  the Hill estimates, but other positive estimates for the EVI can also be used. The appropriate positive estimates for the EVI need to be specified in gamma. The generalised QQ-plot then plots

$$(\log((n+1)/(k+1)), \log(X_{n-k,n}H_{k,n}))$$

for  $k = 1, \dots, n - 1$ .

See Section 4.2.2 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

gqq.the	Vector of the theoretical quantiles from a standard exponential distribution.
gqq.emp	Vector of the empirical quantiles from the logarithm of the UH scores.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

Beirlant, J., Vynckier, P. and Teugels, J.L. (1996). "Excess Function and Estimation of the Extreme-value Index." *Bernoulli*, 2, 293–318.

**See Also**

[ParetoQQ](#), [Hill](#)

**Examples**

```

data(soa)

# Compute Hill estimator
H <- Hill(soa$size[1:5000], plot=FALSE)$gamma

# Generalised QQ-plot
genQQ(soa$size[1:5000], gamma=H)

```

GPD

*The generalised Pareto distribution***Description**

Density, distribution function, quantile function and random generation for the Generalised Pareto Distribution (GPD).

**Usage**

```

dgpd(x, gamma, mu = 0, sigma, log = FALSE)
pgpd(x, gamma, mu = 0, sigma, lower.tail = TRUE, log.p = FALSE)
qgpd(p, gamma, mu = 0, sigma, lower.tail = TRUE, log.p = FALSE)
rgpd(n, gamma, mu = 0, sigma)

```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
gamma	The $\gamma$ parameter of the GPD, a real number.
mu	The $\mu$ parameter of the GPD, a strictly positive number. Default is 0.
sigma	The $\sigma$ parameter of the GPD, a strictly positive number.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the GPD for  $\gamma \neq 0$  is equal to  $F(x) = 1 - (1 + \gamma(x - \mu)/\sigma)^{-1/\gamma}$  for all  $x \geq \mu$  and  $F(x) = 0$  otherwise. When  $\gamma = 0$ , the CDF is given by  $F(x) = 1 - \exp(-(x - \mu)/\sigma)$  for all  $x \geq \mu$  and  $F(x) = 0$  otherwise.

**Value**

dgpd gives the density function evaluated in  $x$ , pgpd the CDF evaluated in  $x$  and qgpd the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rgpd returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**References**

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[tGPD](#), [Pareto](#), [EPD](#), [Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dgpd(x, gamma=1/2, sigma=5), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, pgpd(x, gamma=1/2, sigma=5), xlab="x", ylab="CDF", type="l")
```

---

GPDfit

*Fit GPD using MLE*


---

**Description**

Fit the Generalised Pareto Distribution (GPD) to data using Maximum Likelihood Estimation (MLE).

**Usage**

```
GPDfit(data, start = c(0.1, 1), warnings = FALSE)
```

**Arguments**

data	Vector of $n$ observations.
start	Vector of length 2 containing the starting values for the optimisation. The first element is the starting value for the estimator of $\gamma$ and the second element is the starting value for the estimator of $\sigma$ . Default is $c(0.1, 1)$ .
warnings	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.

**Details**

See Section 4.2.2 in Albrecher et al. (2017) for more details.

**Value**

A vector with the MLE estimate for the  $\gamma$  parameter of the GPD as the first component and the MLE estimate for the  $\sigma$  parameter of the GPD as the second component.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur and R code from Klaus Herrmann.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[GPDmle](#), [EPDfit](#)

**Examples**

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Fit GPD to last 500 observations
res <- GPDfit(SOAdata=sort(soa$size)[500])
```

---

GPDmle

*GPD-ML estimator*

---

**Description**

Fit the Generalised Pareto Distribution (GPD) to the exceedances (peaks) over a threshold using Maximum Likelihood Estimation (MLE). Optionally, these estimates are plotted as a function of  $k$ .

**Usage**

```
GPDmle(data, start = c(0.1,1), warnings = FALSE, logk = FALSE,
        plot = FALSE, add = FALSE, main = "POT estimates of the EVI", ...)
```

```
POT(data, start = c(0.1,1), warnings = FALSE, logk = FALSE,
     plot = FALSE, add = FALSE, main = "POT estimates of the EVI", ...)
```



**Arguments**

data	Vector of $n$ observations.
start	Vector of length 2 containing the starting values for the optimisation. The first element is the starting value for the estimator of $\gamma$ and the second element is the starting value for the estimator of $\sigma$ . Default is $c(0.1, 1)$ .
warnings	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( $\text{logk}=\text{TRUE}$ ) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "POT estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The POT function is the same function but with a different name for compatibility with the old S-Plus code.

For each value of  $k$ , we look at the exceedances over the  $(k + 1)$ th largest observation:  $X_{n-k+j,n} - X_{n-k,n}$  for  $j = 1, \dots, k$ , with  $X_{j,n}$  the  $j$ th largest observation and  $n$  the sample size. The GPD is then fitted to these  $k$  exceedances using MLE which yields estimates for the parameters of the GPD:  $\gamma$  and  $\sigma$ .

See Section 4.2.2 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding MLE estimates for the $\gamma$ parameter of the GPD.
sigma	Vector of the corresponding MLE estimates for the $\sigma$ parameter of the GPD.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur and R code from Klaus Herrmann.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[GPDfit](#), [GPDresiduals](#), [EPD](#)

**Examples**

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Plot GPD-ML estimates as a function of k
GPDmle(SOAdata, plot=TRUE)
```

---

GPDresiduals

*GPD residual plot*

---

**Description**

Residual plot to check GPD fit for peaks over a threshold.

**Usage**

```
GPDresiduals(data, t, gamma, sigma, plot = TRUE,
             main = "GPD residual plot", ...)
```

**Arguments**

<code>data</code>	Vector of $n$ observations.
<code>t</code>	The used threshold.
<code>gamma</code>	Estimate for the EVI obtained from GPDmle.
<code>sigma</code>	Estimate for $\sigma$ obtained from GPDmle.
<code>plot</code>	Logical indicating if the residuals should be plotted, default is FALSE.
<code>main</code>	Title for the plot, default is "GPD residual plot".
<code>...</code>	Additional arguments for the <code>plot</code> function, see <a href="#">plot</a> for more details.

**Details**

Consider the POT values  $Y = X - t$  and the transformed variable

$$R = 1/\gamma \log(1 + \gamma/\sigma Y),$$

when  $\gamma \neq 0$  and

$$R = Y/\sigma,$$

otherwise. We can assess the goodness-of-fit of the GPD when modelling POT values  $Y = X - t$  by constructing an exponential QQ-plot of the transformed variable  $R$  since  $R$  is standard exponentially distributed if  $Y$  follows the GPD.

See Section 4.2.2 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

res.the            Vector of the theoretical quantiles from a standard exponential distribution.  
res.emp            Vector of the empirical quantiles of  $R$ , see Details.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[GPDfit](#), [ExpQQ](#)

**Examples**

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Plot POT-MLE estimates as a function of k
pot <- GPDmle(SOAdata, plot=TRUE)

# Residual plot
k <- 200
GPDresiduals(SOAdata, sort(SOAdata)[length(SOAdata)-k], pot$gamma[k], pot$sigma[k])
```

---

Hill

*Hill estimator*

---

**Description**

Computes the Hill estimator for positive extreme value indices (Hill, 1975) as a function of the tail parameter  $k$ . Optionally, these estimates are plotted as a function of  $k$ .

**Usage**

```
Hill(data, k = TRUE, logk = FALSE, plot = FALSE, add = FALSE,
      main = "Hill estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
k	Logical indicating if the Hill estimates are plotted as a function of the tail parameter $k$ (k=TRUE) or as a function of $\log(X_{n-k})$ . Default is TRUE.
logk	Logical indicating if the Hill estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ (logk=FALSE) when k=TRUE. Default is FALSE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Hill estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The Hill estimator can be seen as the estimator of slope in the upper right corner ( $k$  last points) of the Pareto QQ-plot when using constrained least squares (the regression line has to pass through the point  $(-\log((k+1)/(n+1)), \log X_{n-k})$ ). It is given by

$$H_{k,n} = 1/k \sum_{j=1}^k \log X_{n-j+1,n} - \log X_{n-k,n}.$$

See Section 4.2.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding Hill estimates.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.
- Hill, B. M. (1975). "A simple general approach to inference about the tail of a distribution." *Annals of Statistics*, 3, 1163–1173.

**See Also**

[ParetoQQ](#), [Hill.2oQV](#), [genHill](#)

**Examples**

```
data(norwegianfire)

# Plot Hill estimates as a function of k
Hill(norwegianfire$size[norwegianfire$year==76],plot=TRUE)
```

Hill.2oQV

*Bias-reduced MLE (Quantile view)***Description**

Computes bias-reduced ML estimates of gamma based on the quantile view.

**Usage**

```
Hill.2oQV(data, start = c(1,1,1), warnings = FALSE, logk = FALSE,
          plot = FALSE, add = FALSE, main = "Estimates of the EVI", ...)
```

**Arguments**

<code>data</code>	Vector of $n$ observations.
<code>start</code>	A vector of length 3 containing starting values for the first numerical optimisation (see Details). The elements are the starting values for the estimators of $\gamma$ , $\mu$ and $\sigma$ , respectively. Default is <code>c(1, 1, 1)</code> .
<code>warnings</code>	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.
<code>logk</code>	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is FALSE.
<code>plot</code>	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
<code>add</code>	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
<code>main</code>	Title for the plot, default is "Estimates of the EVI".
<code>...</code>	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

See Section 4.2.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the ML estimates for the EVI for each value of $k$ .
b	Vector of the ML estimates for the parameter $b$ in the regression model for each value of $k$ .
beta	Vector of the ML estimates for the parameter $\beta$ in the regression model for each value of $k$ .

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur and R code from Klaus Herrmann.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Dierckx, G., Goegebeur Y. and Matthys, G. (1999). "Tail Index Estimation and an Exponential Regression Model." *Extremes*, 2, 177–200.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**Examples**

```
data(norwegianfire)

# Plot bias-reduced MLE (QV) as a function of k
Hill.2oQV(norwegianfire$size[norwegianfire$year==76],plot=TRUE)
```

---

Hill.kopt

*Select optimal threshold for Hill estimator*


---

**Description**

Select optimal threshold for the Hill estimator by minimising the Asymptotic Mean Squared Error (AMSE).

**Usage**

```
Hill.kopt(data, start = c(1, 1, 1), warnings = FALSE, logk = FALSE,
          plot = FALSE, add = FALSE, main = "AMSE plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
start	A vector of length 3 containing starting values for the first numerical optimisation (see <a href="#">Hill.2oQV</a> for more details). Default is <code>c(1, 1, 1)</code> .
warnings	Logical indicating if possible warnings from the optimisation function are shown, default is FALSE.
logk	Logical indicating if the AMSE values are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the AMSE values should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the optimal value for $k$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "AMSE plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

See Section 4.2.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
AMSE	Vector of the AMSE values for each value of $k$ .
kopt	Optimal value of $k$ corresponding to minimal <i>AMSE</i> value.
gammaopt	Optimal value of the Hill estimator corresponding to minimal <i>AMSE</i> value.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.
- Beirlant J., Vynckier, P. and Teugels, J. (1996). "Tail Index Estimation, Pareto Quantile Plots, and Regression Diagnostics." *Journal of the American Statistical Association*, 91, 1659–1667.

**See Also**

[Hill](#), [Hill.2oQV](#)

**Examples**

```

data(norwegianfire)

# Plot Hill estimator as a function of k
Hill(norwegianfire$size[norwegianfire$year==76],plot=TRUE)

# Add optimal value of k
Hill.kopt(norwegianfire$size[norwegianfire$year==76],add=TRUE)

```

icHill

*Hill estimator for interval censored data***Description**

Computes the Hill estimator for positive extreme value indices, adapted for interval censoring, as a function of the tail parameter  $k$ . Optionally, these estimates are plotted as a function of  $k$ .

**Usage**

```

icHill(L, U, censored, trunclower = 0, truncupper = Inf,
       logk = FALSE, plot = TRUE, add = FALSE, main = "Hill estimates of the EVI", ...)

```

**Arguments**

L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals.
censored	A logical vector of length $n$ indicating if an observation is censored.
trunclower	Lower truncation point. Default is 0.
truncupper	Upper truncation point. Default is Inf (no upper truncation).
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Hill estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.



## Details

This estimator is given by

$$H^{TB}(x) = \left( \int_x^\infty (1 - \hat{F}^{TB}(u))/u du \right) / (1 - \hat{F}^{TB}(x)),$$

where  $\hat{F}^{TB}$  is the Turnbull estimator for the CDF. More specifically, we use the values  $x = \hat{Q}^{TB}(p)$  for  $p = 1/(n+1), \dots, (n-1)/(n+1)$  where  $\hat{Q}^{TB}(p)$  is the empirical quantile function corresponding to the Turnbull estimator. We then denote

$$H_{k,n}^{TB} = H^{TB}(x_{n-k,n})$$

with

$$x_{n-k,n} = \hat{Q}^{TB}((n-k)/(n+1)) = \hat{Q}^{TB}(1 - (k+1)/(n+1)).$$

Right censored data should be entered as L=1 and U=truncupper, and left censored data should be entered as L=trunclower and U=u.

If the **interval** package is installed, the `icfit` function is used to compute the Turnbull estimator. Otherwise, `survfit.formula` from **survival** is used.

Use `Hill` for non-censored data or `cHill` for right censored data.

See Section 4.3 in Albrecher et al. (2017) for more details.

## Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding Hill estimates.
X	Vector of thresholds $x_{n-k,n}$ used when estimating $\gamma$ .

## Author(s)

Tom Reynkens

## References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

## See Also

`cHill`, `Hill`, `MeanExcess_TB`, `icParetoQQ`, `Turnbull`, `icfit`

**Examples**

```

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Hill estimator adapted for interval censoring
icHill(Z, U, censored, ylim=c(0,1))

# Hill estimator adapted for right censoring
cHill(Z, censored, lty=2, add=TRUE)

# True value of gamma
abline(h=1/2, lty=3, col="blue")

# Legend
legend("topright", c("icHill", "cHill"), lty=1:2)

```

---

icParetoQQ

*Pareto quantile plot for interval censored data*


---

**Description**

Pareto QQ-plot adapted for interval censored data using the Turnbull estimator.

**Usage**

```
icParetoQQ(L, U = L, censored, trunclower = 0, truncupper = Inf,
           plot = TRUE, main = "Pareto QQ-plot", ...)
```

**Arguments**

L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals.
censored	A logical vector of length $n$ indicating if an observation is censored.
trunclower	Lower truncation point. Default is 0.

truncupper	Upper truncation point. Default is Inf (no upper truncation).
plot	Logical indicating if the quantiles should be plotted in a Pareto QQ-plot, default is TRUE.
main	Title for the plot, default is "Pareto QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The Pareto QQ-plot adapted for interval censoring is given by

$$(-\log(1 - F^{TB}(x_{j,n})), \log x_{j,n})$$

for  $j = 1, \dots, n-1$ , where  $\hat{F}^{TB}$  is the Turnbull estimator for the CDF and  $x_{i,n} = \hat{Q}^{TB}(i/(n+1))$  with  $\hat{Q}^{TB}(p)$  the empirical quantile function corresponding to the Turnbull estimator.

Right censored data should be entered as L=1 and U=truncupper, and left censored data should be entered as L=trunclower and U=u.

If the **interval** package is installed, the [icfit](#) function is used to compute the Turnbull estimator. Otherwise, [survfit.formula](#) from **survival** is used.

Use [ParetoQQ](#) for non-censored data or [cParetoQQ](#) for right censored data.

See Section 4.3 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

pqq.the	Vector of the theoretical quantiles, see Details.
pqq.emp	Vector of the empirical quantiles from the log-transformed data.

### Author(s)

Tom Reynkens

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

### See Also

[cParetoQQ](#), [ParetoQQ](#), [icHill](#), [Turnbull](#), [icfit](#)

### Examples

```
# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)
```

```

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Pareto QQ-plot adapted for interval censoring
icParetoQQ(Z, U, censored)

# Pareto QQ-plot adapted for right censoring
cParetoQQ(Z, censored)

```

---

KaplanMeier

*Kaplan-Meier estimator*


---

## Description

Computes the Kaplan-Meier estimator for the survival function of right censored data.

## Usage

```
KaplanMeier(x, data, censored, conf.type="plain", conf.int = 0.95)
```

## Arguments

<code>x</code>	Vector with points to evaluate the estimator in.
<code>data</code>	Vector of $n$ observations.
<code>censored</code>	Vector of $n$ logicals indicating if an observation is right censored.
<code>conf.type</code>	Type of confidence interval, see <a href="#">survfit.formula</a> . Default is "plain".
<code>conf.int</code>	Confidence level of the two-sided confidence interval, see <a href="#">survfit.formula</a> . Default is 0.95.

## Details

We consider the random right censoring model where one observes  $Z = \min(X, C)$  where  $X$  is the variable of interest and  $C$  is the censoring variable.

This function is merely a wrapper for [survfit.formula](#) from **survival**.

This estimator is only suitable for right censored data. When the data are interval censored, one can use the Turnbull estimator implemented in [Turnbull](#).

**Value**

A list with following components:

<code>surv</code>	A vector of length <code>length(x)</code> containing the Kaplan-Meier estimator evaluated in the elements of <code>x</code> .
<code>fit</code>	The output from the call to <code>survfit.formula</code> , an object of class <code>survfit</code> .

**Author(s)**

Tom Reynkens

**References**

Kaplan, E. L. and Meier, P. (1958). "Nonparametric Estimation from Incomplete Observations." *Journal of the American Statistical Association*, 53, 457–481.

**See Also**

[survfit.formula](#), [Turnbull](#)

**Examples**

```
data <- c(1, 2.5, 3, 4, 5.5, 6, 7.5, 8.25, 9, 10.5)
censored <- c(0, 1, 0, 0, 1, 0, 1, 1, 0, 1)

x <- seq(0, 12, 0.1)

# Kaplan-Meier estimator
plot(x, KaplanMeier(x, data, censored)$surv, type="s", ylab="Kaplan-Meier estimator")
```

---

LognormalQQ

*Log-normal quantile plot*

---

**Description**

Computes the empirical quantiles of the log-transform of a data vector and the theoretical quantiles of the standard normal distribution. These quantiles are then plotted in a log-normal QQ-plot with the theoretical quantiles on the *x*-axis and the empirical quantiles on the *y*-axis.

**Usage**

```
LognormalQQ(data, plot = TRUE, main = "Log-normal QQ-plot", ...)
```

**Arguments**

<code>data</code>	Vector of $n$ observations.
<code>plot</code>	Logical indicating if the quantiles should be plotted in a log-normal QQ-plot, default is TRUE.
<code>main</code>	Title for the plot, default is "Log-normal QQ-plot".
<code>...</code>	Additional arguments for the <code>plot</code> function, see <a href="#">plot</a> for more details.

**Details**

By definition, a log-transformed log-normal random variable is normally distributed. We can thus obtain a log-normal QQ-plot from a normal QQ-plot by replacing the empirical quantiles of the data vector by the empirical quantiles from the log-transformed data. We hence plot

$$(\Phi^{-1}(i/(n+1)), \log(X_{i,n}))$$

for  $i = 1, \dots, n$ , where  $\Phi$  is the standard normal CDF.

See Section 4.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

<code>lnqq.the</code>	Vector of the theoretical quantiles from a standard normal distribution.
<code>lnqq.emp</code>	Vector of the empirical quantiles from the log-transformed data.

**Author(s)**

Tom Reynkens.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[ExpQQ](#), [ParetoQQ](#), [WeibullQQ](#)

**Examples**

```
data(norwegianfire)

# Log-normal QQ-plot for Norwegian Fire Insurance data for claims in 1976.
LognormalQQ(norwegianfire$size[norwegianfire$year==76])
```

---

LognormalQQ\_der      *Derivative plot of the log-normal QQ-plot*

---

### Description

Computes the derivative plot of the log-normal QQ-plot. These values can be plotted as a function of the data or as a function of the tail parameter  $k$ .

### Usage

```
LognormalQQ_der(data, k = FALSE, plot = TRUE,
                 main = "Derivative plot of log-normal QQ-plot", ...)
```

### Arguments

data	Vector of $n$ observations.
plot	Logical indicating if the derivative values should be plotted, default is TRUE.
k	Logical indicating if the derivative values are plotted as a function of the tail parameter $k$ ( $k=TRUE$ ) or as a function of the logarithm of the data ( $k=FALSE$ ). Default is FALSE.
main	Title for the plot, default is "Derivative plot of log-normal QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The derivative plot of a log-normal QQ-plot is

$$(k, H_{k,n}/N_{k,n})$$

or

$$(\log X_{n-k,n}, H_{k,n}/N_{k,n})$$

with  $H_{k,n}$  the Hill estimates and

$$N_{k,n} = (n + 1)/(k + 1)\phi(\Phi^{-1}(a)) - \Phi^{-1}(a).$$

Here is  $a = 1 - (k + 1)/(n + 1)$ ,  $\phi$  the standard normal PDF and  $\Phi$  the standard normal CDF.

See Section 4.1 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

xval	Vector of the x-values of the plot ( $k$ or $\log X_{n-k,n}$ ).
yval	Vector of the derivative values.

**Author(s)**

Tom Reynkens.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[LognormalQQ](#), [Hill](#), [MeanExcess](#), [ParetoQQ\\_der](#), [WeibullQQ\\_der](#)

**Examples**

```
data(norwegianfire)

# Log-normal QQ-plot for Norwegian Fire Insurance data for claims in 1976.
LognormalQQ(norwegianfire$size[norwegianfire$year==76])

# Derivate plot
LognormalQQ_der(norwegianfire$size[norwegianfire$year==76])
```

---

LStail

*Least Squares tail estimator*


---

**Description**

Computes the Least Squares (LS) estimates of the EVI based on the last  $k$  observations of the generalised QQ-plot.

**Usage**

```
LStail(data, rho = -1, lambda = 0.5, logk = FALSE, plot = FALSE, add = FALSE,
       main = "LS estimates of the EVI", ...)

TSfraction(data, rho = -1, lambda = 0.5, logk = FALSE, plot = FALSE, add = FALSE,
          main = "LS estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
rho	Estimate for $\rho$ , or NULL when $\rho$ needs to be estimated using the method of Beirlant et al. (2002). Default is -1.
lambda	Parameter used in the method of Beirlant et al. (2002), only used when rho=NULL. Default is 0.5.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.



plot	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "LS estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

We estimate  $\gamma$  (EVI) and  $b$  using least squares on the following regression model (Beirlant et al., 2005):  $Z_j = \gamma + b(n/k)(j/k)^{-\rho} + \epsilon_j$  with  $Z_j = (j + 1) \log(UH_{j,n}/UH_{j+1,n})$  and  $UH_{j,n} = X_{n-j,n}H_{j,n}$ , where  $H_{j,n}$  is the Hill estimator with threshold  $X_{n-j,n}$ .

See Section 5.8 of Beirlant et al. (2004) for more details.

The function TSfraction is included for compatibility with the old S-Plus code.

### Value

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding LS estimates for the EVI.
b	Vector of the corresponding LS estimates for $b$ .
rho	Vector of the estimates for $\rho$ when rho=NULL or the given input for rho otherwise.

### Author(s)

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

### References

- Beirlant, J., Dierckx, G. and Guillou, A. (2005). "Estimation of the Extreme Value Index and Regression on Generalized Quantile Plots." *Bernoulli*, 11, 949–970.
- Beirlant, J., Dierckx, G., Guillou, A. and Starica, C. (2002). "On Exponential Representations of Log-spacing of Extreme Order Statistics." *Extremes*, 5, 157–180.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

### See Also

[genQQ](#)

### Examples

```
data(soa)

# LS tail estimator
LStail(soa$size, plot=TRUE, ylim=c(0,0.5))
```

---

MeanExcess	<i>Mean excess function</i>
------------	-----------------------------

---

### Description

Computes the mean excess values for a vector of observations. These mean excess values can then be plotted as a function of the data or as a function of the tail parameter  $k$ .

### Usage

```
MeanExcess(data, plot = TRUE, k = FALSE, main = "Mean excess plot", ...)
```

### Arguments

data	Vector of $n$ observations.
plot	Logical indicating if the mean excess values should be plotted in a mean excess plot, default is TRUE.
k	Logical indicating if the mean excess scores are plotted as a function of the tail parameter $k$ ( $k=TRUE$ ) or as a function of the data ( $k=FALSE$ ). Default is FALSE.
main	Title for the plot, default is "Mean excess plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The mean excess plot is

$$(k, e_{k,n})$$

or

$$(X_{n-k,n}, e_{k,n})$$

with

$$e_{k,n} = 1/k \sum_{j=1}^k X_{n-j+1,n} - X_{n-k,n}.$$

Note that the mean excess plot is the derivative plot of the Exponential QQ-plot.

See Section 4.1 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
X	Vector of the order statistics <code>data[n-k]</code> corresponding to the tail parameters in $k$ .
e	Vector of the mean excess values corresponding to the tail parameters in $k$ .

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[ExpQQ](#), [LognormalQQ\\_der](#), [ParetoQQ\\_der](#), [WeibullQQ\\_der](#)

**Examples**

```
data(norwegianfire)

# Mean excess plots for Norwegian Fire Insurance data for claims in 1976.

# Mean excess values as a function of k
MeanExcess(norwegianfire$size[norwegianfire$year==76], k=TRUE)

# Mean excess values as a function of the data
MeanExcess(norwegianfire$size[norwegianfire$year==76], k=FALSE)
```

---

MeanExcess\_TB

*Mean excess function using Turnbull estimator*

---

**Description**

Computes mean excess values using the Turnbull estimator. These mean excess values can then be plotted as a function of the empirical quantiles (computed using the Turnbull estimator) or as a function of the tail parameter  $k$ .

**Usage**

```
MeanExcess_TB(L, U = L, censored, trunclower = 0, truncupper = Inf,
              plot = TRUE, k = FALSE, intervalpkg = TRUE,
              main = "Mean excess plot", ...)
```

**Arguments**

L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals. By default, they are equal to L.
censored	A logical vector of length $n$ indicating if an observation is censored.
trunclower	Lower truncation point, default is 0.
truncupper	Upper truncation point, default is Inf.
plot	Logical indicating if the mean excess values should be plotted in a mean excess plot, default is TRUE.
k	Logical indicating if the mean excess values are plotted as a function of the tail parameter $k$ ( $k=TRUE$ ) or as a function of the empirical quantiles computed using the Turnbull estimator ( $k=FALSE$ ). Default is FALSE.
intervalpkg	Logical indicating if the Turnbull estimator is computed using the implementation in the <b>interval</b> package if this package is installed. Default is TRUE.
main	Title for the plot, default is "Mean excess plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The mean excess values are given by

$$\hat{e}^{TB}(v) = \left( \int_v^{\infty} 1 - \hat{F}^{TB}(u) du \right) / (1 - \hat{F}^{TB}(v))$$

where  $\hat{F}^{TB}$  is the Turnbull estimator for the CDF. More specifically, we use the values  $v = \hat{Q}^{TB}(p)$  for  $p = 1/(n+1), \dots, (n-1)/(n+1)$  where  $\hat{Q}^{TB}(p)$  is the empirical quantile function corresponding to the Turnbull estimator.

Right censored data should be entered as  $L=1$  and  $U=truncupper$ , and left censored data should be entered as  $L=trunclower$  and  $U=u$ .

If the **interval** package is installed and `intervalpkg=TRUE`, the `icfit` function is used to compute the Turnbull estimator. Otherwise, `survfit.formula` from **survival** is used.

Use `MeanExcess` for non-censored data.

See Section 4.3 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
X	Vector of the empirical quantiles, computed using the Turnbull estimator, corresponding to $(n-k)/(n+1)=1-(k+1)/(n+1)$ .
e	Vector of the mean excess values corresponding to the tail parameters in $k$ .

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[MeanExcess](#), [Turnbull](#), [icfit](#)

**Examples**

```
# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X, Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Mean excess plot
MeanExcess_TB(Z, U, censored, k=FALSE)
```

---

MEfit

*Mixed Erlang fit*

---

**Description**

Create an S3 object using a Mixed Erlang (ME) fit.

**Usage**

```
MEfit(p, shape, theta, M, M_initial = NULL)
```

**Arguments**

p	Vector of mixing weights, denoted by $\alpha$ in Verbelen et al. (2015).
shape	Vector of shape parameters $r$ .
theta	Scale parameter $\theta$ .
M	Number of mixture components.
M_initial	Initial value provided for M. When NULL (default), not included in the object.

**Details**

The rate parameter  $\lambda$  used in Albrecher et al. (2017) is equal to  $1/\theta$ .

See Reynkens et al. (2017) and Section 4.3 of Albrecher et al. (2017) for more details

**Value**

An S3 object which contains the input arguments in a list.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[SpliceFit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#)

**Examples**

```
# Create MEfit object
mefit <- MEfit(p=c(0.65,0.35), shape=c(39,58), theta=16.19, M=2)

# Create EVTfit object
evtfit <- EVTfit(gamma=c(0.76,0.64), endpoint=c(39096, Inf))

# Create SpliceFit object
splicefit <- SpliceFit(const=c(0.5,0.996), trunclower=0, t=c(1020,39096), type=c("ME", "TPa", "Pa"),
                     MEfit=mefit, EVTfit=evtfit)

# Show summary
summary(splicefit)
```

---

Moment	<i>Moment estimator</i>
--------	-------------------------

---

### Description

Compute the moment estimates for real extreme value indices as a function of the tail parameter  $k$ . Optionally, these estimates are plotted as a function of  $k$ .

### Usage

```
Moment(data, logk = FALSE, plot = FALSE, add = FALSE,  
        main = "Moment estimates of the EVI", ...)
```

### Arguments

<code>data</code>	Vector of $n$ observations.
<code>logk</code>	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is <code>FALSE</code> .
<code>plot</code>	Logical indicating if the estimates should be plotted as a function of $k$ , default is <code>FALSE</code> .
<code>add</code>	Logical indicating if the estimates should be added to an existing plot, default is <code>FALSE</code> .
<code>main</code>	Title for the plot, default is "Moment estimates of the EVI".
<code>...</code>	Additional arguments for the <code>plot</code> function, see <a href="#">plot</a> for more details.

### Details

The moment estimator for the EVI is introduced by Dekkers et al. (1989) and is a generalisation of the Hill estimator.

See Section 4.2.2 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

<code>k</code>	Vector of the values of the tail parameter $k$ .
<code>gamma</code>	Vector of the corresponding moment estimates.

### Author(s)

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

## References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.
- Dekkers, A.L.M, Einmahl, J.H.J. and de Haan, L. (1989). "A Moment Estimator for the Index of an Extreme-value Distribution." *Annals of Statistics*, 17, 1833–1855.

## See Also

[Hill](#), [genHill](#)

## Examples

```
data(soa)

# Hill estimator
H <- Hill(soa$size, plot=FALSE)
# Moment estimator
M <- Moment(soa$size)
# Generalised Hill estimator
gH <- genHill(soa$size, gamma=H$gamma)

# Plot estimates
plot(H$k[1:5000], M$gamma[1:5000], xlab="k", ylab=expression(gamma), type="l", ylim=c(0.2,0.5))
lines(H$k[1:5000], gH$gamma[1:5000], lty=2)
legend("topright", c("Moment", "Generalised Hill"), lty=1:2)
```

---

norwegianfire

*Norwegian fire insurance data*

---

## Description

Fire insurance claims for a Norwegian insurance company for the period 1972 to 1992 as studied in Beirlant et al. (1996).

A priority of 500 units was in force.

## Usage

```
data("norwegianfire")
```

## Format

A data frame with 9181 observations on the following 2 variables:

size Size of fire insurance claim (in 1000 NOK).

year Year of claim occurrence (expressed as *yy* instead of *19yy*).



**Source**

Beirlant, J., Teugels, J. L. and Vynckier, P. (1996). *Practical Analysis of Extreme Values*, Leuven University Press.

**References**

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**Examples**

```
data(norwegianfire)

# Exponential QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ExpQQ(norwegianfire$size[norwegianfire$year==76])

# Pareto QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ParetoQQ(norwegianfire$size[norwegianfire$year==76])
```

---

Pareto

*The Pareto distribution*


---

**Description**

Density, distribution function, quantile function and random generation for the Pareto distribution (type I).

**Usage**

```
dpareto(x, shape, scale = 1, log = FALSE)
ppareto(x, shape, scale = 1, lower.tail = TRUE, log.p = FALSE)
qpareto(p, shape, scale = 1, lower.tail = TRUE, log.p = FALSE)
rpareto(n, shape, scale = 1)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
shape	The shape parameter of the Pareto distribution, a strictly positive number.
scale	The scale parameter of the Pareto distribution, a strictly positive number. Its default value is 1.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the Pareto distribution is equal to  $F(x) = 1 - (x/scale)^{-shape}$  for all  $x \geq scale$  and  $F(x) = 0$  otherwise. Both shape and scale need to be strictly positive.

**Value**

dpareto gives the density function evaluated in  $x$ , ppareto the CDF evaluated in  $x$  and qpareto the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rpareto returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**See Also**

[tPareto](#), [GPD](#), [Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(1, 10, 0.01)
plot(x, dpareto(x, shape=2), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(1, 10, 0.01)
plot(x, ppareto(x, shape=2), xlab="x", ylab="CDF", type="l")
```

---

ParetoQQ

*Pareto quantile plot*

---

**Description**

Computes the empirical quantiles of the log-transform of a data vector and the theoretical quantiles of the standard exponential distribution. These quantiles are then plotted in a Pareto QQ-plot with the theoretical quantiles on the  $x$ -axis and the empirical quantiles on the  $y$ -axis.

**Usage**

```
ParetoQQ(data, plot = TRUE, main = "Pareto QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
plot	Logical indicating if the quantiles should be plotted in a Pareto QQ-plot, default is TRUE.
main	Title for the plot, default is "Pareto QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

It can be easily seen that a log-transformed Pareto random variable is exponentially distributed. We can hence obtain a Pareto QQ-plot from an exponential QQ-plot by replacing the empirical quantiles from the data vector by the empirical quantiles from the log-transformed data. We hence plot

$$(-\log(1 - i/(n + 1)), \log X_{i,n})$$

for  $i = 1, \dots, n$ , with  $X_{i,n}$  the  $i$ -th order statistic of the data.

See Section 4.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

pqq.the	Vector of the theoretical quantiles from a standard exponential distribution.
pqq.emp	Vector of the empirical quantiles from the log-transformed data.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[ParetoQQ\\_der](#), [ExpQQ](#), [genQQ](#), [LognormalQQ](#), [WeibullQQ](#)

**Examples**

```
data(norwegianfire)

# Exponential QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ExpQQ(norwegianfire$size[norwegianfire$year==76])

# Pareto QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ParetoQQ(norwegianfire$size[norwegianfire$year==76])
```

ParetoQQ\_der

*Derivative plot of the Pareto QQ-plot***Description**

Computes the derivative plot of the Pareto QQ-plot. These values can be plotted as a function of the data or as a function of the tail parameter  $k$ .

**Usage**

```
ParetoQQ_der(data, k = FALSE, plot = TRUE,
             main = "Derivative plot of Pareto QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
plot	Logical indicating if the derivative values should be plotted, default is TRUE.
k	Logical indicating if the derivative values are plotted as a function of the tail parameter $k$ (k=TRUE) or as a function of the logarithm of the data (k=FALSE). Default is FALSE.
main	Title for the plot, default is "Derivative plot of Pareto QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The derivative plot of a Pareto QQ-plot is

$$(k, H_{k,n})$$

or

$$(\log X_{n-k,n}, H_{k,n})$$

with  $H_{k,n}$  the Hill estimates.

See Section 4.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

xval	Vector of the x-values of the plot ( $k$ or $\log X_{n-k,n}$ ).
yval	Vector of the derivative values $H_{k,n}$ .

**Author(s)**

Tom Reynkens.

## References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

## See Also

[ParetoQQ](#), [Hill](#), [MeanExcess](#), [LognormalQQ\\_der](#), [WeibullQQ\\_der](#)

## Examples

```
data(norwegianfire)

# Pareto QQ-plot for Norwegian Fire Insurance data for claims in 1976.
ParetoQQ(norwegianfire$size[norwegianfire$year==76])

# Derivate plot
ParetoQQ_der(norwegianfire$size[norwegianfire$year==76])
```

---

pClas

*Classical estimators for the CDF*

---

## Description

Compute approximations of the CDF using the normal approximation, normal-power approximation, shifted Gamma approximation or normal approximation to the shifted Gamma distribution.

## Usage

```
pClas(x, mean = 0, variance = 1, skewness = NULL,
      method = c("normal", "normal-power", "shifted Gamma", "shifted Gamma normal"),
      lower.tail = TRUE, log.p = FALSE)
```

## Arguments

x	Vector of points to approximate the CDF in.
mean	Mean of the distribution, default is 0.
variance	Variance of the distribution, default is 1.
skewness	Skewness coefficient of the distribution, this argument is not used for the normal approximation. Default is NULL meaning no skewness coefficient is provided.
method	Approximation method to use, one of "normal", "normal-power", "shifted Gamma" or "shifted Gamma normal". Default is "normal".
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

- The *normal approximation* for the CDF of the r.v.  $X$  is defined as

$$F_X(x) \approx \Phi((x - \mu)/\sigma)$$

where  $\mu$  and  $\sigma^2$  are the mean and variance of  $X$ , respectively.

- This approximation can be improved when the skewness parameter

$$\nu = E((X - \mu)^3)/\sigma^3$$

is available. The *normal-power approximation* of the CDF is then given by

$$F_X(x) \approx \Phi(\sqrt{9/\nu^2 + 6z/\nu + 1} - 3/\nu)$$

for  $z = (x - \mu)/\sigma \geq 1$  and  $9/\nu^2 + 6z/\nu + 1 \geq 0$ .

- The *shifted Gamma approximation* uses the approximation

$$X \approx \Gamma(4/\nu^2, 2/(\nu \times \sigma)) + \mu - 2\sigma/\nu.$$

Here, we need that  $\nu > 0$ .

- The *normal approximation to the shifted Gamma distribution* approximates the CDF of  $X$  as

$$F_X(x) \approx \Phi(\sqrt{16/\nu^2 + 8z/\nu} - \sqrt{16/\nu^2 - 1})$$

for  $z = (x - \mu)/\sigma \geq 1$ . We need again that  $\nu > 0$ .

See Section 6.2 of Albrecher et al. (2017) for more details.

**Value**

Vector of estimates for the probabilities  $F(x) = P(X \leq x)$ .

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[pEdge](#), [pGC](#)

**Examples**

```

# Chi-squared sample
X <- rchisq(1000, 2)

x <- seq(0, 10, 0.01)

# Classical approximations
p1 <- pClas(x, mean(X), var(X))
p2 <- pClas(x, mean(X), var(X), mean((X-mean(X))^3)/sd(X)^3, method="normal-power")
p3 <- pClas(x, mean(X), var(X), mean((X-mean(X))^3)/sd(X)^3, method="shifted Gamma")
p4 <- pClas(x, mean(X), var(X), mean((X-mean(X))^3)/sd(X)^3, method="shifted Gamma normal")

# True probabilities
p <- pchisq(x, 2)

# Plot true and estimated probabilities
plot(x, p, type="l", ylab="F(x)", ylim=c(0,1), col="red")
lines(x, p1, lty=2)
lines(x, p2, lty=3, col="green")
lines(x, p3, lty=4)
lines(x, p4, lty=5, col="blue")

legend("bottomright", c("True CDF", "normal approximation", "normal-power approximation",
                        "shifted Gamma approximation", "shifted Gamma normal approximation"),
      lty=1:5, col=c("red", "black", "green", "black", "blue"), lwd=2)

```

pEdge

*Edgeworth approximation***Description**

Edgeworth approximation of the CDF using the first four moments.

**Usage**

```
pEdge(x, moments = c(0, 1, 0, 3), raw = TRUE, lower.tail = TRUE, log.p = FALSE)
```

**Arguments**

x	Vector of points to approximate the CDF in.
moments	The first four raw moments if raw=TRUE. By default the first four raw moments of the standard normal distribution are used. When raw=FALSE, the mean $\mu = E(X)$ , variance $\sigma^2 = E((X - \mu)^2)$ , skewness (third standardised moment, $\nu = E((X - \mu)^3)/\sigma^3$ ) and kurtosis (fourth standardised moment, $k = E((X - \mu)^4)/\sigma^4$ ).
raw	When TRUE (default), the first four raw moments are provided in moments. Otherwise, the mean, variance, skewness and kurtosis are provided in moments.

lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

### Details

Denote the standard normal PDF and CDF respectively by  $\phi$  and  $\Phi$ . Let  $\mu$  be the first moment,  $\sigma^2 = E((X - \mu)^2)$  the variance,  $\mu_3 = E((X - \mu)^3)$  the third central moment and  $\mu_4 = E((X - \mu)^4)$  the fourth central moment of the random variable  $X$ . The corresponding cumulants are given by  $\kappa_1 = \mu$ ,  $\kappa_2 = \sigma^2$ ,  $\kappa_3 = \mu_3$  and  $\kappa_4 = \mu_4 - 3\sigma^4$ .

Now consider the random variable  $Z = (X - \mu)/\sigma$ , which has cumulants 0, 1,  $\nu = \kappa_3/\sigma^3$  and  $k = \kappa_4/\sigma^4 = \mu_4/\sigma^4 - 3$ .

The Edgeworth approximation for the CDF of  $X$  ( $F(x)$ ) is given by

$$\hat{F}_E(x) = \Phi(z) + \phi(z)(-\nu/6h_2(z) - (3k \times h_3(z) + \gamma_3^2 h_5(z))/72)$$

with  $h_2(z) = z^2 - 1$ ,  $h_3(z) = z^3 - 3z$ ,  $h_5(z) = z^5 - 10z^3 + 15z$  and  $z = (x - \mu)/\sigma$ .

See Section 6.2 of Albrecher et al. (2017) for more details.

### Value

Vector of estimates for the probabilities  $F(x) = P(X \leq x)$ .

### Author(s)

Tom Reynkens

### References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Cheah, P.K., Fraser, D.A.S. and Reid, N. (1993). "Some Alternatives to Edgeworth." *The Canadian Journal of Statistics*, 21(2), 131–138.

### See Also

[pGC](#), [pEdge](#)

### Examples

```
# Chi-squared sample
X <- rchisq(1000, 2)

x <- seq(0, 10, 0.01)

# Empirical moments
moments = c(mean(X), mean(X^2), mean(X^3), mean(X^4))
```



```

# Gram-Charlier approximation
p1 <- pGC(x, moments)

# Edgeworth approximation
p2 <- pEdge(x, moments)

# Normal approximation
p3 <- pClas(x, mean(X), var(X))

# True probabilities
p <- pchisq(x, 2)

# Plot true and estimated probabilities
plot(x, p, type="l", ylab="F(x)", ylim=c(0,1), col="red")
lines(x, p1, lty=2)
lines(x, p2, lty=3)
lines(x, p3, lty=4, col="blue")

legend("bottomright", c("True CDF", "GC approximation",
                        "Edgeworth approximation", "Normal approximation"),
      col=c("red", "black", "black", "blue"), lty=1:4, lwd=2)

```

pGC

*Gram-Charlier approximation***Description**

Gram-Charlier approximation of the CDF using the first four moments.

**Usage**

```
pGC(x, moments = c(0, 1, 0, 3), raw = TRUE, lower.tail = TRUE, log.p = FALSE)
```

**Arguments**

x	Vector of points to approximate the CDF in.
moments	The first four raw moments if raw=TRUE. By default the first four raw moments of the standard normal distribution are used. When raw=FALSE, the mean $\mu = E(X)$ , variance $\sigma^2 = E((X - \mu)^2)$ , skewness (third standardised moment, $\nu = E((X - \mu)^3)/\sigma^3$ ) and kurtosis (fourth standardised moment, $k = E((X - \mu)^4)/\sigma^4$ ).
raw	When TRUE (default), the first four raw moments are provided in moments. Otherwise, the mean, variance, skewness and kurtosis are provided in moments.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

### Details

Denote the standard normal PDF and CDF respectively by  $\phi$  and  $\Phi$ . Let  $\mu$  be the first moment,  $\sigma^2 = E((X - \mu)^2)$  the variance,  $\mu_3 = E((X - \mu)^3)$  the third central moment and  $\mu_4 = E((X - \mu)^4)$  the fourth central moment of the random variable  $X$ . The corresponding cumulants are given by  $\kappa_1 = \mu$ ,  $\kappa_2 = \sigma^2$ ,  $\kappa_3 = \mu_3$  and  $\kappa_4 = \mu_4 - 3\sigma^4$ .

Now consider the random variable  $Z = (X - \mu)/\sigma$ , which has cumulants 0, 1,  $\nu = \kappa_3/\sigma^3$  and  $k = \kappa_4/\sigma^4 = \mu_4/\sigma^4 - 3$ .

The Gram-Charlier approximation for the CDF of  $X$  ( $F(x)$ ) is given by

$$\hat{F}_{GC}(x) = \Phi(z) + \phi(z)(-\nu/6h_2(z) - k/24h_3(z))$$

with  $h_2(z) = z^2 - 1$ ,  $h_3(z) = z^3 - 3z$  and  $z = (x - \mu)/\sigma$ .

See Section 6.2 of Albrecher et al. (2017) for more details.

### Value

Vector of estimates for the probabilities  $F(x) = P(X \leq x)$ .

### Author(s)

Tom Reynkens

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Cheah, P.K., Fraser, D.A.S. and Reid, N. (1993). "Some Alternatives to Edgeworth." *The Canadian Journal of Statistics*, 21(2), 131–138.

### See Also

[pEdge](#), [pClas](#)

### Examples

```
# Chi-squared sample
X <- rchisq(1000, 2)

x <- seq(0, 10, 0.01)

# Empirical moments
moments = c(mean(X), mean(X^2), mean(X^3), mean(X^4))

# Gram-Charlier approximation
p1 <- pGC(x, moments)

# Edgeworth approximation
p2 <- pEdge(x, moments)
```

```

# Normal approximation
p3 <- pClas(x, mean(X), var(X))

# True probabilities
p <- pchisq(x, 2)

# Plot true and estimated probabilities
plot(x, p, type="l", ylab="F(x)", ylim=c(0,1), col="red")
lines(x, p1, lty=2)
lines(x, p2, lty=3)
lines(x, p3, lty=4, col="blue")

legend("bottomright", c("True CDF", "GC approximation",
                        "Edgeworth approximation", "Normal approximation"),
      col=c("red", "black", "black", "blue"), lty=1:4, lwd=2)

```

---

Prob	<i>Weissman estimator of small exceedance probabilities and large return periods</i>
------	--

---

### Description

Compute estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the approach of Weissman (1978).

### Usage

```

Prob(data, gamma, q, plot = FALSE, add = FALSE,
     main = "Estimates of small exceedance probability", ...)

```

```

Return(data, gamma, q, plot = FALSE, add = FALSE,
       main = "Estimates of large return period", ...)

```

```

Weissman.p(data, gamma, q, plot = FALSE, add = FALSE,
           main = "Estimates of small exceedance probability", ...)

```

```

Weissman.r(data, gamma, q, plot = FALSE, add = FALSE,
           main = "Estimates of large return period", ...)

```

### Arguments

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI, typically Hill estimates are used.
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.

add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile" for Prob and "Estimates of large return period" for Return.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

See Section 4.2.1 of Albrecher et al. (2017) for more details.

Weissman.p and Weissman.r are the same functions as Prob and Return but with a different name for compatibility with the old S-Plus code.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for Prob.
R	Vector of the corresponding estimates for the return period, only returned for Return.
q	The used large quantile.

### Author(s)

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

Weissman, I. (1978). "Estimation of Parameters and Large Quantiles Based on the  $k$  Largest Observations." *Journal of the American Statistical Association*, 73, 812–815.

### See Also

[Quant](#)

### Examples

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Hill estimator
H <- Hill(SOAdata)
```

```

# Exceedance probability
q <- 10^6
# Weissman estimator
Prob(SOAdata, gamma=H$gamma, q=q, plot=TRUE)

# Return period
q <- 10^6
# Weissman estimator
Return(SOAdata, gamma=H$gamma, q=q, plot=TRUE)

```

---

ProbEPD	<i>Estimator of small exceedance probabilities and large return periods using EPD</i>
---------	---

---

### Description

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the parameters from the EPD fit.

### Usage

```

ProbEPD(data, q, gamma, kappa, tau, plot = FALSE, add = FALSE,
        main = "Estimates of small exceedance probability", ...)

ReturnEPD(data, q, gamma, kappa, tau, plot = FALSE, add = FALSE,
         main = "Estimates of large return period", ...)

```

### Arguments

data	Vector of $n$ observations.
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">EPD</a> .
kappa	Vector of $n - 1$ estimates for $\kappa$ obtained from <a href="#">EPD</a> .
tau	Vector of $n - 1$ estimates for $\tau$ obtained from <a href="#">EPD</a> .
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for ProbEPD and "Estimates of large return period" for ReturnEPD.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

See Section 4.2.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for ProbEPD.
R	Vector of the corresponding estimates for the return period, only returned for ReturnEPD.
q	The used large quantile.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Joossens, E. and Segers, J. (2009). "Second-Order Refined Peaks-Over-Threshold Modelling for Heavy-Tailed Distributions." *Journal of Statistical Planning and Inference*, 139, 2800–2815.

**See Also**

[EPD](#), [Prob](#)

**Examples**

```
data(secura)

# EPD estimates for the EVI
epd <- EPD(secura$size, plot=TRUE)

# Compute exceedance probabilities
q <- 10^7
ProbEPD(secura$size, q=q, gamma=epd$gamma, kappa=epd$kappa, tau=epd$tau, plot=TRUE)

# Compute return periods
ReturnEPD(secura$size, q=q, gamma=epd$gamma, kappa=epd$kappa, tau=epd$tau,
           plot=TRUE, ylim=c(0,10^4))
```

---

ProbGH	<i>Estimator of small exceedance probabilities and large return periods using generalised Hill</i>
--------	--

---

### Description

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the generalised Hill estimates for the EVI.

### Usage

```
ProbGH(data, gamma, q, plot = FALSE, add = FALSE,
       main = "Estimates of small exceedance probability", ...)
```

```
ReturnGH(data, gamma, q, plot = FALSE, add = FALSE,
         main = "Estimates of large return period", ...)
```

### Arguments

data	Vector of $n$ observations.
gamma	Vector of $n - 2$ estimates for the EVI obtained from <a href="#">genHill</a> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for ProbGH and "Estimates of large return period" for ReturnGH.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

See Section 4.2.2 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for ProbGH.
R	Vector of the corresponding estimates for the return period, only returned for ReturnGH.
q	The used large quantile.

**Author(s)**

Tom Reynkens.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

Beirlant, J., Vynckier, P. and Teugels, J.L. (1996). "Excess Function and Estimation of the Extreme-value Index". *Bernoulli*, 2, 293–318.

**See Also**

[QuantGH](#), [genHill](#), [ProbMOM](#), [Prob](#)

**Examples**

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Hill estimator
H <- Hill(SOAdata)
# Generalised Hill estimator
gH <- genHill(SOAdata, H$gamma)

# Exceedance probability
q <- 10^7
ProbGH(SOAdata, gamma=gH$gamma, q=q, plot=TRUE)

# Return period
q <- 10^7
ReturnGH(SOAdata, gamma=gH$gamma, q=q, plot=TRUE)
```

---

ProbGPD

*Estimator of small exceedance probabilities and large return periods using GPD-MLE*

---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the GPD fit for the peaks over a threshold.



**Usage**

```
ProbGPD(data, gamma, sigma, q, plot = FALSE, add = FALSE,
        main = "Estimates of small exceedance probability", ...)
```

```
ReturnGPD(data, gamma, sigma, q, plot = FALSE, add = FALSE,
          main = "Estimates of large return period", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">GPDmle</a> .
sigma	Vector of $n - 1$ estimates for $\sigma$ obtained from <a href="#">GPDmle</a> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability" for ProbGPD and "Estimates of large return period" for ReturnGPD.
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

See Section 4.2.2 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates, only returned for ProbGPD.
R	Vector of the corresponding estimates for the return period, only returned for ReturnGPD.
q	The used large quantile.

**Author(s)**

Tom Reynkens.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**See Also**

[QuantGPD](#), [GPDmle](#), [Prob](#)

**Examples**

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# GPD-ML estimator
pot <- GPDmle(SOAdata)

# Exceedance probability
q <- 10^7
ProbGPD(SOAdata, gamma=pot$gamma, sigma=pot$sigma, q=q, plot=TRUE)

# Return period
q <- 10^7
ReturnGPD(SOAdata, gamma=pot$gamma, sigma=pot$sigma, q=q, plot=TRUE)
```

---

ProbMOM

*Estimator of small exceedance probabilities and large return periods using MOM*

---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  or large return period  $1/P(X > q)$  using the Method of Moments estimates for the EVI.

**Usage**

```
ProbMOM(data, gamma, q, plot = FALSE, add = FALSE,
        main = "Estimates of small exceedance probability", ...)

ReturnMOM(data, gamma, q, plot = FALSE, add = FALSE,
          main = "Estimates of large return period", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">Moment</a> .
q	The used large quantile (we estimate $P(X > q)$ or $1/P(X > q)$ for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.

main Title for the plot, default is "Estimates of small exceedance probability" for ProbMOM and "Estimates of large return period" for ReturnMOM.

... Additional arguments for the plot function, see [plot](#) for more details.

### Details

See Section 4.2.2 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k Vector of the values of the tail parameter  $k$ .

P Vector of the corresponding probability estimates, only returned for ProbMOM.

R Vector of the corresponding estimates for the return period, only returned for ReturnMOM.

q The used large quantile.

### Author(s)

Tom Reynkens.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

Dekkers, A.L.M, Einmahl, J.H.J. and de Haan, L. (1989). "A Moment Estimator for the Index of an Extreme-value Distribution." *Annals of Statistics*, 17, 1833–1855.

### See Also

[QuantMOM](#), [Moment](#), [ProbGH](#), [Prob](#)

### Examples

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# MOM estimator
M <- Moment(SOAdata)

# Exceedance probability
q <- 10^7
ProbMOM(SOAdata, gamma=M$gamma, q=q, plot=TRUE)
```

```
# Return period
q <- 10^7
ReturnMOM(SOAdata, gamma=M$gamma, q=q, plot=TRUE)
```

---

 ProbReg

---

*Estimator of small tail probability in regression*


---

### Description

Estimator of small tail probability  $1 - F_i(q)$  in the regression case where  $\gamma$  is constant and the regression modelling is thus only solely placed on the scale parameter.

### Usage

```
ProbReg(Z, A, q, plot = FALSE, add = FALSE,
        main = "Estimates of small exceedance probability", ...)
```

### Arguments

Z	Vector of $n$ observations (from the response variable).
A	Vector of $n - 1$ estimates for $A(i/n)$ obtained from <a href="#">ScaleReg</a> .
q	The used large quantile (we estimate $P(X_i > q)$ ) for $q$ large).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The estimator is defined as

$$1 - \hat{F}_i(q) = \hat{A}(i/n)(k+1)/(n+1)(q/Z_{n-k,n})^{-1/H_{k,n}},$$

with  $H_{k,n}$  the Hill estimator. Here, it is assumed that we have equidistant covariates  $x_i = i/n$ .

See Section 4.4.1 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates.
q	The used large quantile.

**Author(s)**

Tom Reynkens.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[QuantReg](#), [ScaleReg](#), [Prob](#)

**Examples**

```
data(norwegianfire)

Z <- norwegianfire$size[norwegianfire$year==76]

i <- 100
n <- length(Z)

# Scale estimator in i/n
A <- ScaleReg(i/n, Z, h=0.5, kernel = "epanechnikov")$A

# Small exceedance probability
q <- 10^6
ProbReg(Z, A, q, plot=TRUE)

# Large quantile
p <- 10^(-5)
QuantReg(Z, A, p, plot=TRUE)
```

---

Quant

*Weissman estimator of extreme quantiles*

---

**Description**

Compute estimates of an extreme quantile  $Q(1 - p)$  using the approach of Weissman (1978).

**Usage**

```
Quant(data, gamma, p, plot = FALSE, add = FALSE,
      main = "Estimates of extreme quantile", ...)

Weissman.q(data, gamma, p, plot = FALSE, add = FALSE,
           main = "Estimates of extreme quantile", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI, typically Hill estimates are used.
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates as a function of $k$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

See Section 4.2.1 of Albrecher et al. (2017) for more details.

Weissman.q is the same function but with a different name for compatibility with the old S-Plus code.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

**Author(s)**

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.
- Weissman, I. (1978). "Estimation of Parameters and Large Quantiles Based on the  $k$  Largest Observations." *Journal of the American Statistical Association*, 73, 812–815.

**See Also**

[Prob](#), [Quant](#), [2oQV](#)

**Examples**

```

data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Hill estimator
H <- Hill(SOAdata)
# Bias-reduced estimator (QV)
H_QV <- Hill.2oQV(SOAdata)

# Exceedance probability
p <- 10^(-5)
# Weissman estimator
Quant(SOAdata, gamma=H$gamma, p=p, plot=TRUE)

# Second order Weissman estimator (QV)
Quant.2oQV(SOAdata, gamma=H_QV$gamma, beta=H_QV$beta, b=H_QV$b, p=p,
           add=TRUE, lty=2)

```

Quant.2oQV

*Second order refined Weissman estimator of extreme quantiles (QV)***Description**

Compute second order refined Weissman estimator of extreme quantiles  $Q(1-p)$  using the quantile view.

**Usage**

```
Quant.2oQV(data, gamma, b, beta, p, plot = FALSE, add = FALSE,
           main = "Estimates of extreme quantile", ...)
```

```
Weissman.q.2oQV(data, gamma, b, beta, p, plot = FALSE, add = FALSE,
                main = "Estimates of extreme quantile", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">Hill.2oQV</a> .
b	Vector of $n - 1$ estimates for $b$ obtained from <a href="#">Hill.2oQV</a> .
beta	Vector of $n - 1$ estimates for $\beta$ obtained from <a href="#">Hill.2oQV</a> .
p	The exceedance probability of the quantile (we estimate $Q(1-p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.

main            Title for the plot, default is "Estimates of extreme quantile".  
 ...            Additional arguments for the plot function, see [plot](#) for more details.

### Details

See Section 4.2.1 of Albrecher et al. (2017) for more details.

Weissman.q.2oQV is the same function but with a different name for compatibility with the old S-Plus code.

### Value

A list with following components:

k                Vector of the values of the tail parameter  $k$ .  
 Q                Vector of the corresponding quantile estimates.  
 p                The used exceedance probability.

### Author(s)

Tom Reynkens based on S-Plus code from Yuri Goegebeur.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

### See Also

[Quant](#), [Hill.2oQV](#)

### Examples

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Hill estimator
H <- Hill(SOAdata)
# Bias-reduced estimator (QV)
H_QV <- Hill.2oQV(SOAdata)

# Exceedance probability
p <- 10^(-5)
# Weissman estimator
Quant(SOAdata, gamma=H$gamma, p=p, plot=TRUE)

# Second order Weissman estimator (QV)
```



```
Quant.2oQV(SOAdata, gamma=H_QV$gamma, beta=H_QV$beta, b=H_QV$b, p=p,
           add=TRUE, lty=2)
```

QuantGH

*Estimator of extreme quantiles using generalised Hill***Description**

Compute estimates of an extreme quantile  $Q(1 - p)$  using generalised Hill estimates of the EVI.

**Usage**

```
QuantGH(data, gamma, p, plot = FALSE, add = FALSE,
         main = "Estimates of extreme quantile", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 2$ estimates for the EVI obtained from <a href="#">genHill</a> .
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

See Section 4.2.2 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

**Author(s)**

Tom Reynkens.

## References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.
- Beirlant, J., Vynckier, P. and Teugels, J.L. (1996). "Excess Function and Estimation of the Extreme-value Index". *Bernoulli*, 2, 293–318.

## See Also

[ProbGH](#), [genHill](#), [QuantMOM](#), [Quant](#)

## Examples

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# Hill estimator
H <- Hill(SOAdata)
# Generalised Hill estimator
gH <- genHill(SOAdata, H$gamma)

# Large quantile
p <- 10^(-5)
QuantGH(SOAdata, p=p, gamma=gH$gamma, plot=TRUE)
```

---

QuantGPD

*Estimator of extreme quantiles using GPD-MLE*

---

## Description

Computes estimates of an extreme quantile  $Q(1-p)$  using the GPD fit for the peaks over a threshold.

## Usage

```
QuantGPD(data, gamma, sigma, p, plot = FALSE, add = FALSE,
          main = "Estimates of extreme quantile", ...)
```

## Arguments

- data**            Vector of  $n$  observations.
- gamma**           Vector of  $n - 1$  estimates for the EVI obtained from [GPDmle](#).
- sigma**           Vector of  $n - 1$  estimates for  $\sigma$  obtained from [GPDmle](#).
- p**                The exceedance probability of the quantile (we estimate  $Q(1 - p)$  for  $p$  small).

plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

See Section 4.2.2 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

### Author(s)

Tom Reynkens.

### References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

### See Also

[ProbGPD](#), [GPDmle](#), [Quant](#)

### Examples

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# GPD-ML estimator
pot <- GPDmle(SOAdata)

# Large quantile
p <- 10^(-5)
QuantGPD(SOAdata, p=p, gamma=pot$gamma, sigma=pot$sigma, plot=TRUE)
```

QuantMOM

*Estimator of extreme quantiles using MOM***Description**

Compute estimates of an extreme quantile  $Q(1 - p)$  using the Method of Moments estimates of the EVI.

**Usage**

```
QuantMOM(data, gamma, p, plot = FALSE, add = FALSE,
          main = "Estimates of extreme quantile", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">Moment</a> .
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

See Section 4.2.2 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

**Author(s)**

Tom Reynkens.

## References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.
- Dekkers, A.L.M, Einmahl, J.H.J. and de Haan, L. (1989). "A Moment Estimator for the Index of an Extreme-value Distribution." *Annals of Statistics*, 17, 1833–1855.

## See Also

[ProbMOM](#), [Moment](#), [QuantGH](#), [Quant](#)

## Examples

```
data(soa)

# Look at last 500 observations of SOA data
SOAdata <- sort(soa$size)[length(soa$size)-(0:499)]

# MOM estimator
M <- Moment(SOAdata)

# Large quantile
p <- 10^(-5)
QuantMOM(SOAdata, p=p, gamma=M$gamma, plot=TRUE)
```

---

QuantReg

*Estimator of extreme quantiles in regression*

---

## Description

Estimator of extreme quantile  $Q_i(1-p)$  in the regression case where  $\gamma$  is constant and the regression modelling is thus only solely placed on the scale parameter.

## Usage

```
QuantReg(Z, A, p, plot = FALSE, add = FALSE,
         main = "Estimates of extreme quantile", ...)
```

## Arguments

- |      |  |
|------|--|
| Z    | Vector of $n$ observations (from the response variable).                                       |
| A    | Vector of $n - 1$ estimates for $A(i/n)$ obtained from <a href="#">ScaleReg</a> .              |
| p    | The exceedance probability of the quantile (we estimate $Q_i(1-p)$ for $p$ small).             |
| plot | Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE. |

add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The estimator is defined as

$$\hat{Q}_i(1-p) = Z_{n-k,n}((k+1)/((n+1) \times p) \hat{A}(i/n))^{H_{k,n}},$$

with  $H_{k,n}$  the Hill estimator. Here, it is assumed that we have equidistant covariates  $x_i = i/n$ .

See Section 4.4.1 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

### Author(s)

Tom Reynkens.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

### See Also

[ProbReg](#), [ScaleReg](#), [Quant](#)

### Examples

```
data(norwegianfire)

Z <- norwegianfire$size[norwegianfire$year==76]

i <- 100
n <- length(Z)

# Scale estimator in i/n
A <- ScaleReg(i/n, Z, h=0.5, kernel = "epanechnikov")$A

# Small exceedance probability
q <- 10^6
ProbReg(Z, A, q, plot=TRUE)
```

```
# Large quantile
p <- 10^(-5)
QuantReg(Z, A, p, plot=TRUE)
```

---

Scale	<i>Scale estimator</i>
-------	------------------------

---

### Description

Computes the estimator for the scale parameter as described in Beirlant et al. (2016).

### Usage

```
Scale(data, gamma = NULL, logk = FALSE, plot = FALSE, add = FALSE,
      main = "Estimates of scale parameter", ...)
```

### Arguments

<code>data</code>	Vector of $n$ observations.
<code>gamma</code>	Vector of $n - 1$ estimates for the EVI. When NULL (the default value), Hill estimates are computed.
<code>logk</code>	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is FALSE.
<code>plot</code>	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
<code>add</code>	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
<code>main</code>	Title for the plot, default is "Estimates of scale parameter".
<code>...</code>	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The scale estimates are computed based on the following model for the CDF:  $1 - F(x) = Ax^{-1/\gamma}$ , where  $A := C^{1/\gamma}$  is the scale parameter:

$$\hat{A}_{k,n} = (k + 1)/(n + 1)X_{n-k,n}^{1/H_{k,n}}$$

where  $H_{k,n}$  are the Hill estimates.

See Section 4.2.1 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

<code>k</code>	Vector of the values of the tail parameter $k$ .
<code>A</code>	Vector of the corresponding scale estimates.
<code>C</code>	Vector of the corresponding estimates for $C$ , see Details.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Schoutens, W., De Spiegeleer, J., Reynkens, T. and Herrmann, K. (2016). "Hunting for Black Swans in the European Banking Sector Using Extreme Value Analysis." In: Jan Kallsen and Antonis Papapantoleon (eds.), *Advanced Modelling in Mathematical Finance*, Springer International Publishing, Switzerland, pp. 147–166.

**See Also**

[ScaleEPD](#), [Scale.2o](#), [Hill](#)

**Examples**

```
data(secura)

# Hill estimator
H <- Hill(secura$size)

# Scale estimator
S <- Scale(secura$size, gamma=H$gamma, plot=FALSE)

# Plot logarithm of scale
plot(S$k, log(S$A), xlab="k", ylab="log(Scale)", type="l")
```

---

Scale.2o

*Bias-reduced scale estimator using second order Hill estimator*


---

**Description**

Computes the bias-reduced estimator for the scale parameter using the second-order Hill estimator.

**Usage**

```
Scale.2o(data, gamma, b, beta, logk = FALSE, plot = FALSE, add = FALSE,
         main = "Estimates of scale parameter", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <code>Hill.2oQV</code> .
b	Vector of $n - 1$ estimates for $B$ obtained from <code>Hill.2oQV</code> .
beta	Vector of $n - 1$ estimates for $\beta$ obtained from <code>Hill.2oQV</code> .



logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of scale parameter".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The scale estimates are computed based on the following model for the CDF:  $1-F(x) = Ax^{-1/\gamma}(1+bx^{-\beta}(1+o(1)))$ , where  $A := C^{1/\gamma}$  is the scale parameter.

See Section 4.2.1 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
A	Vector of the corresponding scale estimates.
C	Vector of the corresponding estimates for $C$ , see details.

### Author(s)

Tom Reynkens

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Schoutens, W., De Spiegeleer, J., Reynkens, T. and Herrmann, K. (2016). "Hunting for Black Swans in the European Banking Sector Using Extreme Value Analysis." In: Jan Kallsen and Antonis Papapantoleon (eds.), *Advanced Modelling in Mathematical Finance*, Springer International Publishing, Switzerland, pp. 147–166.

### See Also

[Scale](#), [ScaleEPD](#), [Hill.2oQV](#)

### Examples

```
data(secura)

# Hill estimator
H <- Hill(secura$size)
# Bias-reduced Hill estimator
H2o <- Hill.2oQV(secura$size)
```

```

# Scale estimator
S <- Scale(secura$size, gamma=H$gamma, plot=FALSE)
# Bias-reduced scale estimator
S2o <- Scale.2o(secura$size, gamma=H2o$gamma, b=H2o$b,
               beta=H2o$beta, plot=FALSE)

# Plot logarithm of scale
plot(S$k,log(S$A), xlab="k", ylab="log(Scale)", type="l")
lines(S2o$k,log(S2o$A), lty=2)

```

---

ScaleEPD

*Bias-reduced scale estimator using EPD estimator*


---

### Description

Computes the bias-reduced estimator for the scale parameter using the EPD estimator (Beirlant et al., 2016).

### Usage

```
ScaleEPD(data, gamma, kappa, logk = FALSE, plot = FALSE, add = FALSE,
         main = "Estimates of scale parameter", ...)
```

### Arguments

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from EPD.
kappa	Vector of $n - 1$ estimates for $\kappa$ obtained from EPD.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ ( <code>logk=TRUE</code> ) or as a function of $k$ . Default is <code>FALSE</code> .
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is <code>FALSE</code> .
add	Logical indicating if the estimates should be added to an existing plot, default is <code>FALSE</code> .
main	Title for the plot, default is "Estimates of scale parameter".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The scale estimates are computed based on the following model for the CDF:  $1 - F(x) = Ax^{-1/\gamma}(1 + bx^{-\beta}(1 + o(1)))$ , where  $A := C^{1/\gamma}$  is the scale parameter. Using the EPD approach we replace  $bx^{-\beta}$  by  $\kappa/\gamma$ .

See Section 4.2.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
A	Vector of the corresponding scale estimates.
C	Vector of the corresponding estimates for $C$ , see details.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Schoutens, W., De Spiegeleer, J., Reynkens, T. and Herrmann, K. (2016). "Hunting for Black Swans in the European Banking Sector Using Extreme Value Analysis." In: Jan Kallsen and Antonis Papapantoleon (eds.), *Advanced Modelling in Mathematical Finance*, Springer International Publishing, Switzerland, pp. 147–166.

**See Also**

[Scale](#), [Scale.2o](#), [EPD](#)

**Examples**

```
data(secura)

# Hill estimator
H <- Hill(secura$size)
# EPD estimator
epd <- EPD(secura$size)

# Scale estimator
S <- Scale(secura$size, gamma=H$gamma, plot=FALSE)
# Bias-reduced scale estimator
Sepd <- ScaleEPD(secura$size, gamma=epd$gamma, kappa=epd$kappa, plot=FALSE)

# Plot logarithm of scale
plot(S$k, log(S$A), xlab="k", ylab="log(Scale)", type="l")
lines(Sepd$k, log(Sepd$A), lty=2)
```

---

ScaleReg	<i>Scale estimator in regression</i>
----------	--------------------------------------

---

### Description

Estimator of the scale parameter in the regression case where  $\gamma$  is constant and the regression modelling is thus placed solely on the scale parameter.

### Usage

```
ScaleReg(s, Z, kernel = c("normal", "uniform", "triangular", "epanechnikov", "biweight"),
        h, plot = TRUE, add = FALSE, main = "Estimates of scale parameter", ...)
```

### Arguments

s	Point to evaluate the scale estimator in.
Z	Vector of $n$ observations (from the response variable).
kernel	The kernel used in the estimator. One of "normal" (default), "uniform", "triangular", "epanechnikov" and "biweight".
h	The bandwidth used in the kernel function.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of scale parameter".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The scale estimator is computed as

$$\hat{A}(s) = 1/(k+1) \sum_{i=1}^n 1_{Z_i > Z_{n-k,n}} K_h(s - i/n)$$

with  $K_h(x) = K(x/h)/h$ ,  $K$  the kernel function and  $h$  the bandwidth. Here, it is assumed that we have equidistant covariates  $x_i = i/n$ .

See Section 4.4.1 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
A	Vector of the corresponding scale estimates.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[ProbReg](#), [QuantReg](#), [scale](#), [Hill](#)

**Examples**

```
data(norwegianfire)

Z <- norwegianfire$size[norwegianfire$year==76]

i <- 100
n <- length(Z)

# Scale estimator in i/n
A <- ScaleReg(i/n, Z, h=0.5, kernel = "epanechnikov")$A

# Small exceedance probability
q <- 10^6
ProbReg(Z, A, q, plot=TRUE)

# Large quantile
p <- 10^(-5)
QuantReg(Z, A, p, plot=TRUE)
```

---

secura

*Secura dataset*

---

**Description**

Automobile claims from 1988 to 2001, gathered from several European insurance companies, exceeding 1 200 000 Euro. Note that the data were, among others, corrected for inflation.

**Usage**

```
data("secura")
```

**Format**

A data frame with 371 observations on the following 2 variables:

year Year of claim occurrence.

size Size of automobile insurance claim (in EUR).

## References

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

## Examples

```
data(secura)

# Exponential QQ-plot of Secura data
ExpQQ(secura$size)

# Pareto QQ-plot of Secura data
ParetoQQ(secura$size)

# Mean excess plot of Secura data (function of k)
MeanExcess(secura$size, k=TRUE)

# Mean excess plot of Secura data (function of order statistics)
MeanExcess(secura$size, k=FALSE)
```

---

soa

*SOA Group Medical Insurance Large Claims Database*

---

## Description

The SOA Group Medical Insurance Large Claims Database records, among others, all the claim amounts exceeding 25,000 USD in the year 1991.

## Usage

```
data("soa")
```

## Format

A data frame with 75789 observations on the following variable:

size Claim size (in USD).

## Source

Grazier, K. L. and G'Sell Associates (1997). *Group Medical Insurance Large Claims Database Collection and Analysis*. SOA Monograph M-HB97-1, Society of Actuaries, Schaumburg.

Society of Actuaries, <https://www.soa.org/resources/experience-studies/2000-2004/91-92-group-medical-cla>

## References

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**Examples**

```

data(soa)

# Histogram of log-claim amount
hist(log(soa$size),breaks=seq(10,16,0.2),xlab="log(Claim size)")

# Exponential QQ-plot of claim amount
ExpQQ(soa$size)

# Mean excess plot of claim amount (function of k)
MeanExcess(soa$size, k=TRUE)

# Mean excess plot of claim amount (function of order statistics)
MeanExcess(soa$size, k=FALSE)

```

---

Splice	<i>Spliced distribution</i>
--------	-----------------------------

---

**Description**

Density, distribution function, quantile function and random generation for the fitted spliced distribution.

**Usage**

```

dSplice(x, splicefit, log = FALSE)

pSplice(x, splicefit, lower.tail = TRUE, log.p = FALSE)

qSplice(p, splicefit, lower.tail = TRUE, log.p = FALSE)

rSplice(n, splicefit)

```

**Arguments**

x	Vector of points to evaluate the CDF or PDF in.
p	Vector of probabilities.
n	Number of observations.
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> , <a href="#">SpliceFitPareto</a> or <a href="#">SpliceFitGPD</a> .
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

See Reynkens et al. (2017) and Section 4.3 in Albrecher et al. (2017) for details.

**Value**

dSplice gives the density function evaluated in  $x$ , pSplice the CDF evaluated in  $x$  and qSplice the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rSplice returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens with R code from Roel Verbelen for the mixed Erlang PDF, CDF and quantiles.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758.

**See Also**

[VaR](#), [SpliceFit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#), [SpliceECDF](#), [SpliceLL](#), [SplicePP](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")
```



```

p <- seq(0, 1, 0.01)

# Plot of splicing quantiles
plot(p, qSplice(p, splicefit), type="l", xlab="p", ylab="Q(p)")

# Plot of VaR
plot(p, VaR(p, splicefit), type="l", xlab="p", ylab=bquote(VaR[p]))

# Random sample from spliced distribution
x <- rSplice(1000, splicefit)

## End(Not run)

```

---

SpliceECDF

*Plot of fitted and empirical survival function*


---

### Description

This function plots the fitted survival function of the spliced distribution together with the empirical survival function (determined using the Empirical CDF (ECDF)). Moreover,  $100(1 - \alpha)\%$  confidence bands are added.

### Usage

```
SpliceECDF(x, X, splicefit, alpha = 0.05, ...)
```

### Arguments

x	Vector of points to plot the functions at.
X	Data used for fitting the distribution.
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> or <a href="#">SpliceFitGPD</a> .
alpha	$100(1 - \alpha)\%$ is the confidence level for the confidence bands. Default is $\alpha = 0.05$ .
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

Use [SpliceTB](#) for censored data.

Confidence bands are determined using the Dvoretzky-Kiefer-Wolfowitz inequality (Massart, 1990).

See Reynkens et al. (2017) and Section 4.3.1 in Albrecher et al. (2017) for more details.

### Author(s)

Tom Reynkens

## References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Massart, P. (1990). The Tight Constant in the Dvoretzky-Kiefer-Wolfowitz Inequality. *Annals of Probability*, 18, 1269–1283.
- Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.
- Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758.

## See Also

[SpliceTB](#), [pSplice](#), [ecdf](#), [SpliceFitPareto](#), [SpliceFitGPD](#), [SpliceLL](#), [SplicePP](#), [SpliceQQ](#)

## Examples

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and empirical survival function
SpliceECDF(x, X, splicefit)

# Log-log plot with empirical survival function and fitted survival function
SpliceLL(x, X, splicefit)

# PP-plot of empirical survival function and fitted survival function
SplicePP(X, splicefit)

# PP-plot of empirical survival function and
# fitted survival function with log-scales
SplicePP(X, splicefit, log=TRUE)
```

```
# Splicing QQ-plot
SpliceQQ(X, splicefit)

## End(Not run)
```

---

SpliceFit

*Splicing fit*


---

### Description

Create an S3 object using ME-Pa or ME-GPD splicing fit obtained from [SpliceFitPareto](#), [SpliceFiticPareto](#) or [SpliceFitGPD](#).

### Usage

```
SpliceFit(const, trunclower, t, type, MEfit, EVTfit, loglik = NULL, IC = NULL)
```

### Arguments

const	Vector of splicing constants or a single splicing constant.
trunclower	Lower truncation point.
t	Vector of splicing points or a single splicing point.
type	Vector of types of the distributions: "ME" and then for each fitted EVT distribution: "Pa" (Pareto), "TPa" (truncated Pareto) or "GPD" (GPD).
MEfit	<a href="#">MEfit</a> object with details on the mixed Erlang fit.
EVTfit	<a href="#">EVTfit</a> object with details on the EVT fit.
loglik	Log-likelihood of the fitted model. When NULL (default), not included in the object.
IC	Information criteria of the fitted model. When NULL (default), not included in the object. This vector should have length 1, 2 or 3 when included.

### Details

See Reynkens et al. (2017) and Section 4.3 in Albrecher et al. (2017) for details.

### Value

An S3 object containing the above input arguments and values for  $\pi$ , the splicing weights. These splicing weights are equal to

$$\pi_1 = const_1, \pi_2 = const_2 - const_1, \dots, \pi_{l+1} = 1 - const_l = 1 - (\pi_1 + \dots + \pi_l)$$

when  $l \geq 2$  and

$$\pi_1 = const_1, \pi_2 = 1 - const_1 = 1 - \pi_1$$

when  $l = 1$ , where  $l$  is the length of const.

A summary method is available.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[MEfit](#), [EVTfit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#)

**Examples**

```
# Create MEfit object
mefit <- MEfit(p=c(0.65,0.35), shape=c(39,58), theta=16.19, M=2)

# Create EVTfit object
evtfit <- EVTfit(gamma=c(0.76,0.64), endpoint=c(39096, Inf))

# Create SpliceFit object
splicefit <- SpliceFit(const=c(0.5,0.996), trunclower=0, t=c(1020,39096), type=c("ME","TPa","Pa"),
                      MEfit=mefit, EVTfit=evtfit)

# Show summary
summary(splicefit)
```

---

SpliceFitGPD

*Splicing of mixed Erlang and GPD using POT-MLE*

---

**Description**

Fit spliced distribution of a mixed Erlang distribution and a Generalised Pareto Distribution (GPD). The parameters of the GPD are determined using the POT-MLE approach.

**Usage**

```
SpliceFitGPD(X, const = NULL, tsplice = NULL, M = 3, s = 1:10, trunclower = 0,
             ncores = NULL, criterium = c("BIC","AIC"), reduceM = TRUE,
             eps = 10^(-3), beta_tol = 10^(-5), maxiter = Inf)
```

**Arguments**

X	Data used for fitting the distribution.
const	The probability of the quantile where the ME distribution will be spliced with the GPD distribution. Default is NULL meaning the input from <code>tsplICE</code> is used.
tsplICE	The point where the ME distribution will be spliced with the GPD distribution. Default is NULL meaning the input from <code>const</code> is used.
M	Initial number of Erlang mixtures, default is 3. This number can change when determining an optimal mixed Erlang fit using an information criterion.
s	Vector of spread factors for the EM algorithm, default is 1:10. We loop over these factors when determining an optimal mixed Erlang fit using an information criterion, see Verbelen et al. (2016).
truncLower	Lower truncation point. Default is 0.
ncores	Number of cores to use when determining an optimal mixed Erlang fit using an information criterion. When NULL (default), $\max(nc-1, 1)$ cores are used where <code>nc</code> is the number of cores as determined by <code>detectCores</code> .
criterium	Information criterion used to select the number of components of the ME fit and <code>s</code> . One of "AIC" and "BIC" (default).
reduceM	Logical indicating if <code>M</code> should be reduced based on the information criterion, default is TRUE.
eps	Covergence threshold used in the EM algorithm (ME part). Default is $10^{-3}$ .
beta_tol	Threshold for the mixing weights below which the corresponding shape parameter vector is considered neglectable (ME part). Default is $10^{-5}$ .
maxiter	Maximum number of iterations in a single EM algorithm execution (ME part). Default is Inf meaning no maximum number of iterations.

**Details**

See Reynkens et al. (2017), Section 4.3.1 of Albrecher et al. (2017) and Verbelen et al. (2015) for details. The code follows the notation of the latter. Initial values follow from Verbelen et al. (2016).

**Value**

A `SpliceFit` object.

**Author(s)**

Tom Reynkens with R code from Roel Verbelen for fitting the mixed Erlang distribution.

**References**

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758.

Verbelen, R., Antonio, K. and Claeskens, G. (2016). "Multivariate Mixtures of Erlangs for Density Estimation Under Censoring." *Lifetime Data Analysis*, 22, 429–455.

### See Also

[SpliceFitPareto](#), [SpliceFiticPareto](#), [Splice](#), [GPDfit](#)

### Examples

```
## Not run:

# GPD random sample
X <- rgpd(1000, gamma = 0.5, sigma = 2)

# Splice ME and GPD
splicefit <- SpliceFitGPD(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and empirical survival function
SpliceECDF(x, X, splicefit)

# Log-log plot with empirical survival function and fitted survival function
SpliceLL(x, X, splicefit)

# PP-plot of empirical survival function and fitted survival function
SplicePP(X, splicefit)

# PP-plot of empirical survival function and
# fitted survival function with log-scales
SplicePP(X, splicefit, log=TRUE)

# Splicing QQ-plot
SpliceQQ(X, splicefit)

## End(Not run)
```

---

SpliceFiticPareto      *Splicing of mixed Erlang and Pareto for interval censored data*

---

### Description

Fit spliced distribution of a mixed Erlang distribution and a Pareto distribution adapted for interval censoring and truncation.

### Usage

```
SpliceFiticPareto(L, U, censored, tsplice, M = 3, s = 1:10, trunclower = 0,
  truncupper = Inf, ncores = NULL, criterium = c("BIC", "AIC"),
  reduceM = TRUE, eps = 10-3, beta_tol = 10-5, maxiter = Inf,
  cpp = FALSE)
```

### Arguments

L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals.
censored	A logical vector of length $n$ indicating if an observation is censored.
tsplice	The splicing point $t$ .
M	Initial number of Erlang mixtures, default is 3. This number can change when determining an optimal mixed Erlang fit using an information criterion.
s	Vector of spread factors for the EM algorithm, default is 1:10. We loop over these factors when determining an optimal mixed Erlang fit using an information criterion, see Verbelen et al. (2016).
trunclower	Lower truncation point. Default is 0.
truncupper	Upper truncation point. Default is Inf (no upper truncation).
ncores	Number of cores to use when determining an optimal mixed Erlang fit using an information criterion. When NULL (default), $\max(nc-1, 1)$ cores are used where $nc$ is the number of cores as determined by <code>detectCores</code> .
criterium	Information criterion used to select the number of components of the ME fit and $s$ . One of "AIC" and "BIC" (default).
reduceM	Logical indicating if $M$ should be reduced based on the information criterion, default is TRUE.
eps	Covergence threshold used in the EM algorithm. Default is $10^{-3}$ .
beta_tol	Threshold for the mixing weights below which the corresponding shape parameter vector is considered neglectable (ME part). Default is $10^{-5}$ .
maxiter	Maximum number of iterations in a single EM algorithm execution. Default is Inf meaning no maximum number of iterations.
cpp	Use C++ implementation (cpp=TRUE) or R implementation (cpp=FALSE) of the algorithm. Default is FALSE meaning the plain R implementation is used.

**Details**

See Reynkens et al. (2017), Section 4.3.2 of Albrecher et al. (2017) and Verbelen et al. (2015) for details. The code follows the notation of the latter. Initial values follow from Verbelen et al. (2016).

Right censored data should be entered as  $L=1$  and  $U=\text{truncupper}$ , and left censored data should be entered as  $L=\text{trunclower}$  and  $U=u$ .

**Value**

A [SpliceFit](#) object.

**Author(s)**

Tom Reynkens based on R code from Roel Verbelen for fitting the mixed Erlang distribution (without splicing).

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758.

Verbelen, R., Antonio, K. and Claeskens, G. (2016). "Multivariate Mixtures of Erlangs for Density Estimation Under Censoring." *Lifetime Data Analysis*, 22, 429–455.

**See Also**

[SpliceFitPareto](#), [SpliceFitGPD](#), [Splice](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
```



```

U[censored] <- Inf

# Splice ME and Pareto
splicefit <- SpliceFitPareto(L=Z, U=U, censored=censored, tsplice=quantile(Z,0.9))

x <- seq(0,20,0.1)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and Turnbull survival function
SpliceTB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# Log-log plot with Turnbull survival function and fitted survival function
SpliceLL_TB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and fitted survival function
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and
# fitted survival function with log-scales
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit, log=TRUE)

# QQ-plot using Turnbull survival function and fitted survival function
SpliceQQ_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

## End(Not run)

```

---

SpliceFitPareto

*Splicing of mixed Erlang and Pareto*


---

### Description

Fit spliced distribution of a mixed Erlang distribution and Pareto distribution(s). The shape parameter(s) of the Pareto distribution(s) is determined using the Hill estimator.

### Usage

```

SpliceFitPareto(X, const = NULL, tsplice = NULL, M = 3, s = 1:10, trunclower = 0,
truncupper = Inf, EVTtruncation = FALSE, ncores = NULL,
criterion = c("BIC", "AIC"), reduceM = TRUE,
eps = 10-3, beta_tol = 10-5, maxiter = Inf)

```

```
SpliceFitHill(X, const = NULL, tsplce = NULL, M = 3, s = 1:10, trunclower = 0,
truncupper = Inf, EVTtruncation = FALSE, ncores = NULL,
criterium = c("BIC","AIC"), reduceM = TRUE,
eps = 10^(-3), beta_tol = 10^(-5), maxiter = Inf)
```

### Arguments

X	Data used for fitting the distribution.
const	Vector of length $l$ containing the probabilities of the quantiles where the distributions will be spliced (splicing points). The ME distribution will be spliced with $l$ Pareto distributions. Default is NULL meaning the input from <code>tsplce</code> is used.
tsplce	Vector of length $l$ containing the splicing points. The ME distribution will be spliced with $l$ Pareto distributions. Default is NULL meaning the input from <code>const</code> is used.
M	Initial number of Erlang mixtures, default is 3. This number can change when determining an optimal mixed Erlang fit using an information criterion.
s	Vector of spread factors for the EM algorithm, default is 1:10. We loop over these factors when determining an optimal mixed Erlang fit using an information criterion, see Verbelen et al. (2016).
trunclower	Lower truncation point. Default is 0.
truncupper	Upper truncation point. Default is Inf (no upper truncation). When <code>truncupper=Inf</code> and <code>EVTtruncation=TRUE</code> , the truncation point is estimated using the approach of Beirlant et al. (2016).
EVTtruncation	Logical indicating if the $l$ -th Pareto distribution is a truncated Pareto distribution. Default is FALSE.
ncores	Number of cores to use when determining an optimal mixed Erlang fit using an information criterion. When NULL (default), $\max(nc-1, 1)$ cores are used where $nc$ is the number of cores as determined by <code>detectCores</code> .
criterium	Information criterion used to select the number of components of the ME fit and $s$ . One of "AIC" and "BIC" (default).
reduceM	Logical indicating if M should be reduced based on the information criterion, default is TRUE.
eps	Covergence threshold used in the EM algorithm (ME part). Default is $10^{-3}$ .
beta_tol	Threshold for the mixing weights below which the corresponding shape parameter vector is considered neglectable (ME part). Default is $10^{-5}$ .
maxiter	Maximum number of iterations in a single EM algorithm execution (ME part). Default is Inf meaning no maximum number of iterations.

### Details

See Reynkens et al. (2017), Section 4.3.1 of Albrecher et al. (2017) and Verbelen et al. (2015) for details. The code follows the notation of the latter. Initial values follow from Verbelen et al. (2016).

The `SpliceFitHill` function is the same function but with a different name for compatibility with old versions of the package.

Use `SpliceFiticPareto` when censoring is present.

### Value

A `SpliceFit` object.

### Author(s)

Tom Reynkens with R code from Roel Verbelen for fitting the mixed Erlang distribution.

### References

- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.
- Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.
- Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758.
- Verbelen, R., Antonio, K. and Claeskens, G. (2016). "Multivariate Mixtures of Erlangs for Density Estimation Under Censoring." *Lifetime Data Analysis*, 22, 429–455.

### See Also

[SpliceFiticPareto](#), [SpliceFitGPD](#), [Splice](#), [Hill](#), [trHill](#)

### Examples

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")
```

```

# Fitted survival function and empirical survival function
SpliceECDF(x, X, splicefit)

# Log-log plot with empirical survival function and fitted survival function
SpliceLL(x, X, splicefit)

# PP-plot of empirical survival function and fitted survival function
SplicePP(X, splicefit)

# PP-plot of empirical survival function and
# fitted survival function with log-scales
SplicePP(X, splicefit, log=TRUE)

# Splicing QQ-plot
SpliceQQ(X, splicefit)

## End(Not run)

```

---

SpliceLL

*LL-plot with fitted and empirical survival function*


---

### Description

This function plots the logarithm of the empirical survival function (determined using the Empirical CDF (ECDF)) versus the logarithm of the data. Moreover, the logarithm of the fitted survival function of the spliced distribution is added.

### Usage

```
SpliceLL(x = sort(X), X, splicefit, plot = TRUE, main = "Splicing LL-plot", ...)
```

### Arguments

<code>x</code>	Vector of points to plot the fitted survival function at. By default we plot it at the data points.
<code>X</code>	Data used for fitting the distribution.
<code>splicefit</code>	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> or <a href="#">SpliceFitGPD</a> .
<code>plot</code>	Logical indicating if the splicing LL-plot should be made, default is TRUE.
<code>main</code>	Title for the plot, default is "Splicing LL-plot".
<code>...</code>	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The LL-plot consists of the points

$$(\log(x_{i,n}), \log(1 - \hat{F}(x_{i,n})))$$

for  $i = 1, \dots, n$  with  $n$  the length of the data,  $x_{i,n}$  the  $i$ -th smallest observation and  $\hat{F}$  the empirical distribution function. Then, the line

$$(\log(x), \log(1 - \hat{F}_{spliced}(x))),$$

with  $\hat{F}_{spliced}$  the fitted spliced distribution function, is added.

Use `SpliceLL_TB` for censored data.

See Reynkens et al. (2017) and Section 4.3.1 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

<code>logX</code>	Vector of the logarithms of the sorted data.
<code>sll.the</code>	Vector of the theoretical log-probabilities $\log(1 - \hat{F}_{spliced}(x))$ .
<code>logx</code>	Vector of the logarithms of the points to plot the fitted survival function at.
<code>sll.emp</code>	Vector of the empirical log-probabilities $\log(1 - \hat{F}(x_{i,n}))$ .

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[SpliceLL\\_TB](#), [pSplice](#), [ecdf](#), [SpliceFitPareto](#), [SpliceFitGPD](#), [SpliceECDF](#), [SplicePP](#), [SpliceQQ](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)
```

```

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and empirical survival function
SpliceECDF(x, X, splicefit)

# Log-log plot with empirical survival function and fitted survival function
SpliceLL(x, X, splicefit)

# PP-plot of empirical survival function and fitted survival function
SplicePP(X, splicefit)

# PP-plot of empirical survival function and
# fitted survival function with log-scales
SplicePP(X, splicefit, log=TRUE)

# Splicing QQ-plot
SpliceQQ(X, splicefit)

## End(Not run)

```

---

SpliceLL\_TB

*LL-plot with fitted and Turnbull survival function*


---

### Description

This function plots the logarithm of the Turnbull survival function (which is suitable for interval censored data) versus the logarithm of the data. Moreover, the logarithm of the fitted survival function of the spliced distribution is added.

### Usage

```

SpliceLL_TB(x = sort(L), L, U = L, censored, splicefit, plot = TRUE,
            main = "Splicing LL-plot", ...)

```

**Arguments**

x	Vector of points to plot the fitted survival function at. By default we plot it at the points L.
L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals. By default, they are equal to L.
censored	A logical vector of length $n$ indicating if an observation is censored.
splicefit	A <code>SpliceFit</code> object, e.g. output from <code>SpliceFiticPareto</code> .
plot	Logical indicating if the splicing LL-plot should be made, default is TRUE.
main	Title for the plot, default is "Splicing LL-plot".
...	Additional arguments for the plot function, see <code>plot</code> for more details.

**Details**

The LL-plot consists of the points

$$(\log(L_{i,n}), \log(1 - \hat{F}^{TB}(L_{i,n})))$$

for  $i = 1, \dots, n$  with  $n$  the length of the data,  $x_{i,n}$  the  $i$ -th smallest observation and  $\hat{F}^{TB}$  the Turnbull estimator for the distribution function. Then, the line

$$(\log(x), \log(1 - \hat{F}_{spliced}(x))),$$

with  $\hat{F}_{spliced}$  the fitted spliced distribution function, is added.

Right censored data should be entered as `L=1` and `U=truncupper`, and left censored data should be entered as `L=trunclower` and `U=u`. The limits `trunclower` and `truncupper` are obtained from the `SpliceFit` object.

If the **interval** package is installed, the `icfit` function is used to compute the Turnbull estimator. Otherwise, `survfit.formula` from **survival** is used.

Use `SpliceLL` for non-censored data.

See Reynkens et al. (2017) and Section 4.3.2 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

<code>logX</code>	Vector of the logarithms of the sorted left boundaries of the intervals.
<code>sll.the</code>	Vector of the theoretical log-probabilities $\log(1 - \hat{F}_{spliced}(x))$ .
<code>logx</code>	Vector of the logarithms of the points to plot the fitted survival function at.
<code>sll.emp</code>	Vector of the empirical log-probabilities $\log(1 - \hat{F}^{TB}(x_{i,n}))$ .

**Author(s)**

Tom Reynkens

## References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

## See Also

[SpliceLL](#), [pSplice](#), [Turnbull](#), [icfit](#), [SpliceFiticPareto](#), [SpliceTB](#), [SplicePP\\_TB](#), [SpliceQQ\\_TB](#)

## Examples

```
## Not run:

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Splice ME and Pareto
splicefit <- SpliceFiticPareto(L=Z, U=U, censored=censored, tsplice=quantile(Z,0.9))

x <- seq(0,20,0.1)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and Turnbull survival function
SpliceTB(x, L=Z, U=U, censored=censored, splicefit=splicefit)
```



```

# Log-log plot with Turnbull survival function and fitted survival function
SpliceLL_TB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and fitted survival function
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and
# fitted survival function with log-scales
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit, log=TRUE)

# QQ-plot using Turnbull survival function and fitted survival function
SpliceQQ_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

## End(Not run)

```

---

SplicePP

*PP-plot with fitted and empirical survival function*


---

## Description

This function plots the fitted survival function of the spliced distribution versus the empirical survival function (determined using the Empirical CDF (ECDF)).

## Usage

```

SplicePP(X, splicefit, x = sort(X), log = FALSE, plot = TRUE,
         main = "Splicing PP-plot", ...)

```

## Arguments

X	Data used for fitting the distribution.
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> or <a href="#">SpliceFitGPD</a> .
x	Vector of points to plot the functions at. By default we plot them at the data points.
log	Logical indicating if minus the logarithms of the survival probabilities are plotted versus each other, default is FALSE.
plot	Logical indicating if the splicing PP-plot should be made, default is TRUE.
main	Title for the plot, default is "Splicing PP-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The PP-plot consists of the points

$$(1 - \hat{F}(x_{i,n}), 1 - \hat{F}_{spliced}(x_{i,n}))$$

for  $i = 1, \dots, n$  with  $n$  the length of the data,  $x_{i,n}$  the  $i$ -th smallest observation,  $\hat{F}$  the empirical distribution function and  $\hat{F}_{spliced}$  the fitted spliced distribution function. The minus-log version of the PP-plot consists of

$$(-\log(1 - \hat{F}(x_{i,n})), -\log(1 - \hat{F}_{spliced}(x_{i,n}))).$$

Use [SplicePP\\_TB](#) for censored data.

See Reynkens et al. (2017) and Section 4.3.1 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

spp.the	Vector of the theoretical probabilities $1 - \hat{F}_{spliced}(x_{i,n})$ (when <code>log=FALSE</code> ) or $-\log(1 - \hat{F}_{spliced}(x_{i,n}))$ (when <code>log=TRUE</code> ).
spp.emp	Vector of the empirical probabilities $1 - \hat{F}(x_{i,n})$ (when <code>log=FALSE</code> ) or $-\log(1 - \hat{F}(x_{i,n}))$ (when <code>log=TRUE</code> ).

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[SplicePP\\_TB](#), [pSplice](#), [ecdf](#), [SpliceFitPareto](#), [SpliceFitGPD](#), [SpliceECDF](#), [SpliceLL](#), [SpliceQQ](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
```

```

splicefit <- SpliceFitPareto(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and empirical survival function
SpliceECDF(x, X, splicefit)

# Log-log plot with empirical survival function and fitted survival function
SpliceLL(x, X, splicefit)

# PP-plot of empirical survival function and fitted survival function
SplicePP(X, splicefit)

# PP-plot of empirical survival function and
# fitted survival function with log-scales
SplicePP(X, splicefit, log=TRUE)

# Splicing QQ-plot
SpliceQQ(X, splicefit)

## End(Not run)

```

---

SplicePP\_TB

*PP-plot with fitted and Turnbull survival function*


---

### Description

This function plots the fitted survival function of the spliced distribution versus the Turnbull survival function (which is suitable for interval censored data).

### Usage

```
SplicePP_TB(L, U = L, censored, splicefit, x = NULL, log = FALSE, plot = TRUE,
            main = "Splicing PP-plot", ...)
```

### Arguments

**L** Vector of length  $n$  with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.

U	Vector of length $n$ with the upper boundaries of the intervals. By default, they are equal to L.
censored	A logical vector of length $n$ indicating if an observation is censored.
splicefit	A <code>SpliceFit</code> object, e.g. output from <code>SpliceFiticPareto</code> .
x	Vector of points to plot the functions at. When NULL, the default, the empirical quantiles for $1/(n+1), \dots, n/(n+1)$ , obtained using the Turnbull estimator, are used.
log	Logical indicating if minus the logarithms of the survival probabilities are plotted versus each other, default is FALSE.
plot	Logical indicating if the splicing PP-plot should be made, default is TRUE.
main	Title for the plot, default is "Splicing PP-plot".
...	Additional arguments for the plot function, see <code>plot</code> for more details.

### Details

The PP-plot consists of the points

$$(1 - \hat{F}^{TB}(x_i), 1 - \hat{F}_{spliced}(x_i))$$

for  $i = 1, \dots, n$  with  $n$  the length of the data,  $x_i = \hat{Q}^{TB}(p_i)$  where  $p_i = i/(n+1)$ ,  $\hat{Q}^{TB}$  is the quantile function obtained using the Turnbull estimator,  $\hat{F}^{TB}$  the Turnbull estimator for the distribution function and  $\hat{F}_{spliced}$  the fitted spliced distribution function. The minus-log version of the PP-plot consists of

$$(-\log(1 - \hat{F}^{TB}(x_i)), -\log(1 - \hat{F}_{spliced}(x_i))).$$

Right censored data should be entered as L=1 and U=truncupper, and left censored data should be entered as L=trunclower and U=u. The limits trunclower and truncupper are obtained from the `SpliceFit` object.

If the `interval` package is installed, the `icfit` function is used to compute the Turnbull estimator. Otherwise, `survfit.formula` from `survival` is used.

Use `SplicePP` for non-censored data.

See Reynkens et al. (2017) and Section 4.3.2 in Albrecher et al. (2017) for more details.

### Value

A list with following components:

spp.the	Vector of the theoretical probabilities $1 - \hat{F}_{spliced}(x_i)$ (when log=FALSE) or $-\log(1 - \hat{F}_{spliced}(x_i))$ (when log=TRUE).
spp.emp	Vector of the empirical probabilities $1 - \hat{F}^{TB}(x_i)$ (when log=FALSE) or $-\log(1 - \hat{F}^{TB}(x_i))$ (when log=TRUE).

### Author(s)

Tom Reynkens

## References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

## See Also

[SplicePP](#), [pSplice](#), [Turnbull](#), [icfit](#), [SpliceFiticPareto](#), [SpliceTB](#), [SpliceLL\\_TB](#), [SpliceQQ\\_TB](#)

## Examples

```
## Not run:

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Splice ME and Pareto
splicefit <- SpliceFiticPareto(L=Z, U=U, censored=censored, tsplice=quantile(Z,0.9))

x <- seq(0,20,0.1)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and Turnbull survival function
SpliceTB(x, L=Z, U=U, censored=censored, splicefit=splicefit)
```

```

# Log-log plot with Turnbull survival function and fitted survival function
SpliceLL_TB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and fitted survival function
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and
# fitted survival function with log-scales
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit, log=TRUE)

# QQ-plot using Turnbull survival function and fitted survival function
SpliceQQ_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

## End(Not run)

```

---

SpliceQQ

*Splicing quantile plot*


---

### Description

Computes the empirical quantiles of a data vector and the theoretical quantiles of the fitted spliced distribution. These quantiles are then plotted in a splicing QQ-plot with the theoretical quantiles on the  $x$ -axis and the empirical quantiles on the  $y$ -axis.

### Usage

```
SpliceQQ(X, splicefit, p = NULL, plot = TRUE, main = "Splicing QQ-plot", ...)
```

### Arguments

<code>X</code>	Vector of $n$ observations.
<code>splicefit</code>	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> or <a href="#">SpliceFitGPD</a> .
<code>p</code>	Vector of probabilities used in the QQ-plot. If NULL, the default, we take $p$ equal to $1/(n+1), \dots, n/(n+1)$ .
<code>plot</code>	Logical indicating if the quantiles should be plotted in a splicing QQ-plot, default is TRUE.
<code>main</code>	Title for the plot, default is "Splicing QQ-plot".
<code>...</code>	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

This QQ-plot is given by

$$(Q(p_j), \hat{Q}(p_j)),$$

for  $j = 1, \dots, n$  where  $Q$  is the quantile function of the fitted splicing model and  $\hat{Q}$  is the empirical quantile function and  $p_j = j/(n+1)$ .

See Reynkens et al. (2017) and Section 4.3.1 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

sqq.the	Vector of the theoretical quantiles of the fitted spliced distribution.
sqq.emp	Vector of the empirical quantiles from the data.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[SpliceQQ\\_TB](#), [qSplice](#), [SpliceFitPareto](#), [SpliceFitGPD](#), [SpliceECDF](#), [SpliceLL](#), [SplicePP](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

x <- seq(0, 20, 0.01)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and empirical survival function
SpliceECDF(x, X, splicefit)

# Log-log plot with empirical survival function and fitted survival function
```

```

SpliceLL(x, X, splicefit)

# PP-plot of empirical survival function and fitted survival function
SplicePP(X, splicefit)

# PP-plot of empirical survival function and
# fitted survival function with log-scales
SplicePP(X, splicefit, log=TRUE)

# Splicing QQ-plot
SpliceQQ(X, splicefit)

## End(Not run)

```

---

SpliceQQ\_TB

*Splicing quantile plot using Turnbull estimator*


---

### Description

This function plots the fitted quantile function of the spliced distribution versus quantiles based on the Turnbull survival function (which is suitable for interval censored data).

### Usage

```

SpliceQQ_TB(L, U = L, censored, splicefit, p = NULL,
            plot = TRUE, main = "Splicing QQ-plot", ...)

```

### Arguments

L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals. By default, they are equal to L.
censored	A logical vector of length $n$ indicating if an observation is censored.
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFiticPareto</a> .
p	Vector of probabilities used in the QQ-plot. If NULL, the default, we take $p$ equal to $1/(n+1), \dots, n/(n+1)$ .
plot	Logical indicating if the quantiles should be plotted in a splicing QQ-plot, default is TRUE.
main	Title for the plot, default is "Splicing QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.



## Details

This QQ-plot is given by

$$(Q(p_j), \hat{Q}^{TB}(p_j)),$$

for  $j = 1, \dots, n$  where  $Q$  is the quantile function of the fitted splicing model,  $\hat{Q}^{TB}$  the quantile function obtained using the Turnbull estimator and  $p_j = j/(n + 1)$ .

If the **interval** package is installed, the `icfit` function is used to compute the Turnbull estimator. Otherwise, `survfit.formula` from **survival** is used.

Right censored data should be entered as `L=1` and `U=truncupper`, and left censored data should be entered as `L=trunclower` and `U=u`. The limits `trunclower` and `truncupper` are obtained from the `SpliceFit` object.

Use `SpliceQQ` for non-censored data.

See Reynkens et al. (2017) and Section 4.3.2 in Albrecher et al. (2017) for more details.

## Value

A list with following components:

<code>sqq.the</code>	Vector of the theoretical quantiles of the fitted spliced distribution.
<code>sqq.emp</code>	Vector of the empirical quantiles from the data (based on the Turnbull estimator).

## Author(s)

Tom Reynkens

## References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

## See Also

[SpliceQQ](#), [qSplice](#), [Turnbull](#), [icfit](#), [SpliceFit](#), [Pareto](#), [SpliceTB](#), [SplicePP\\_TB](#), [SpliceLL\\_TB](#)

## Examples

```
## Not run:

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)
```

```

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Splice ME and Pareto
splicefit <- SpliceFiticPareto(L=Z, U=U, censored=censored, tsplice=quantile(Z,0.9))

x <- seq(0,20,0.1)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and Turnbull survival function
SpliceTB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# Log-log plot with Turnbull survival function and fitted survival function
SpliceLL_TB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and fitted survival function
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and
# fitted survival function with log-scales
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit, log=TRUE)

# QQ-plot using Turnbull survival function and fitted survival function
SpliceQQ_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

## End(Not run)

```

**Description**

This function plots the fitted survival function of the spliced distribution together with the Turnbull survival function (which is suitable for interval censored data). Moreover,  $100(1 - \alpha)\%$  confidence intervals are added.

**Usage**

```
SpliceTB(x = sort(L), L, U = L, censored, splicefit, alpha = 0.05, ...)
```

**Arguments**

x	Vector of points to plot the functions at. By default we plot it at the points L.
L	Vector of length $n$ with the lower boundaries of the intervals for interval censored data or the observed data for right censored data.
U	Vector of length $n$ with the upper boundaries of the intervals. By default, they are equal to L.
censored	A logical vector of length $n$ indicating if an observation is censored.
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFiticPareto</a> .
alpha	$100(1 - \alpha)\%$ is the confidence level for the confidence intervals. Default is $\alpha = 0.05$ .
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

Right censored data should be entered as  $L=1$  and  $U=\text{truncupper}$ , and left censored data should be entered as  $L=\text{trunclower}$  and  $U=u$ . The limits `trunclower` and `truncupper` are obtained from the [SpliceFit](#) object.

Use [SpliceECDF](#) for non-censored data.

See Reynkens et al. (2017) and Section 4.3.2 in Albrecher et al. (2017) for more details.

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

**See Also**

[SpliceECDF](#), [pSplice](#), [Turnbull](#), [SpliceFiticPareto](#), [SpliceLL\\_TB](#), [SplicePP\\_TB](#), [SpliceQQ\\_TB](#)

**Examples**

```
## Not run:

# Pareto random sample
X <- rpareto(500, shape=2)

# Censoring variable
Y <- rpareto(500, shape=1)

# Observed sample
Z <- pmin(X,Y)

# Censoring indicator
censored <- (X>Y)

# Right boundary
U <- Z
U[censored] <- Inf

# Splice ME and Pareto
splicefit <- SpliceFiticPareto(L=Z, U=U, censored=censored, tsplice=quantile(Z,0.9))

x <- seq(0,20,0.1)

# Plot of spliced CDF
plot(x, pSplice(x, splicefit), type="l", xlab="x", ylab="F(x)")

# Plot of spliced PDF
plot(x, dSplice(x, splicefit), type="l", xlab="x", ylab="f(x)")

# Fitted survival function and Turnbull survival function
SpliceTB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# Log-log plot with Turnbull survival function and fitted survival function
SpliceLL_TB(x, L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and fitted survival function
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

# PP-plot of Turnbull survival function and
# fitted survival function with log-scales
SplicePP_TB(L=Z, U=U, censored=censored, splicefit=splicefit, log=TRUE)

# QQ-plot using Turnbull survival function and fitted survival function
SpliceQQ_TB(L=Z, U=U, censored=censored, splicefit=splicefit)

## End(Not run)
```

stdf

Non-parametric estimators of the STDF

**Description**

Non-parametric estimators of the stable tail dependence function (STDF):  $\hat{l}_k(x)$  and  $\tilde{l}_k(x)$ .

**Usage**

```
stdf(x, k, X, alpha = 0.5)
```

```
stdf2(x, k, X)
```

**Arguments**

**x** A  $d$ -dimensional point to estimate the STDF in.  
**k** Value of the tail index  $k$ .  
**X** A data matrix of dimensions  $n$  by  $d$  with observations in the rows.  
**alpha** The parameter  $\alpha$  of the estimator  $\hat{l}_k(x)$  (stdf), default is 0.5. This argument is not used in stdf2.

**Details**

The stable tail dependence function in  $x$  can be estimated by

$$\hat{l}_k(x) = 1/k \sum_{i=1}^n 1_{\{\exists j \in \{1, \dots, d\}: \hat{F}_j(X_{i,j}) > 1 - k/nx_j\}}$$

with

$$\hat{F}_j(X_{i,j}) = (R_{i,j} - \alpha)/n$$

where  $R_{i,j}$  is the rank of  $X_{i,j}$  among the  $n$  observations in the  $j$ -th dimension:

$$R_{i,j} = \sum_{m=1}^n 1_{\{X_{m,j} \leq X_{i,j}\}}.$$

This estimator is implemented in stdf.

The second estimator is given by

$$\tilde{l}_k(x) = 1/k \sum_{i=1}^n 1_{\{X_{i,1} \geq X_{n-[kx_1]+1,n}^{(1)} \text{ or } \dots \text{ or } X_{i,d} \geq X_{n-[kx_d]+1,n}^{(d)}\}}$$

where  $X_{i,n}^{(j)}$  is the  $i$ -th smallest observation in the  $j$ -th dimension. This estimator is implemented in stdf2.

See Section 4.5 of Beirlant et al. (2016) for more details.

**Value**

stdf returns the estimate  $\hat{l}_k(x)$  and stdf2 returns the estimate  $\tilde{l}_k(x)$ .

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant J., Goegebeur Y., Segers, J. and Teugels, J. (2004). *Statistics of Extremes: Theory and Applications*, Wiley Series in Probability, Wiley, Chichester.

**Examples**

```
# Generate data matrix
X <- cbind(rpareto(100,2), rpareto(100,3))

# Tail index
k <- 20

# Point to evaluate the STDF in
x <- c(2,3)

# First estimate
stdf(x, k, X)

# Second estimate
stdf2(x, k, X)
```

---

tBurr

*The truncated Burr distribution*


---

**Description**

Density, distribution function, quantile function and random generation for the truncated Burr distribution (type XII).

**Usage**

```
dtburr(x, alpha, rho, eta = 1, endpoint = Inf, log = FALSE)
ptburr(x, alpha, rho, eta = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtburr(p, alpha, rho, eta = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtburr(n, alpha, rho, eta = 1, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
alpha	The $\alpha$ parameter of the truncated Burr distribution, a strictly positive number.
rho	The $\rho$ parameter of the truncated Burr distribution, a strictly negative number.
eta	The $\eta$ parameter of the truncated Burr distribution, a strictly positive number. The default value is 1.
endpoint	Endpoint of the truncated Burr distribution. The default value is Inf for which the truncated Burr distribution corresponds to the ordinary Burr distribution.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated Burr distribution is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of the ordinary Burr distribution and  $T$  is the endpoint (truncation point) of the truncated Burr distribution.

**Value**

dtburr gives the density function evaluated in  $x$ , ptburr the CDF evaluated in  $x$  and qtburr the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rtburr returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**See Also**

[Burr, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dtburr(x, alpha=2, rho=-1, endpoint=9), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, ptburr(x, alpha=2, rho=-1, endpoint=9), xlab="x", ylab="CDF", type="l")
```

tExp

*The truncated exponential distribution***Description**

Density, distribution function, quantile function and random generation for the truncated exponential distribution.

**Usage**

```
dtxp(x, rate = 1, endpoint = Inf, log = FALSE)
ptexp(x, rate = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtexp(p, rate = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtexp(n, rate = 1, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
rate	The rate parameter for the exponential distribution, default is 1.
endpoint	Endpoint of the truncated exponential distribution. The default value is Inf for which the truncated exponential distribution corresponds to the ordinary exponential distribution.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated exponential distribution is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of the ordinary exponential distribution and  $T$  is the endpoint (truncation point) of the truncated exponential distribution.

**Value**

dtxp gives the density function evaluated in  $x$ , ptexp the CDF evaluated in  $x$  and qtexp the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rtexp returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.



**See Also**

[Exponential, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dtexp(x, rate = 2, endpoint=5), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, ptexp(x, rate = 2, endpoint=5), xlab="x", ylab="CDF", type="l")
```

---

tFrechet

*The truncated Fréchet distribution*


---

**Description**

Density, distribution function, quantile function and random generation for the truncated Fréchet distribution.

**Usage**

```
dtfrechet(x, shape, loc = 0, scale = 1, endpoint = Inf, log = FALSE)
ptfrechet(x, shape, loc = 0, scale = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtfrechet(p, shape, loc = 0, scale = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtfrechet(n, shape, loc = 0, scale = 1, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
shape	Shape parameter of the Fréchet distribution.
loc	Location parameter of the Fréchet distribution, default is 0.
scale	Scale parameter of the Fréchet distribution, default is 1.
endpoint	Endpoint of the truncated Fréchet distribution. The default value is Inf for which the truncated Fréchet distribution corresponds to the ordinary Fréchet distribution.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated Fréchet distribution is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of an ordinary Fréchet distribution and  $T$  is the endpoint (truncation point) of the truncated Fréchet distribution.

**Value**

`dtfrechet` gives the density function evaluated in  $x$ , `ptfrechet` the CDF evaluated in  $x$  and `qtfrechet` the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

`rtfrechet` returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**See Also**

[Fréchet, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(1, 10, 0.01)
plot(x, dtfrechet(x, shape=2, endpoint=5), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(1, 10, 0.01)
plot(x, ptfrechet(x, shape=2, endpoint=5), xlab="x", ylab="CDF", type="l")
```

---

tGPD

*The truncated generalised Pareto distribution*


---

**Description**

Density, distribution function, quantile function and random generation for the truncated Generalised Pareto Distribution (GPD).

**Usage**

```
dtgpd(x, gamma, mu = 0, sigma, endpoint = Inf, log = FALSE)
ptgpd(x, gamma, mu = 0, sigma, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtgpd(p, gamma, mu = 0, sigma, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtgpd(n, gamma, mu = 0, sigma, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
gamma	The $\gamma$ parameter of the GPD, a real number.
mu	The $\mu$ parameter of the GPD, a strictly positive number. Default is 0.
sigma	The $\sigma$ parameter of the GPD, a strictly positive number.
endpoint	Endpoint of the truncated GPD. The default value is Inf for which the truncated GPD corresponds to the ordinary GPD.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated GPD is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of the ordinary GPD and  $T$  is the endpoint (truncation point) of the truncated GPD.

**Value**

dtgpd gives the density function evaluated in  $x$ , ptgpd the CDF evaluated in  $x$  and qtgpd the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rtgpd returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens

**See Also**

[tGPD, Pareto, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dtgpd(x, gamma=1/2, sigma=5, endpoint=8), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, ptgpd(x, gamma=1/2, sigma=5, endpoint=8), xlab="x", ylab="CDF", type="l")
```

tlnorm

*The truncated log-normal distribution***Description**

Density, distribution function, quantile function and random generation for the truncated log-normal distribution.

**Usage**

```
dtlnorm(x, meanlog = 0, sdlog = 1, endpoint = Inf, log = FALSE)
ptlnorm(x, meanlog = 0, sdlog = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtlnorm(p, meanlog = 0, sdlog = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtlnorm(n, meanlog = 0, sdlog = 1, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
meanlog	Mean of the distribution on the log scale, default is 0.
sdlog	Standard deviation of the distribution on the log scale, default is 1.
endpoint	Endpoint of the truncated log-normal distribution. The default value is Inf for which the truncated log-normal distribution corresponds to the ordinary log-normal distribution.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated log-normal distribution is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of the ordinary log-normal distribution and  $T$  is the endpoint (truncation point) of the truncated log-normal distribution.

**Value**

dtlnorm gives the density function evaluated in  $x$ , ptlnorm the CDF evaluated in  $x$  and qtlnorm the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rtlnorm returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**See Also**[Lognormal, Distributions](#)**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dtlnorm(x, endpoint=9), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, ptlnorm(x, endpoint=9), xlab="x", ylab="CDF", type="l")
```

tPareto

*The truncated Pareto distribution***Description**

Density, distribution function, quantile function and random generation for the truncated Pareto distribution.

**Usage**

```
dtpareto(x, shape, scale = 1, endpoint = Inf, log = FALSE)
ptpareto(x, shape, scale = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtpareto(p, shape, scale = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtpareto(n, shape, scale = 1, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
shape	The shape parameter of the truncated Pareto distribution, a strictly positive number.
scale	The scale parameter of the truncated Pareto distribution, a strictly positive number. Its default value is 1.
endpoint	Endpoint of the truncated Pareto distribution. The default value is Inf for which the truncated Pareto distribution corresponds to the ordinary Pareto distribution.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated Pareto distribution is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of an ordinary Pareto distribution and  $T$  is the endpoint (truncation point) of the truncated Pareto distribution.

**Value**

dtpareto gives the density function evaluated in  $x$ , ptpareto the CDF evaluated in  $x$  and qtpareto the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rtpareto returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens

**See Also**

[Pareto, Distributions](#)

**Examples**

```
# Plot of the PDF
x = seq(1,10,0.01)
plot(x, dtpareto(x, shape=2, endpoint=10), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x = seq(1,10,0.01)
plot(x, ptpareto(x, shape=2, endpoint=10), xlab="x", ylab="CDF", type="l")
```

---

trDT

*Truncation odds*


---

**Description**

Estimates of truncation odds of the truncated probability mass under the untruncated distribution using truncated Hill.

**Usage**

```
trDT(data, r = 1, gamma, plot = FALSE, add = FALSE, main = "Estimates of DT", ...)
```

**Arguments**

**data** Vector of  $n$  observations.

**r** Trimming parameter, default is 1 (no trimming).

**gamma** Vector of  $n - 1$  estimates for the EVI obtained from [trHill](#).

plot	Logical indicating if the estimates of $D_T$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $D_T$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of DT".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The truncation odds is defined as

$$D_T = (1 - F(T))/F(T)$$

with  $T$  the upper truncation point and  $F$  the CDF of the untruncated distribution (e.g. Pareto distribution).

We estimate this truncation odds as

$$\hat{D}_T = \max\{(k+1)/(n+1)(R_{r,k,n}^{1/\gamma_k} - 1/(k+1))/(1 - R_{r,k,n}^{1/\gamma_k}), 0\}$$

with  $R_{r,k,n} = X_{n-k,n}/X_{n-r+1,n}$ .

See Beirlant et al. (2016) or Section 4.2.3 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
DT	Vector of the corresponding estimates for the truncation odds $D_T$ .

### Author(s)

Tom Reynkens based on R code of Dries Cornilly.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

### See Also

[trHill](#), [trEndpoint](#), [trQuant](#), [trDTMLE](#)

**Examples**

```
# Sample from truncated Pareto distribution.
# truncated at 99% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.99, shape=shape))

# Truncated Hill estimator
trh <- trHill(X, plot=TRUE, ylim=c(0,2))

# Truncation odds
dt <- trDT(X, gamma=trh$gamma, plot=TRUE, ylim=c(0,0.05))
```

trDTMLE

*Truncation odds***Description**

Estimates of truncation odds of the truncated probability mass under the untruncated distribution using truncated MLE.

**Usage**

```
trDTMLE(data, gamma, tau, plot = FALSE, add = FALSE, main = "Estimates of DT", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trMLE</a> .
tau	Vector of $n - 1$ estimates for the $\tau$ obtained from <a href="#">trMLE</a> .
plot	Logical indicating if the estimates of $D_T$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $D_T$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of DT".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The truncation odds is defined as

$$D_T = (1 - F(T))/F(T)$$

with  $T$  the upper truncation point and  $F$  the CDF of the untruncated distribution (e.g. GPD).

We estimate this truncation odds as

$$\hat{D}_T = \max\{(k+1)/(n+1)((1 + \hat{\tau}_k E_{1,k})^{-1/\hat{\xi}_k} - 1/(k+1))/(1 - (1 + \hat{\tau}_k E_{1,k})^{-1/\hat{\xi}_k}), 0\}$$

with  $E_{1,k} = X_{n,n} - X_{n-k,n}$ .

See Beirlant et al. (2017) for more details.



**Value**

A list with following components:

`k`                    Vector of the values of the tail parameter  $k$ .  
`DT`                    Vector of the corresponding estimates for the truncation odds  $D_T$ .

**Author(s)**

Tom Reynkens.

**References**

Beirlant, J., Fraga Alves, M. I. and Reynkens, T. (2017). "Fitting Tails Affected by Truncation". *Electronic Journal of Statistics*, 11(1), 2026–2065.

**See Also**

[trMLE](#), [trEndpointMLE](#), [trProbMLE](#), [trQuantMLE](#), [trTestMLE](#), [trDT](#)

**Examples**

```
# Sample from GPD truncated at 99% quantile
gamma <- 0.5
sigma <- 1.5
X <- rtgpd(n=250, gamma=gamma, sigma=sigma, endpoint=qgpd(0.99, gamma=gamma, sigma=sigma))

# Truncated ML estimator
trmle <- trMLE(X, plot=TRUE, ylim=c(0,2))

# Truncation odds
dtmle <- trDTMLE(X, gamma=trmle$gamma, tau=trmle$tau, plot=TRUE, ylim=c(0,0.05))
```

---

<code>trEndpoint</code>	<i>Estimator of endpoint</i>
-------------------------	------------------------------

---

**Description**

Estimator of endpoint using truncated Hill estimates.

**Usage**

```
trEndpoint(data, r = 1, gamma, plot = FALSE, add = FALSE,
           main = "Estimates of endpoint", ...)
```

**Arguments**

data	Vector of $n$ observations.
r	Trimming parameter, default is 1 (no trimming).
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trHill</a> .
plot	Logical indicating if the estimates of $T$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $T$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of endpoint".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The endpoint is estimated as

$$\hat{T}_{k,n} = \max\{X_{n-k,n}(((X_{n-k,n}/X_{n,n})^{1/\hat{\gamma}_k} - 1/(k+1))/(1 - 1/(k+1)))^{-\hat{\gamma}_k}, X_{n,n}\}$$

with  $\hat{\gamma}_k$  the Hill estimates adapted for truncation.

See Beirlant et al. (2016) or Section 4.2.3 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Tk	Vector of the corresponding estimates for the endpoint $T$ .

**Author(s)**

Tom Reynkens based on R code of Dries Cornilly.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

**See Also**

[trHill](#), [trDT](#), [trEndpointMLE](#)

**Examples**

```
# Sample from truncated Pareto distribution.
# truncated at 99% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.99, shape=shape))

# Truncated Hill estimator
trh <- trHill(X, plot=TRUE, ylim=c(0,2))

# Endpoint
trEndpoint(X, gamma=trh$gamma, plot=TRUE, ylim=c(8,12))
abline(h=qpareto(0.99, shape=shape), lty=2)
```

---

trEndpointMLE	<i>Estimator of endpoint</i>
---------------	------------------------------

---

**Description**

Estimator of endpoint using truncated ML estimates.

**Usage**

```
trEndpointMLE(data, gamma, tau, plot = FALSE, add = FALSE,
              main = "Estimates of endpoint", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trMLE</a> .
tau	Vector of $n - 1$ estimates for the $\tau$ obtained from <a href="#">trMLE</a> .
plot	Logical indicating if the estimates of $T$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $T$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of endpoint".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The endpoint is estimated as

$$\hat{T}_k = X_{n-k,n} + 1/\hat{\tau}_k [((1 - 1/k)/((1 + \hat{\tau}_k(X_{n,n} - X_{n-k,n}))^{-1/\hat{\xi}_k} - 1/k))^{\hat{\xi}_k} - 1]$$

with  $\hat{\gamma}_k$  and  $\hat{\tau}_k$  the truncated ML estimates for  $\gamma$  and  $\tau$ .

See Beirlant et al. (2017) for more details.

**Value**

A list with following components:

`k`                    Vector of the values of the tail parameter  $k$ .  
`Tk`                    Vector of the corresponding estimates for the endpoint  $T$ .

**Author(s)**

Tom Reynkens.

**References**

Beirlant, J., Fraga Alves, M. I. and Reynkens, T. (2017). "Fitting Tails Affected by Truncation". *Electronic Journal of Statistics*, 11(1), 2026–2065.

**See Also**

[trMLE](#), [trDTMLE](#), [trProbMLE](#), [trQuantMLE](#), [trTestMLE](#), [trEndpoint](#)

**Examples**

```
# Sample from GPD truncated at 99% quantile
gamma <- 0.5
sigma <- 1.5
X <- rtgpd(n=250, gamma=gamma, sigma=sigma, endpoint=qgpd(0.99, gamma=gamma, sigma=sigma))

# Truncated ML estimator
trmle <- trMLE(X, plot=TRUE, ylim=c(0,2))

# Endpoint
trEndpointMLE(X, gamma=trmle$gamma, tau=trmle$tau, plot=TRUE, ylim=c(0,50))
abline(h=qgpd(0.99, gamma=gamma, sigma=sigma), lty=2)
```

---

trHill

*Hill estimator for upper truncated data*


---

**Description**

Computes the Hill estimator for positive extreme value indices, adapted for upper truncation, as a function of the tail parameter  $k$  (Aban et al. 2006; Beirlant et al., 2016). Optionally, these estimates are plotted as a function of  $k$ .

**Usage**

```
trHill(data, r = 1, tol = 1e-08, maxiter = 100, logk = FALSE,
       plot = FALSE, add = FALSE, main = "Estimates of the EVI", ...)
```

**Arguments**

data	Vector of $n$ observations.
r	Trimming parameter, default is 1 (no trimming).
tol	Numerical tolerance for stopping criterion used in Newton-Raphson iterations, default is $1e-08$ .
maxiter	Maximum number of Newton-Raphson iterations, default is 100.
logk	Logical indicating if the estimates are plotted as a function of $\log(k)$ (logk=TRUE) or as a function of $k$ . Default is FALSE.
plot	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The truncated Hill estimator is the MLE for  $\gamma$  under the truncated Pareto distribution.

To estimate the EVI using the truncated Hill estimator an equation needs to be solved. Beirlant et al. (2016) propose to use Newton-Raphson iterations to solve this equation. We take the trimmed Hill estimates as starting values for this algorithm. The trimmed Hill estimator is defined as

$$H_{r,k,n} = 1/(k - r + 1) \sum_{j=r}^k \log(X_{n-j+1,n}) - \log(X_{n-k,n})$$

for  $1 \leq r < k < n$  and is a basic extension of the Hill estimator for upper truncated data (the ordinary Hill estimator is obtained for  $r = 1$ ).

The equation that needs to be solved is

$$H_{r,k,n} = \gamma + R_{r,k,n}^{1/\gamma} \log(R_{r,k,n}) / (1 - R_{r,k,n}^{1/\gamma})$$

with  $R_{r,k,n} = X_{n-k,n} / X_{n-r+1,n}$ .

See Beirlant et al. (2016) or Section 4.2.3 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding estimates for $\gamma$ .
H	Vector of corresponding trimmed Hill estimates.

**Author(s)**

Tom Reynkens based on R code of Dries Cornilly.

## References

- Aban, I.B., Meerschaert, M.M. and Panorska, A.K. (2006). "Parameter Estimation for the Truncated Pareto Distribution." *Journal of the American Statistical Association*, 101, 270–277.
- Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.
- Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

## See Also

[Hill](#), [trDT](#), [trEndpoint](#), [trProb](#), [trQuant](#), [trMLE](#)

## Examples

```
# Sample from truncated Pareto distribution.
# truncated at 99% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.99, shape=shape))

# Truncated Hill estimator
trh <- trHill(X, plot=TRUE, ylim=c(0,2))
```

---

trMLE

*MLE estimator for upper truncated data*


---

## Description

Computes the ML estimator for the extreme value index, adapted for upper truncation, as a function of the tail parameter  $k$  (Beirlant et al., 2017). Optionally, these estimates are plotted as a function of  $k$ .

## Usage

```
trMLE(data, start = c(1, 1), eps = 10−10,
       plot = TRUE, add = FALSE, main = "Estimates for EVI", ...)
```

## Arguments

data	Vector of $n$ observations.
start	Starting values for $\gamma$ and $\tau$ for the numerical optimisation.
eps	Numerical tolerance, see Details. By default it is equal to $10^{-10}$ .
plot	Logical indicating if the estimates of $\gamma$ should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates of $\gamma$ should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of the EVI".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

## Details

We compute the MLE for the  $\gamma$  and  $\sigma$  parameters of the truncated GPD. For numerical reasons, we compute the MLE for  $\tau = \gamma/\sigma$  and transform this estimate to  $\sigma$ .

The log-likelihood is given by

$$(k-1) \ln \tau - (k-1) \ln \xi - (1 + 1/\xi) \sum_{j=2}^k \ln(1 + \tau E_{j,k}) - (k-1) \ln(1 - (1 + \tau E_{1,k})^{-1/\xi})$$

with  $E_{j,k} = X_{n-j+1,n} - X_{n-k,n}$ .

In order to meet the restrictions  $\sigma = \xi/\tau > 0$  and  $1 + \tau E_{j,k} > 0$  for  $j = 1, \dots, k$ , we require the estimates of these quantities to be larger than the numerical tolerance value eps.

See Beirlant et al. (2017) for more details.

## Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
gamma	Vector of the corresponding estimates for $\gamma$ .
tau	Vector of the corresponding estimates for $\tau$ .
sigma	Vector of the corresponding estimates for $\sigma$ .
conv	Convergence indicator of <code>optim</code> .

## Author(s)

Tom Reynkens.

## References

Beirlant, J., Fraga Alves, M. I. and Reynkens, T. (2017). "Fitting Tails Affected by Truncation". *Electronic Journal of Statistics*, 11(1), 2026–2065.

## See Also

[trDTMLE](#), [trEndpointMLE](#), [trProbMLE](#), [trQuantMLE](#), [trTestMLE](#), [trHill](#), [GPDmle](#)

## Examples

```
# Sample from GPD truncated at 99% quantile
gamma <- 0.5
sigma <- 1.5
X <- rtgpd(n=250, gamma=gamma, sigma=sigma, endpoint=qgpd(0.99, gamma=gamma, sigma=sigma))

# Truncated ML estimator
trmle <- trMLE(X, plot=TRUE, ylim=c(0,2))
```

trParetoQQ

*Truncated Pareto quantile plot***Description**

Extension of the Pareto QQ-plot as described in Beirlant et al. (2016).

**Usage**

```
trParetoQQ(data, r = 1, DT, kstar = NULL, plot = TRUE, main = "TPa QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
r	Trimming parameter, default is 1 (no trimming).
DT	Vector of $n - 1$ estimates for the truncation odds $D_T$ obtained from <a href="#">trDT</a> .
kstar	Value for $k$ used to construct the plot. When NULL (default), a value will be chosen by maximising the correlation between the empirical and theoretical quantiles (see Details).
plot	Logical indicating if the quantiles should be plotted in a Pareto QQ-plot, default is TRUE.
main	Title for the plot, default is "TPa QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The Pareto QQ-plot for truncated data plots

$$(-\log(\hat{D}_{T,r,k^*,n} + j/(n+1)), \log(X_{n-j+1,n}))$$

for  $j = 1, \dots, n$ .

The value for  $k^*$  can be given by the user or can be determined automatically. In the latter case, we use the  $k^*$  that maximises the absolute value of the correlation between  $-\log(\hat{D}_{T,r,k^*,n} + j/(n+1))$  and  $\log(X_{n-j+1,n})$  for  $j = 1, \dots, k$  and  $k^* > 10$ .

When taking  $D_T = 0$ , one obtains the ordinary Pareto QQ-plot.

Note that the definition here differs slightly from the one in Beirlant et al. (2016). We plot the empirical and theoretical quantiles the other way around and therefore have to add a minus (before the log).

See Beirlant et al. (2016) for more details.



**Value**

A list with following components:

pqq.the	Vector of theoretical quantiles $-\log(\hat{D}_{T,r,k^*,n} + j/(n+1))$ , see Details.
pqq.emp	Vector of the empirical quantiles from the log-transformed data.
kstar	Optimal value for $k$ or input argument kstar, see Details.
DTstar	Estimate of $D_T$ corresponding to kstar.

**Author(s)**

Tom Reynkens.

**References**

Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

**See Also**

[ParetoQQ](#), [trDT](#)

**Examples**

```
# Endpoint of truncated Pareto distribution
endpoint <- qpareto(0.99, shape=2)

# Generate sample from truncated Pareto distribution
X <- rtpareto(1000, shape=2, endpoint=endpoint)

# Ordinary Pareto QQ-plot
ParetoQQ(X)

# Truncated Hill estimates
gamma <- trHill(X)$gamma

# Estimates for truncation odds
dt <- trDT(X, gamma=gamma)$DT

# Truncated Pareto QQ-plot
trParetoQQ(X, DT=dt)
```

trProb

*Estimator of small exceedance probabilities using truncated Hill***Description**

Computes estimates of a small exceedance probability  $P(X > q)$  using the estimates for the EVI obtained from the Hill estimator adapted for upper truncation.

**Usage**

```
trProb(data, r = 1, gamma, q, warnings = TRUE, plot = FALSE, add = FALSE,
       main = "Estimates of small exceedance probability", ...)
```

**Arguments**

data	Vector of $n$ observations.
r	Trimming parameter, default is 1 (no trimming).
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trHill</a> .
q	The used large quantile (we estimate $P(X > q)$ for $q$ large).
warnings	Logical indicating if warnings are shown, default is TRUE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The probability is estimated as

$$\hat{P}(X > q) = (k + 1)/(n + 1)((q/X_{n-k,n})^{-1/\gamma_k} - R_{r,k,n}^{1/\hat{\gamma}_k})/(1 - R_{r,k,n}^{1/\hat{\gamma}_k})$$

with  $R_{r,k,n} = X_{n-k,n}/X_{n-r+1,n}$  and  $\hat{\gamma}_k$  the Hill estimates adapted for truncation.

See Beirlant et al. (2016) or Section 4.2.3 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates.
q	The used large quantile.

**Author(s)**

Tom Reynkens based on R code of Dries Cornilly.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

**See Also**

[trHill](#), [trQuant](#), [Prob](#), [trProbMLE](#)

**Examples**

```
# Sample from truncated Pareto distribution.
# truncated at 99% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.99, shape=shape))

# Truncated Hill estimator
trh <- trHill(X, plot=TRUE, ylim=c(0,2))

# Small probability
trProb(X, gamma=trh$gamma, q=8, plot=TRUE)
```

---

trProbMLE

---

*Estimator of small exceedance probabilities using truncated MLE*


---

**Description**

Computes estimates of a small exceedance probability  $P(X > q)$  using the estimates for the EVI obtained from the ML estimator adapted for upper truncation.

**Usage**

```
trProbMLE(data, gamma, tau, DT, q, plot = FALSE, add = FALSE,
          main = "Estimates of small exceedance probability", ...)
```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trMLE</a> .
tau	Vector of $n - 1$ estimates for the $\tau$ obtained from <a href="#">trMLE</a> .
DT	Vector of $n - 1$ estimates for the truncation odds obtained from <a href="#">trDTMLE</a> .
q	The used large quantile (we estimate $P(X > q)$ for $q$ large).

plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of small exceedance probability".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

The probability is estimated as

$$\hat{p}_{T,k}(q) = (1 + \hat{D}_{T,k})(k + 1)/(n + 1)(1 + \hat{\tau}_k(q - X_{n-k,n}))^{-1/\hat{\xi}_k} - \hat{D}_{T,k}$$

with  $\hat{\gamma}_k$  and  $\hat{\tau}_k$  the ML estimates adapted for truncation and  $\hat{D}_T$  the estimates for the truncation odds.

See Beirlant et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
P	Vector of the corresponding probability estimates.
q	The used large quantile.

### Author(s)

Tom Reynkens.

### References

Beirlant, J., Fraga Alves, M. I. and Reynkens, T. (2017). "Fitting Tails Affected by Truncation". *Electronic Journal of Statistics*, 11(1), 2026–2065.

### See Also

[trMLE](#), [trDTMLE](#), [trQuantMLE](#), [trEndpointMLE](#), [trTestMLE](#), [trProb](#), [Prob](#)

### Examples

```
# Sample from GPD truncated at 99% quantile
gamma <- 0.5
sigma <- 1.5
X <- rtgpd(n=250, gamma=gamma, sigma=sigma, endpoint=qgpd(0.99, gamma=gamma, sigma=sigma))

# Truncated ML estimator
trmle <- trMLE(X, plot=TRUE, ylim=c(0,2))

# Truncation odds
```

```
dtmle <- trDTMLE(X, gamma=trmle$gamma, tau=trmle$tau, plot=FALSE)

# Small exceedance probability
trProbMLE(X, gamma=trmle$gamma, tau=trmle$tau, DT=dtmle$DT, plot=TRUE, q=26, ylim=c(0,0.005))
```

---

trQuant *Estimator of large quantiles using truncated Hill*

---

### Description

trQuant computes estimates of large quantiles  $Q(1 - p)$  of the truncated distribution using the estimates for the EVI obtained from the Hill estimator adapted for upper truncation. trQuantW computes estimates of large quantiles  $Q_W(1 - p)$  of the parent distribution  $W$  which is unobserved.

### Usage

```
trQuant(data, r = 1, rough = TRUE, gamma, DT, p, plot = FALSE, add = FALSE,
        main = "Estimates of extreme quantile", ...)

trQuantW(data, gamma, DT, p, plot = FALSE, add = FALSE,
         main = "Estimates of extreme quantile", ...)
```

### Arguments

data	Vector of $n$ observations (truncated data).
r	Trimming parameter, default is 1 (no trimming).
rough	Logical indicating if rough truncation is present, default is TRUE.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trHill</a> .
DT	Vector of $n - 1$ estimates for the truncation odds obtained from <a href="#">trDT</a> .
p	The exceedance probability of the quantile (we estimate $Q(1 - p)$ for $p$ small).
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

We observe the truncated r.v.  $X =_d W|W < T$  where  $T$  is the truncation point and  $W$  the untruncated r.v.

Under rough truncation, the quantiles for  $X$  are estimated using

$$\hat{Q}(1 - p) = X_{n-k,n}((\hat{D}_T + (k + 1)/(n + 1))/(\hat{D}_T + p))^{\hat{\gamma}^k},$$

with  $\hat{\gamma}_k$  the Hill estimates adapted for truncation and  $\hat{D}_T$  the estimates for the truncation odds.

Under light truncation, the quantiles are estimated using the Weissman estimator with the Hill estimates replaced by the truncated Hill estimates:

$$\hat{Q}(1-p) = X_{n-k,n}((k+1)/((n+1)p))^{\hat{\gamma}_k}.$$

To decide between light and rough truncation, one can use the test implemented in [trTest](#).

The quantiles for  $W$  are estimated using

$$\hat{Q}_W(1-p) = X_{n-k,n}((\hat{D}_T + (k+1)/(n+1))/(p(1 + \hat{D}_T))^{\hat{\gamma}_k}.$$

See Beirlant et al. (2016) or Section 4.2.3 of Albrecher et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

### Author(s)

Tom Reynkens based on R code of Dries Cornilly.

### References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

### See Also

[trHill](#), [trDT](#), [trProb](#), [trEndpoint](#), [trTest](#), [Quant](#), [trQuantMLE](#)

### Examples

```
# Sample from truncated Pareto distribution.
# truncated at 99% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.99, shape=shape))

# Truncated Hill estimator
trh <- trHill(X, plot=TRUE, ylim=c(0,2))

# Truncation odds
dt <- trDT(X, gamma=trh$gamma, plot=TRUE, ylim=c(0,2))

# Large quantile
```

```

p <- 10^(-5)
# Truncated distribution
trQuant(X, gamma=trh$gamma, DT=dt$DT, p=p, plot=TRUE)
# Original distribution
trQuantW(X, gamma=trh$gamma, DT=dt$DT, p=p, plot=TRUE, ylim=c(0,1000))

```

trQuantMLE

*Estimator of large quantiles using truncated MLE***Description**

This function computes estimates of large quantiles  $Q(1-p)$  of the truncated distribution using the ML estimates adapted for upper truncation. Moreover, estimates of large quantiles  $Q_Y(1-p)$  of the original distribution  $Y$ , which is unobserved, are also computed.

**Usage**

```

trQuantMLE(data, gamma, tau, DT, p, Y = FALSE, plot = FALSE, add = FALSE,
            main = "Estimates of extreme quantile", ...)

```

**Arguments**

data	Vector of $n$ observations.
gamma	Vector of $n-1$ estimates for the EVI obtained from <a href="#">trMLE</a> .
tau	Vector of $n-1$ estimates for the $\tau$ obtained from <a href="#">trMLE</a> .
DT	Vector of $n-1$ estimates for the truncation odds obtained from <a href="#">trDTMLE</a> .
p	The exceedance probability of the quantile (we estimate $Q(1-p)$ or $Q_Y(1-p)$ for $p$ small).
Y	Logical indicating if quantiles from the truncated distribution ( $Q(1-p)$ ) or from the parent distribution ( $Q_Y(1-p)$ ) are computed. Default is TRUE.
plot	Logical indicating if the estimates should be plotted as a function of $k$ , default is FALSE.
add	Logical indicating if the estimates should be added to an existing plot, default is FALSE.
main	Title for the plot, default is "Estimates of extreme quantile".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

We observe the truncated r.v.  $X =_d Y|Y < T$  where  $T$  is the truncation point and  $Y$  the untruncated r.v.

Under rough truncation, the quantiles for  $X$  are estimated using

$$\hat{Q}_{T,k}(1-p) = X_{n-k,n} + 1/(\hat{\tau}_k) \left[ \left( \frac{\hat{D}_{T,k} + (k+1)/(n+1)}{\hat{D}_{T,k} + p} \right)^{\hat{\xi}_k} - 1 \right],$$

with  $\hat{\gamma}_k$  and  $\hat{\tau}_k$  the ML estimates adapted for truncation and  $\hat{D}_T$  the estimates for the truncation odds.

The quantiles for  $Y$  are estimated using

$$\hat{Q}_{Y,k}(1-p) = X_{n-k,n} + 1/(\hat{\tau}_k) \left( \left[ \frac{\hat{D}_{T,k} + (k+1)/(n+1)}{p(\hat{D}_{T,k} + 1)} \right]^{\hat{\xi}_k} - 1 \right).$$

See Beirlant et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
Q	Vector of the corresponding quantile estimates.
p	The used exceedance probability.

### Author(s)

Tom Reynkens.

### References

Beirlant, J., Fraga Alves, M. I. and Reynkens, T. (2017). "Fitting Tails Affected by Truncation". *Electronic Journal of Statistics*, 11(1), 2026–2065.

### See Also

[trMLE](#), [trDTMLE](#), [trProbMLE](#), [trEndpointMLE](#), [trTestMLE](#), [trQuant](#), [Quant](#)

### Examples

```
# Sample from GPD truncated at 99% quantile
gamma <- 0.5
sigma <- 1.5
X <- rtgpd(n=250, gamma=gamma, sigma=sigma, endpoint=qgpd(0.99, gamma=gamma, sigma=sigma))

# Truncated ML estimator
trmle <- trMLE(X, plot=TRUE, ylim=c(0,2))

# Truncation odds
dtmle <- trDTMLE(X, gamma=trmle$gamma, tau=trmle$tau, plot=FALSE)

# Large quantile of X
trQuantMLE(X, gamma=trmle$gamma, tau=trmle$tau, DT=dtmle$DT, plot=TRUE, p=0.005, ylim=c(15,30))

# Large quantile of Y
trQuantMLE(X, gamma=trmle$gamma, tau=trmle$tau, DT=dtmle$DT, plot=TRUE, p=0.005, ylim=c(0,300),
            Y=TRUE)
```



---

trTest *Test for truncated Pareto-type tails*

---

### Description

Test between non-truncated Pareto-type tails (*light truncation*) and truncated Pareto-type tails (*rough truncation*).

### Usage

```
trTest(data, alpha = 0.05, plot = TRUE, main = "Test for truncation", ...)
```

### Arguments

data	Vector of $n$ observations.
alpha	The used significance level, default is 0.05.
plot	Logical indicating if the P-values should be plotted as a function of $k$ , default is FALSE.
main	Title for the plot, default is "Test for truncation".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

We want to test  $H_0 : X$  has non-truncated Pareto tails vs.  $H_1 : X$  has truncated Pareto tails. Let

$$E_{k,n}(\gamma) = 1/k \sum_{j=1}^k (X_{n-k,n}/X_{n-j+1,n})^{1/\gamma},$$

with  $X_{i,n}$  the  $i$ -th order statistic. The test statistic is then

$$T_{k,n} = \sqrt{12k}(E_{k,n}(H_{k,n}) - 1/2)/(1 - E_{k,n}(H_{k,n}))$$

which is asymptotically standard normally distributed. We reject  $H_0$  on level  $\alpha$  if

$$T_{k,n} < -z_\alpha$$

where  $z_\alpha$  is the  $100(1 - \alpha)\%$  quantile of a standard normal distribution. The corresponding P-value is thus given by

$$\Phi(T_{k,n})$$

with  $\Phi$  the CDF of a standard normal distribution.

See Beirlant et al. (2016) or Section 4.2.3 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

k	Vector of the values of the tail parameter $k$ .
testVal	Corresponding test values.
critVal	Critical value used for the test, i.e. $qnorm(1-\alpha/2)$ .
Pval	Corresponding P-values.
Reject	Logical vector indicating if the null hypothesis is rejected for a certain value of $k$ .

**Author(s)**

Tom Reynkens.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Beirlant, J., Fraga Alves, M.I. and Gomes, M.I. (2016). "Tail fitting for Truncated and Non-truncated Pareto-type Distributions." *Extremes*, 19, 429–462.

**See Also**

[trHill](#), [trTestMLE](#)

**Examples**

```
# Sample from truncated Pareto distribution.
# truncated at 95% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.95, shape=shape))

# Test for truncation
trTest(X)

# Sample from truncated Pareto distribution.
# truncated at 99% quantile
shape <- 2
X <- rtpareto(n=1000, shape=shape, endpoint=qpareto(0.99, shape=shape))

# Test for truncation
trTest(X)
```

---

trTestMLE	<i>Test for truncated GPD tails</i>
-----------	-------------------------------------

---

### Description

Test between non-truncated GPD tails (*light truncation*) and truncated GPD tails (*rough truncation*).

### Usage

```
trTestMLE(data, gamma, tau, alpha = 0.05, plot = TRUE, main = "Test for truncation", ...)
```

### Arguments

data	Vector of $n$ observations.
gamma	Vector of $n - 1$ estimates for the EVI obtained from <a href="#">trMLE</a> .
tau	Vector of $n - 1$ estimates for the $\tau$ obtained from <a href="#">trMLE</a> .
alpha	The used significance level, default is $0.05$ .
plot	Logical indicating if the P-values should be plotted as a function of $k$ , default is FALSE.
main	Title for the plot, default is "Test for truncation".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

### Details

We want to test  $H_0 : X$  has non-truncated GPD tails vs.  $H_1 : X$  has truncated GPD tails. Let  $\hat{\gamma}_k$  and  $\hat{\tau}_k$  be the truncated MLE estimates for  $\gamma$  and  $\tau$ . The test statistic is then

$$T_{k,n} = k(1 + \hat{\tau}(X_{n,n} - X_{-k,n}))^{-1/\hat{\xi}_k}$$

which is asymptotically standard exponentially distributed. We reject  $H_0$  on level  $\alpha$  if  $T_{k,n} > \ln(1/\alpha)$ . The corresponding P-value is given by  $\exp(-T_{k,n})$ .

See Beirlant et al. (2017) for more details.

### Value

A list with following components:

k	Vector of the values of the tail parameter $k$ .
testVal	Corresponding test values.
critVal	Critical value used for the test, i.e. $\ln(1/\alpha)$ .
Pval	Corresponding P-values.
Reject	Logical vector indicating if the null hypothesis is rejected for a certain value of $k$ .

**Author(s)**

Tom Reynkens.

**References**

Beirlant, J., Fraga Alves, M. I. and Reynkens, T. (2017). "Fitting Tails Affected by Truncation". *Electronic Journal of Statistics*, 11(1), 2026–2065.

**See Also**

[trMLE](#), [trDTMLE](#), [trProbMLE](#), [trEndpointMLE](#), [trTestMLE](#), [trTest](#)

**Examples**

```
# Sample from GPD truncated at 99% quantile
gamma <- 0.5
sigma <- 1.5
X <- rtgpd(n=250, gamma=gamma, sigma=sigma, endpoint=qgpd(0.99, gamma=gamma, sigma=sigma))

# Truncated ML estimator
trmle <- trMLE(X, plot=TRUE, ylim=c(0,2))

# Test for truncation
trTestMLE(X, gamma=trmle$gamma, tau=trmle$tau)
```

---

Turnbull

*Turnbull estimator*

---

**Description**

Computes the Turnbull estimator for the survival function of interval censored data.

**Usage**

```
Turnbull(x, L, R, censored, trunclower = 0, truncupper = Inf,
         conf.type = "plain", conf.int = 0.95)
```

**Arguments**

x	Vector with points to evaluate the estimator in.
L	Vector of length $n$ with the lower boundaries of the intervals.
R	Vector of length $n$ with the upper boundaries of the intervals.
censored	Vector of $n$ logicals indicating if an observation is interval censored.
trunclower	Lower truncation point, default is 0.
truncupper	Upper truncation point, default is Inf.
conf.type	Type of confidence interval, see <a href="#">survfit.formula</a> . Default is "plain".
conf.int	Confidence level of the two-sided confidence interval, see <a href="#">survfit.formula</a> . Default is 0.95.

**Details**

We consider the random interval censoring model where one observes  $L \leq R$  and where the variable of interest  $X$  lies between  $L$  and  $R$ .

Right censored data should be entered as  $L=1$  and  $R=\text{truncupper}$ , and right censored data should be entered as  $L=\text{trunclower}$  and  $R=r$ .

This function calls `survfit.formula` from **survival**.

See Section 4.3.2 in Albrecher et al. (2017) for more details.

**Value**

A list with following components:

<code>surv</code>	A vector of length <code>length(x)</code> containing the Turnbull estimator evaluated in the elements of <code>x</code> .
<code>fit</code>	The output from the call to <code>survfit.formula</code> , an object of class <code>survfit</code> .

**Author(s)**

Tom Reynkens

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Turnbull, B. W. (1974). "Nonparametric Estimation of a Survivorship Function with Doubly Censored Data." *Journal of the American Statistical Association*, 69, 169–173.

Turnbull, B. W. (1976). "The Empirical Distribution Function with Arbitrarily Grouped, Censored and Truncated Data." *Journal of the Royal Statistical Society: Series B (Methodological)*, 38, 290–295.

**See Also**

[survfit.formula](#), [KaplanMeier](#)

**Examples**

```
L <- 1:10
R <- c(1, 2.5, 3, 4, 5.5, 6, 7.5, 8.25, 9, 10.5)
censored <- c(0, 1, 0, 0, 1, 0, 1, 1, 0, 1)

x <- seq(0, 12, 0.1)

# Turnbull estimator
plot(x, Turnbull(x, L, R, censored)$cdf, type="s", ylab="Turnbull estimator")
```

tWeibull

*The truncated Weibull distribution***Description**

Density, distribution function, quantile function and random generation for the truncated Weibull distribution.

**Usage**

```
dtweibull(x, shape, scale = 1, endpoint = Inf, log = FALSE)
ptweibull(x, shape, scale = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
qtweibull(p, shape, scale = 1, endpoint = Inf, lower.tail = TRUE, log.p = FALSE)
rtweibull(n, shape, scale = 1, endpoint = Inf)
```

**Arguments**

x	Vector of quantiles.
p	Vector of probabilities.
n	Number of observations.
shape	The shape parameter of the Weibull distribution, a strictly positive number.
scale	The scale parameter of the Weibull distribution, a strictly positive number, default is 1.
endpoint	Endpoint of the truncated Weibull distribution. The default value is Inf for which the truncated Weibull distribution corresponds to the ordinary Weibull distribution.
log	Logical indicating if the densities are given as $\log(f)$ , default is FALSE.
lower.tail	Logical indicating if the probabilities are of the form $P(X \leq x)$ (TRUE) or $P(X > x)$ (FALSE). Default is TRUE.
log.p	Logical indicating if the probabilities are given as $\log(p)$ , default is FALSE.

**Details**

The Cumulative Distribution Function (CDF) of the truncated Weibull distribution is equal to  $F_T(x) = F(x)/F(T)$  for  $x \leq T$  where  $F$  is the CDF of the ordinary Weibull distribution and  $T$  is the endpoint (truncation point) of the truncated Weibull distribution.

**Value**

dtweibull gives the density function evaluated in  $x$ , ptweibull the CDF evaluated in  $x$  and qtweibull the quantile function evaluated in  $p$ . The length of the result is equal to the length of  $x$  or  $p$ .

rtweibull returns a random sample of length  $n$ .

**Author(s)**

Tom Reynkens.

**See Also**

[Weibull, Distributions](#)

**Examples**

```
# Plot of the PDF
x <- seq(0, 10, 0.01)
plot(x, dtweibull(x, shape=2, scale=0.5, endpoint=1), xlab="x", ylab="PDF", type="l")

# Plot of the CDF
x <- seq(0, 10, 0.01)
plot(x, ptweibull(x, shape=2, scale=0.5, endpoint=1), xlab="x", ylab="CDF", type="l")
```

---

 VaR

*VaR of splicing fit*


---

**Description**

Compute Value-at-Risk ( $VaR_{1-p} = Q(1-p)$ ) of the fitted spliced distribution.

**Usage**

```
VaR(p, splicefit)
```

**Arguments**

p	The exceedance probability (we estimate $VaR_{1-p} = Q(1-p)$ ).
splicefit	A <a href="#">SpliceFit</a> object, e.g. output from <a href="#">SpliceFitPareto</a> , <a href="#">SpliceFiticPareto</a> or <a href="#">SpliceFitGPD</a> .

**Details**

See Reynkens et al. (2017) and Section 4.6 of Albrecher et al. (2017) for details.

Note that `VaR(p, splicefit)` corresponds to `qSplice(p, splicefit, lower.tail = FALSE)`.

**Value**

Vector of quantiles  $VaR_{1-p} = Q(1-p)$ .

**Author(s)**

Tom Reynkens with R code from Roel Verbelen for the mixed Erlang quantiles.

## References

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

Reynkens, T., Verbelen, R., Beirlant, J. and Antonio, K. (2017). "Modelling Censored Losses Using Splicing: a Global Fit Strategy With Mixed Erlang and Extreme Value Distributions". *Insurance: Mathematics and Economics*, 77, 65–77.

Verbelen, R., Gong, L., Antonio, K., Badescu, A. and Lin, S. (2015). "Fitting Mixtures of Erlangs to Censored and Truncated Data Using the EM Algorithm." *Astin Bulletin*, 45, 729–758

## See Also

[qSplice](#), [CTE](#), [SpliceFit](#), [SpliceFitPareto](#), [SpliceFiticPareto](#), [SpliceFitGPD](#)

## Examples

```
## Not run:

# Pareto random sample
X <- rpareto(1000, shape = 2)

# Splice ME and Pareto
splicefit <- SpliceFitPareto(X, 0.6)

p <- seq(0,1,0.01)

# Plot of quantiles
plot(p, qSplice(p, splicefit), type="l", xlab="p", ylab="Q(p)")

# Plot of VaR
plot(p, VaR(p, splicefit), type="l", xlab="p", ylab=bquote(VaR[1-p]))

## End(Not run)
```

---

WeibullQQ

*Weibull quantile plot*

---

## Description

Computes the empirical quantiles of the log-transform of a data vector and the theoretical quantiles of the standard Weibull distribution. These quantiles are then plotted in a Weibull QQ-plot with the theoretical quantiles on the  $x$ -axis and the empirical quantiles on the  $y$ -axis.

## Usage

```
WeibullQQ(data, plot = TRUE, main = "Weibull QQ-plot", ...)
```



**Arguments**

data	Vector of $n$ observations.
plot	Logical indicating if the quantiles should be plotted in a Weibull QQ-plot, default is TRUE.
main	Title for the plot, default is "Weibull QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The Weibull QQ-plot is given by

$$(\log(-\log(1 - i/(n + 1))), \log X_{i,n})$$

for  $i = 1, \dots, n$ , with  $X_{i,n}$  the  $i$ -th order statistic of the data.

See Section 4.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

wqq.the	Vector of the theoretical quantiles from a standard Weibull distribution.
wqq.emp	Vector of the empirical quantiles from the log-transformed data.

**Author(s)**

Tom Reynkens.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[WeibullQQ\\_der](#), [ExpQQ](#), [LognormalQQ](#), [ParetoQQ](#)

**Examples**

```
data(norwegianfire)

# Weibull QQ-plot for Norwegian Fire Insurance data for claims in 1976.
WeibullQQ(norwegianfire$size[norwegianfire$year==76])

# Derivative of Weibull QQ-plot for Norwegian Fire Insurance data for claims in 1976.
WeibullQQ_der(norwegianfire$size[norwegianfire$year==76])
```

WeibullQQ\_der

*Derivative plot of the Weibull QQ-plot***Description**

Computes the derivative plot of the Weibull QQ-plot. These values can be plotted as a function of the data or as a function of the tail parameter  $k$ .

**Usage**

```
WeibullQQ_der(data, k = FALSE, plot = TRUE,
              main = "Derivative plot of Weibull QQ-plot", ...)
```

**Arguments**

data	Vector of $n$ observations.
plot	Logical indicating if the derivative values should be plotted, default is TRUE.
k	Logical indicating if the derivative values are plotted as a function of the tail parameter $k$ (k=TRUE) or as a function of the logarithm of the data (k=FALSE). Default is FALSE.
main	Title for the plot, default is "Derivative plot of Weibull QQ-plot".
...	Additional arguments for the plot function, see <a href="#">plot</a> for more details.

**Details**

The derivative plot of a Weibull QQ-plot is

$$(k, H_{k,n}/W_{k,n})$$

or

$$(\log X_{n-k,n}, H_{k,n}/W_{k,n})$$

with  $H_{k,n}$  the Hill estimates and

$$W_{k,n} = 1/k \sum_{j=1}^k \log(\log((n+1)/j)) - \log(\log((n+1)/(k+1))).$$

See Section 4.1 of Albrecher et al. (2017) for more details.

**Value**

A list with following components:

xval	Vector of the x-values of the plot ( $k$ or $\log X_{n-k,n}$ ).
yval	Vector of the derivative values.

**Author(s)**

Tom Reynkens.

**References**

Albrecher, H., Beirlant, J. and Teugels, J. (2017). *Reinsurance: Actuarial and Statistical Aspects*, Wiley, Chichester.

**See Also**

[WeibullQQ](#), [Hill](#), [MeanExcess](#), [LognormalQQ\\_der](#), [ParetoQQ\\_der](#)

**Examples**

```
data(norwegianfire)

# Weibull QQ-plot for Norwegian Fire Insurance data for claims in 1976.
WeibullQQ(norwegianfire$size[norwegianfire$year==76])

# Derivative of Weibull QQ-plot for Norwegian Fire Insurance data for claims in 1976.
WeibullQQ_der(norwegianfire$size[norwegianfire$year==76])
```

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