

# Package ‘sjstats’

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**Title** Collection of Convenient Functions for Common Statistical Computations

**Version** 0.19.1

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**Description** Collection of convenient functions for common statistical computations, which are not directly provided by R's base or stats packages. This package aims at providing, first, shortcuts for statistical measures, which otherwise could only be calculated with additional effort (like Cramer's V, Phi, or effect size statistics like Eta or Omega squared), or for which currently no functions available. Second, another focus lies on weighted variants of common statistical measures and tests like weighted standard error, mean, t-test, correlation, and more.

**License** GPL-3

**Depends** R (>= 4.0), utils

**Imports** datawizard, effectsize (>= 0.8.8), insight, parameters, performance, stats

**Suggests** brms, car, coin, ggplot2, lme4, MASS, pscl, pwr, survey, testthat

**URL** <https://strengjacke.github.io/sjstats/>

**BugReports** <https://github.com/strengjacke/sjstats/issues>

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## Contents

anova_stats	2
auto_prior	3
bootstrap	5
boot_ci	6
chisq_gof	8
chi_squared_test	10
cramers_v	13
cv	16
cv_error	16
design_effect	17
efc	18
find_beta	19
gmd	21
inequ_trend	22
is_prime	23
kruskal_wallis_test	24
mann_whitney_test	26
nhanes_sample	29
prop	29
r2	31
samplesize_mixed	32
se_ybar	34
survey_median	34
svyglm.nb	36
svyglm.zip	37
table_values	38
t_test	39
var_pop	42
weight	43
wilcoxon_test	44
<b>Index</b>	<b>47</b>

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anova_stats	<i>Effect size statistics for anova</i>
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### Description

Returns the (partial) eta-squared, (partial) omega-squared, epsilon-squared statistic or Cohen's F for all terms in an anovas. `anova_stats()` returns a tidy summary, including all these statistics and power for each term.

### Usage

```
anova_stats(model, digits = 3)
```

**Arguments**

model            A fitted anova-model of class aov or anova. Other models are coerced to [anova](#).  
 digits           Amount of digits for returned values.

**Value**

A data frame with all statistics is returned (excluding confidence intervals).

**References**

Levine TR, Hullett CR (2002): Eta Squared, Partial Eta Squared, and Misreporting of Effect Size in Communication Research.

Tippey K, Longnecker MT (2016): An Ad Hoc Method for Computing Pseudo-Effect Size for Mixed Model.

**Examples**

```
# load sample data
data(efc)

# fit linear model
fit <- aov(
  c12hour ~ as.factor(e42dep) + as.factor(c172code) + c160age,
  data = efc
)
anova_stats(car::Anova(fit, type = 2))
```

---

 auto\_prior

---

*Create default priors for brms-models*


---

**Description**

This function creates default priors for brms-regression models, based on the same automatic prior-scale adjustment as in **rstanarm**.

**Usage**

```
auto_prior(formula, data, gaussian, locations = NULL)
```

**Arguments**

formula           A formula describing the model, which just needs to contain the model terms, but no notation of interaction, splines etc. Usually, you want only those predictors in the formula, for which automatic priors should be generated. Add informative priors afterwards to the returned `brmsprior`-object.

data	The data that will be used to fit the model.
gaussian	Logical, if the outcome is gaussian or not.
locations	A numeric vector with location values for the priors. If <code>locations = NULL</code> , <code>0</code> is used as location parameter.

### Details

`auto_prior()` is a small, convenient function to create some default priors for brms-models with automatically adjusted prior scales, in a similar way like **rstanarm** does. The default scale for the intercept is 10, for coefficients 2.5. If the outcome is gaussian, both scales are multiplied with `sd(y)`. Then, for categorical variables, nothing more is changed. For numeric variables, the scales are divided by the standard deviation of the related variable.

All prior distributions are *normal* distributions. `auto_prior()` is intended to quickly create default priors with feasible scales. If more precise definitions of priors is necessary, this needs to be done directly with brms-functions like `set_prior()`.

### Value

A `brmsprior`-object.

### Note

As `auto_prior()` also sets priors on the intercept, the model formula used in `brms::brm()` must be rewritten to something like `y ~ 0 + intercept ...`, see [set\\_prior](#).

### Examples

```
data(efc)
efc$c172code <- as.factor(efc$c172code)
efc$c161sex <- as.factor(efc$c161sex)

mf <- formula(neg_c_7 ~ c161sex + c160age + c172code)
auto_prior(mf, efc, TRUE)

## compare to
# m <- rstanarm::stan_glm(mf, data = efc, chains = 2, iter = 200)
# ps <- prior_summary(m)
# ps$prior_intercept$adjusted_scale
# ps$prior$adjusted_scale

## usage
# ap <- auto_prior(mf, efc, TRUE)
# brm(mf, data = efc, prior = ap)

# add informative priors
mf <- formula(neg_c_7 ~ c161sex + c172code)

auto_prior(mf, efc, TRUE) +
  brms::prior(normal(.1554, 40), class = "b", coef = "c160age")
```

```
# example with binary response
efc$neg_c_7d <- ifelse(efc$neg_c_7 < median(efc$neg_c_7, na.rm = TRUE), 0, 1)
mf <- formula(neg_c_7d ~ c161sex + c160age + c172code + e17age)
auto_prior(mf, efc, FALSE)
```

bootstrap

*Generate nonparametric bootstrap replications***Description**

Generates  $n$  bootstrap samples of data and returns the bootstrapped data frames as list-variable.

**Usage**

```
bootstrap(data, n, size)
```

**Arguments**

data	A data frame.
n	Number of bootstraps to be generated.
size	Optional, size of the bootstrap samples. May either be a number between 1 and <code>nrow(data)</code> or a value between 0 and 1 to sample a proportion of observations from data (see 'Examples').

**Details**

By default, each bootstrap sample has the same number of observations as data. To generate bootstrap samples without resampling same observations (i.e. sampling without replacement), use `size` to get bootstrapped data with a specific number of observations. However, specifying the `size`-argument is much less memory-efficient than the bootstrap with replacement. Hence, it is recommended to ignore the `size`-argument, if it is not really needed.

**Value**

A data frame with one column: a list-variable `strap`, which contains resample-objects of class `sj_resample`. These resample-objects are lists with three elements:

1. the original data frame, `data`
2. the rownumbers `id`, i.e. rownumbers of `data`, indicating the resampled rows with replacement
3. the `resample.id`, indicating the index of the resample (i.e. the position of the `sj_resample`-object in the list `strap`)

**Note**

This function applies nonparametric bootstrapping, i.e. the function draws samples with replacement.

There is an `as.data.frame`- and a `print`-method to get or print the resampled data frames. See 'Examples'. The `as.data.frame`- method automatically applies whenever coercion is done because a data frame is required as input. See 'Examples' in `boot_ci`.

**See Also**

`boot_ci` to calculate confidence intervals from bootstrap samples.

**Examples**

```
data(efc)
bs <- bootstrap(efc, 5)

# now run models for each bootstrapped sample
lapply(bs$strap, function(x) lm(neg_c_7 ~ e42dep + c161sex, data = x))

# generate bootstrap samples with 600 observations for each sample
bs <- bootstrap(efc, 5, 600)

# generate bootstrap samples with 70% observations of the original sample size
bs <- bootstrap(efc, 5, .7)

# compute standard error for a simple vector from bootstraps
# use the `as.data.frame()`-method to get the resampled
# data frame
bs <- bootstrap(efc, 100)
bs$c12hour <- unlist(lapply(bs$strap, function(x) {
  mean(as.data.frame(x)$c12hour, na.rm = TRUE)
}))

# bootstrapped standard error
boot_se(bs, "c12hour")

# bootstrapped CI
boot_ci(bs, "c12hour")
```

---

boot\_ci

*Standard error and confidence intervals for bootstrapped estimates*

---

**Description**

Compute nonparametric bootstrap estimate, standard error, confidence intervals and p-value for a vector of bootstrap replicate estimates.

**Usage**

```
boot_ci(data, select = NULL, method = c("dist", "quantile"), ci.lvl = 0.95)
```

```
boot_se(data, select = NULL)
```

```
boot_p(data, select = NULL)
```

```
boot_est(data, select = NULL)
```

**Arguments**

<code>data</code>	A data frame that contains the vector with bootstrapped estimates, or directly the vector (see 'Examples').
<code>select</code>	Optional, unquoted names of variables (as character vector) with bootstrapped estimates. Required, if either <code>data</code> is a data frame (and no vector), and only selected variables from <code>data</code> should be processed.
<code>method</code>	Character vector, indicating if confidence intervals should be based on bootstrap standard error, multiplied by the value of the quantile function of the t-distribution (default), or on sample quantiles of the bootstrapped values. See 'Details' in <code>boot_ci()</code> . May be abbreviated.
<code>ci.lvl</code>	Numeric, the level of the confidence intervals.

**Details**

The methods require one or more vectors of bootstrap replicate estimates as input.

- `boot_est()`: returns the bootstrapped estimate, simply by computing the mean value of all bootstrap estimates.
- `boot_se()`: computes the nonparametric bootstrap standard error by calculating the standard deviation of the input vector.
- The mean value of the input vector and its standard error is used by `boot_ci()` to calculate the lower and upper confidence interval, assuming a t-distribution of bootstrap estimate replicates (for `method = "dist"`, the default, which is  $\text{mean}(x) \pm \text{qt}(.975, \text{df} = \text{length}(x) - 1) * \text{sd}(x)$ ); for `method = "quantile"`, 95\confidence intervals ( $\text{quantile}(x, \text{probs} = \text{c}(0.025, 0.975))$ ). Use `ci.lvl` to change the level for the confidence interval.
- P-values from `boot_p()` are also based on t-statistics, assuming normal distribution.

**Value**

A data frame with either bootstrap estimate, standard error, the lower and upper confidence intervals or the p-value for all bootstrapped estimates.

**References**

Carpenter J, Bithell J. Bootstrap confidence intervals: when, which, what? A practical guide for medical statisticians. *Statist. Med.* 2000; 19:1141-1164

**See Also**

[`bootstrap()`] to generate nonparametric bootstrap samples.

**Examples**

```
data(efc)
bs <- bootstrap(efc, 100)

# now run models for each bootstrapped sample
bs$models <- lapply(
  bs$strap,
  function(.x) lm(neg_c_7 ~ e42dep + c161sex, data = .x)
)

# extract coefficient "dependency" and "gender" from each model
bs$dependency <- vapply(bs$models, function(x) coef(x)[2], numeric(1))
bs$gender <- vapply(bs$models, function(x) coef(x)[3], numeric(1))

# get bootstrapped confidence intervals
boot_ci(bs$dependency)

# compare with model fit
fit <- lm(neg_c_7 ~ e42dep + c161sex, data = efc)
confint(fit)[2, ]

# alternative function calls.
boot_ci(bs$dependency)
boot_ci(bs, "dependency")
boot_ci(bs, c("dependency", "gender"))
boot_ci(bs, c("dependency", "gender"), method = "q")

# compare coefficients
mean(bs$dependency)
boot_est(bs$dependency)
coef(fit)[2]
```

---

chisq\_gof

*Compute model quality*


---

**Description**

For logistic regression models, performs a Chi-squared goodness-of-fit-test.

**Usage**

```
chisq_gof(x, prob = NULL, weights = NULL)
```



**Arguments**

x	A numeric vector or a glm-object.
prob	Vector of probabilities (indicating the population probabilities) of the same length as x's amount of categories / factor levels. Use <code>nrow(table(x))</code> to determine the amount of necessary values for prob. Only used, when x is a vector, and not a glm-object.
weights	Vector with weights, used to weight x.

**Details**

For vectors, this function is a convenient function for the `chisq.test()`, performing goodness-of-fit test. For glm-objects, this function performs a goodness-of-fit test. A well-fitting model shows *no* significant difference between the model and the observed data, i.e. the reported p-values should be greater than 0.05.

**Value**

For vectors, returns the object of the computed `chisq.test`. For glm-objects, an object of class `chisq_gof` with following values: `p.value`, the p-value for the goodness-of-fit test; `z.score`, the standardized z-score for the goodness-of-fit test; `rss`, the residual sums of squares term and `chisq`, the pearson chi-squared statistic.

**References**

Hosmer, D. W., & Lemeshow, S. (2000). Applied Logistic Regression. Hoboken, NJ, USA: John Wiley & Sons, Inc.

**Examples**

```
data(efc)
efc$neg_c_7d <- ifelse(efc$neg_c_7 < median(efc$neg_c_7, na.rm = TRUE), 0, 1)
m <- glm(
  neg_c_7d ~ c161sex + barthtot + c172code,
  data = efc,
  family = binomial(link = "logit")
)

# goodness-of-fit test for logistic regression
chisq_gof(m)

# goodness-of-fit test for vectors against probabilities
# differing from population
chisq_gof(efc$e42dep, c(0.3,0.2,0.22,0.28))

# equal to population
chisq_gof(efc$e42dep, prop.table(table(efc$e42dep)))
```

---

chi\_squared\_test      *Chi-Squared test*

---

### Description

This function performs a  $\chi^2$  test for contingency tables or tests for given probabilities. The returned effects sizes are Cramer's V for tables with more than two rows or columns, Phi ( $\phi$ ) for 2x2 tables, and Fei (*Fei*) for tests against given probabilities (see *Ben-Shachar et al. 2023*).

### Usage

```
chi_squared_test(
  data,
  select = NULL,
  by = NULL,
  probabilities = NULL,
  weights = NULL,
  paired = FALSE,
  ...
)
```

### Arguments

data	A data frame.
select	Name(s) of the continuous variable(s) (as character vector) to be used as samples for the test. <code>select</code> can be one of the following: <ul style="list-style-type: none"> <li>• <code>select</code> can be used in combination with <code>by</code>, in which case <code>select</code> is the name of the continuous variable (and <code>by</code> indicates a grouping factor).</li> <li>• <code>select</code> can also be a character vector of length two or more (more than two names only apply to <code>kruskal_wallis_test()</code>), in which case the two continuous variables are treated as samples to be compared. <code>by</code> must be <code>NULL</code> in this case.</li> <li>• If <code>select</code> is of length <b>two</b> and <code>paired = TRUE</code>, the two samples are considered as <i>dependent</i> and a paired test is carried out.</li> <li>• If <code>select</code> specifies <b>one</b> variable and <code>by = NULL</code>, a one-sample test is carried out (only applicable for <code>t_test()</code> and <code>wilcoxon_test()</code>)</li> <li>• For <code>chi_squared_test()</code>, if <code>select</code> specifies <b>one</b> variable and both <code>by</code> and <code>probabilities</code> are <code>NULL</code>, a one-sample test against given probabilities is automatically conducted, with equal probabilities for each level of <code>select</code>.</li> </ul>
by	Name of the variable indicating the groups. Required if <code>select</code> specifies only one variable that contains all samples to be compared in the test. If <code>by</code> is not a factor, it will be coerced to a factor. For <code>chi_squared_test()</code> , if <code>probabilities</code> is provided, <code>by</code> must be <code>NULL</code> .

probabilities	A numeric vector of probabilities for each cell in the contingency table. The length of the vector must match the number of cells in the table, i.e. the number of unique levels of the variable specified in <code>select</code> . If probabilities is provided, a chi-squared test for given probabilities is conducted. Furthermore, if probabilities is given, <code>by</code> must be <code>NULL</code> . The probabilities must sum to 1.
weights	Name of an (optional) weighting variable to be used for the test.
paired	Logical, if <code>TRUE</code> , a McNemar test is conducted for 2x2 tables. Note that <code>paired</code> only works for 2x2 tables.
...	Additional arguments passed down to <code>chisq.test()</code> .

### Details

The function is a wrapper around `chisq.test()` and `fisher.test()` (for small expected values) for contingency tables, and `chisq.test()` for given probabilities. When probabilities are provided, these are rescaled to sum to 1 (i.e. `rescale.p = TRUE`). When `fisher.test()` is called, simulated p-values are returned (i.e. `simulate.p.value = TRUE`, see `?fisher.test`). If `paired = TRUE` and a 2x2 table is provided, a McNemar test (see `mcnemar.test()`) is conducted.

The weighted version of the chi-squared test is based on the a weighted table, using `xtabs()` as input for `chisq.test()`.

Interpretation of effect sizes are based on rules described in `effectsize::interpret_phi()`, `effectsize::interpret_cramers_v()`, and `effectsize::interpret_fei()`. Use these function directly to get other interpretations, by providing the returned effect size as argument, e.g. `interpret_phi(0.35, rules = "gignac2016")`.

### Value

A data frame with test results. The returned effects sizes are Cramer's V for tables with more than two rows or columns, Phi ( $\phi$ ) for 2x2 tables, and Fei ( $Fei$ ) for tests against given probabilities.

### Which test to use

The following table provides an overview of which test to use for different types of data. The choice of test depends on the scale of the outcome variable and the number of samples to compare.

Samples	Scale of Outcome	Significance Test
1	binary / nominal	<code>chi_squared_test()</code>
1	continuous, not normal	<code>wilcoxon_test()</code>
1	continuous, normal	<code>t_test()</code>
2, independent	binary / nominal	<code>chi_squared_test()</code>
2, independent	continuous, not normal	<code>mann_whitney_test()</code>
2, independent	continuous, normal	<code>t_test()</code>
2, dependent	binary (only 2x2)	<code>chi_squared_test(paired=TRUE)</code>
2, dependent	continuous, not normal	<code>wilcoxon_test()</code>
2, dependent	continuous, normal	<code>t_test(paired=TRUE)</code>
>2, independent	continuous, not normal	<code>kruskal_wallis_test()</code>
>2, independent	continuous, normal	<code>datawizard::means_by_group()</code>
>2, dependent	continuous, not normal	<i>not yet implemented</i> (1)

>2, dependent      continuous, normal      *not yet implemented* (2)

(1) More than two dependent samples are considered as *repeated measurements*. For ordinal or not-normally distributed outcomes, these samples are usually tested using a `friedman.test()`, which requires the samples in one variable, the groups to compare in another variable, and a third variable indicating the repeated measurements (subject IDs).

(2) More than two dependent samples are considered as *repeated measurements*. For normally distributed outcomes, these samples are usually tested using a ANOVA for repeated measurements. A more sophisticated approach would be using a linear mixed model.

## References

- Ben-Shachar, M.S., Patil, I., Thériault, R., Wiernik, B.M., Lüdtke, D. (2023). Phi, Fei, Fo, Fum: Effect Sizes for Categorical Data That Use the Chi-Squared Statistic. *Mathematics*, 11, 1982. doi:10.3390/math11091982
- Bender, R., Lange, S., Ziegler, A. Wichtige Signifikanztests. *Dtsch Med Wochenschr* 2007; 132: e24–e25
- du Prel, J.B., Röhrig, B., Hommel, G., Blettner, M. Auswahl statistischer Testverfahren. *Dtsch Arztebl Int* 2010; 107(19): 343–8

## See Also

- `t_test()` for parametric t-tests of dependent and independent samples.
- `mann_whitney_test()` for non-parametric tests of unpaired (independent) samples.
- `wilcoxon_test()` for Wilcoxon rank sum tests for non-parametric tests of paired (dependent) samples.
- `kruskal_wallis_test()` for non-parametric tests with more than two independent samples.
- `chi_squared_test()` for chi-squared tests (two categorical variables, dependent and independent).

## Examples

```
data(efc)
efc$weight <- abs(rnorm(nrow(efc), 1, 0.3))

# Chi-squared test
chi_squared_test(efc, "c161sex", by = "e16sex")

# weighted Chi-squared test
chi_squared_test(efc, "c161sex", by = "e16sex", weights = "weight")

# Chi-squared test for given probabilities
chi_squared_test(efc, "c161sex", probabilities = c(0.3, 0.7))
```

---

`cramers_v`*Measures of association for contingency tables*

---

## Description

This function calculates various measure of association for contingency tables and returns the statistic and p-value. Supported measures are Cramer's V, Phi, Spearman's rho, Kendall's tau and Pearson's r.

## Usage

```
cramers_v(tab, ...)
```

```
cramer(tab, ...)
```

```
## S3 method for class 'formula'  
cramers_v(  
  formula,  
  data,  
  ci.lvl = NULL,  
  n = 1000,  
  method = c("dist", "quantile"),  
  ...  
)
```

```
phi(tab, ...)
```

```
crosstable_statistics(  
  data,  
  x1 = NULL,  
  x2 = NULL,  
  statistics = c("auto", "cramer", "phi", "spearman", "kendall", "pearson", "fisher"),  
  weights = NULL,  
  ...  
)
```

```
xtab_statistics(  
  data,  
  x1 = NULL,  
  x2 = NULL,  
  statistics = c("auto", "cramer", "phi", "spearman", "kendall", "pearson", "fisher"),  
  weights = NULL,  
  ...  
)
```

**Arguments**

tab	A <code>table()</code> or <code>fable()</code> . Tables of class <code>xtabs()</code> and other will be coerced to <code>fable</code> objects.
...	Other arguments, passed down to the statistic functions <code>chisq.test()</code> , <code>fisher.test()</code> or <code>cor.test()</code> .
formula	A formula of the form <code>lhs ~ rhs</code> where <code>lhs</code> is a numeric variable giving the data values and <code>rhs</code> a factor giving the corresponding groups.
data	A data frame or a table object. If a table object, <code>x1</code> and <code>x2</code> will be ignored. For Kendall's <i>tau</i> , Spearman's <i>rho</i> or Pearson's product moment correlation coefficient, <code>data</code> needs to be a data frame. If <code>x1</code> and <code>x2</code> are not specified, the first two columns of the data frames are used as variables to compute the crosstab.
ci.lvl	Scalar between 0 and 1. If not <code>NULL</code> , returns a data frame including lower and upper confidence intervals.
n	Number of bootstraps to be generated.
method	Character vector, indicating if confidence intervals should be based on bootstrap standard error, multiplied by the value of the quantile function of the t-distribution (default), or on sample quantiles of the bootstrapped values. See 'Details' in <code>boot.ci()</code> . May be abbreviated.
x1	Name of first variable that should be used to compute the contingency table. If <code>data</code> is a table object, this argument will be ignored.
x2	Name of second variable that should be used to compute the contingency table. If <code>data</code> is a table object, this argument will be ignored.
statistics	Name of measure of association that should be computed. May be one of "auto", "cramer", "phi", "spearman", "kendall", "pearson" or "fisher". See 'Details'.
weights	Name of variable in <code>x</code> that indicated the vector of weights that will be applied to weight all observations. Default is <code>NULL</code> , so no weights are used.

**Details**

The p-value for Cramer's V and the Phi coefficient are based on `chisq.test()`. If any expected value of a table cell is smaller than 5, or smaller than 10 and the `df` is 1, then `fisher.test()` is used to compute the p-value, unless `statistics = "fisher"`; in this case, the use of `fisher.test()` is forced to compute the p-value. The test statistic is calculated with `cramers_v()` resp. `phi()`.

Both test statistic and p-value for Spearman's rho, Kendall's tau and Pearson's r are calculated with `cor.test()`.

When `statistics = "auto"`, only Cramer's V or Phi are calculated, based on the dimension of the table (i.e. if the table has more than two rows or columns, Cramer's V is calculated, else Phi).

**Value**

For `phi()`, the table's Phi value. For `[cramers_v()]`, the table's Cramer's V.

For `crosstable_statistics()`, a list with following components:

- `estimate`: the value of the estimated measure of association.

- `p.value`: the p-value for the test.
- `statistic`: the value of the test statistic.
- `stat.name`: the name of the test statistic.
- `stat.html`: if applicable, the name of the test statistic, in HTML-format.
- `df`: the degrees of freedom for the contingency table.
- `method`: character string indicating the name of the measure of association.
- `method.html`: if applicable, the name of the measure of association, in HTML-format.
- `method.short`: the short form of association measure, equals the `statistics`-argument.
- `fisher`: logical, if Fisher's exact test was used to calculate the p-value.

## References

Ben-Shachar, M.S., Patil, I., Thériault, R., Wiernik, B.M., Lüdtke, D. (2023). Phi, Fei, Fo, Fum: Effect Sizes for Categorical Data That Use the Chi-Squared Statistic. *Mathematics*, 11, 1982. [doi:10.3390/math11091982](https://doi.org/10.3390/math11091982)

## Examples

```
# Phi coefficient for 2x2 tables
tab <- table(sample(1:2, 30, TRUE), sample(1:2, 30, TRUE))
phi(tab)

# Cramer's V for nominal variables with more than 2 categories
tab <- table(sample(1:2, 30, TRUE), sample(1:3, 30, TRUE))
cramer(tab)

# formula notation
data(efc)
cramer(e16sex ~ c161sex, data = efc)

# bootstrapped confidence intervals
cramer(e16sex ~ c161sex, data = efc, ci.lvl = .95, n = 100)

# 2x2 table, compute Phi automatically
crosstable_statistics(efc, e16sex, c161sex)

# more dimensions than 2x2, compute Cramer's V automatically
crosstable_statistics(efc, c172code, c161sex)

# ordinal data, use Kendall's tau
crosstable_statistics(efc, e42dep, quol_5, statistics = "kendall")

# calculate Spearman's rho, with continuity correction
crosstable_statistics(efc,
  e42dep,
  quol_5,
  statistics = "spearman",
  exact = FALSE,
  continuity = TRUE
)
```

---

cv *Compute model quality*

---

### Description

Compute the coefficient of variation.

### Usage

```
cv(x, ...)
```

### Arguments

x Fitted linear model of class `lm`, `merMod` (**lme4**) or `lme` (**nlme**).  
... More fitted model objects, to compute multiple coefficients of variation at once.

### Details

The advantage of the `cv` is that it is unitless. This allows coefficient of variation to be compared to each other in ways that other measures, like standard deviations or root mean squared residuals, cannot be.

### Value

Numeric, the coefficient of variation.

### Examples

```
data(efc)
fit <- lm(barthtot ~ c160age + c12hour, data = efc)
cv(fit)
```

---

cv\_error *Test and training error from model cross-validation*

---

### Description

`cv_error()` computes the root mean squared error from a model fitted to kfold cross-validated test-training-data. `cv_compare()` does the same, for multiple formulas at once (by calling `cv_error()` for each formula).

### Usage

```
cv_error(data, formula, k = 5)
```

```
cv_compare(data, formulas, k = 5)
```



**Arguments**

data	A data frame.
formula	The formula to fit the linear model for the test and training data.
k	The number of folds for the kfold-crossvalidation.
formulas	A list of formulas, to fit linear models for the test and training data.

**Details**

`cv_error()` first generates cross-validated test-training pairs, using `crossv_kfold` and then fits a linear model, which is described in `formula`, to the training data. Then, predictions for the test data are computed, based on the trained models. The *training error* is the mean value of the `rmse` for all *trained* models; the *test error* is the `rmse` based on all residuals from the test data.

**Value**

A data frame with the root mean squared errors for the training and test data.

**Examples**

```
data(efc)
cv_error(efc, neg_c_7 ~ barthtot + c161sex)

cv_compare(efc, formulas = list(
  neg_c_7 ~ barthtot + c161sex,
  neg_c_7 ~ barthtot + c161sex + e42dep,
  neg_c_7 ~ barthtot + c12hour
))
```

---

 design\_effect

*Design effects for two-level mixed models*


---

**Description**

Compute the design effect (also called *Variance Inflation Factor*) for mixed models with two-level design.

**Usage**

```
design_effect(n, icc = 0.05)
```

**Arguments**

n	Average number of observations per grouping cluster (i.e. level-2 unit).
icc	Assumed intraclass correlation coefficient for multilevel-model.

**Details**

The formula for the design effect is simply  $(1 + (n - 1) * icc)$ .

**Value**

The design effect (Variance Inflation Factor) for the two-level model.

**References**

Bland JM. 2000. Sample size in guidelines trials. *Fam Pract.* (17), 17-20.

Hsieh FY, Lavori PW, Cohen HJ, Feussner JR. 2003. An Overview of Variance Inflation Factors for Sample-Size Calculation. *Evaluation and the Health Professions* 26: 239-257. doi:10.1177/0163278703255230

Snijders TAB. 2005. Power and Sample Size in Multilevel Linear Models. In: Everitt BS, Howell DC (Hrsg.). *Encyclopedia of Statistics in Behavioral Science*. Chichester, UK: John Wiley and Sons, Ltd. doi:10.1002/0470013192.bsa492

Thompson DM, Fernald DH, Mold JW. 2012. Intraclass Correlation Coefficients Typical of Cluster-Randomized Studies: Estimates From the Robert Wood Johnson Prescription for Health Projects. *The Annals of Family Medicine*;10(3):235-40. doi:10.1370/afm.1347

**Examples**

```
# Design effect for two-level model with 30 observations per
# cluster group (level-2 unit) and an assumed intraclass
# correlation coefficient of 0.05.
design_effect(n = 30)

# Design effect for two-level model with 24 observation per cluster
# group and an assumed intraclass correlation coefficient of 0.2.
design_effect(n = 24, icc = 0.2)
```

---

 efc

---

*Sample dataset from the EUROFAMCARE project*


---

**Description**

German data set from the European study on family care of older people.

**References**

Lamura G, Döhner H, Kofahl C, editors. *Family carers of older people in Europe: a six-country comparative study*. Münster: LIT, 2008.

---

find_beta	<i>Determining distribution parameters</i>
-----------	--

---

### Description

find\_beta(), find\_normal() and find\_cauchy() find the shape, mean and standard deviation resp. the location and scale parameters to describe the beta, normal or cauchy distribution, based on two percentiles. find\_beta2() finds the shape parameters for a Beta distribution, based on a probability value and its standard error or confidence intervals.

### Usage

```
find_beta(x1, p1, x2, p2)
```

```
find_beta2(x, se, ci, n)
```

```
find_cauchy(x1, p1, x2, p2)
```

```
find_normal(x1, p1, x2, p2)
```

### Arguments

x1	Value for the first percentile.
p1	Probability of the first percentile.
x2	Value for the second percentile.
p2	Probability of the second percentile.
x	Numeric, a probability value between 0 and 1. Typically indicates a prevalence rate of an outcome of interest; Or an integer value with the number of observed events. In this case, specify n to indicate the total number of observations.
se	The standard error of x. Either se or ci must be specified.
ci	The upper limit of the confidence interval of x. Either se or ci must be specified.
n	Numeric, number of total observations. Needs to be specified, if x is an integer (number of observed events), and no probability. See 'Examples'.

### Details

These functions can be used to find parameter for various distributions, to define prior probabilities for Bayesian analyses. x1, p1, x2 and p2 are parameters that describe two quantiles. Given this knowledge, the distribution parameters are returned.

Use find\_beta2(), if the known parameters are, e.g. a prevalence rate or similar probability, and its standard deviation or confidence interval. In this case. x should be a probability, for example a prevalence rate of a certain event. se then needs to be the standard error for this probability. Alternatively, ci can be specified, which should indicate the upper limit of the confidence interval of the probability (prevalence rate) x. If the number of events out of a total number of trials is known (e.g. 12 heads out of 30 coin tosses), x can also be the number of observed events, while n

indicates the total amount of trials (in the above example, the function call would be: `find_beta2(x = 12, n = 30)`).

### Value

A list of length two, with the two distribution parameters that can be used to define the distribution, which (best) describes the shape for the given input parameters.

### References

Cook JD. Determining distribution parameters from quantiles. 2010: Department of Biostatistics, Texas ([PDF](#))

### Examples

```
# example from blogpost:
# https://www.johndcook.com/blog/2010/01/31/parameters-from-percentiles/
# 10% of patients respond within 30 days of treatment
# and 80% respond within 90 days of treatment
find_normal(x1 = 30, p1 = .1, x2 = 90, p2 = .8)
find_cauchy(x1 = 30, p1 = .1, x2 = 90, p2 = .8)

parms <- find_normal(x1 = 30, p1 = .1, x2 = 90, p2 = .8)
curve(
  dnorm(x, mean = parms$mean, sd = parms$sd),
  from = 0, to = 200
)

parms <- find_cauchy(x1 = 30, p1 = .1, x2 = 90, p2 = .8)
curve(
  dcauchy(x, location = parms$location, scale = parms$scale),
  from = 0, to = 200
)

find_beta2(x = .25, ci = .5)

shapes <- find_beta2(x = .25, ci = .5)
curve(dbeta(x, shapes[[1]], shapes[[2]]))

# find Beta distribution for 3 events out of 20 observations
find_beta2(x = 3, n = 20)

shapes <- find_beta2(x = 3, n = 20)
curve(dbeta(x, shapes[[1]], shapes[[2]]))
```

---

gmd	<i>Gini's Mean Difference</i>
-----	-------------------------------

---

### Description

`gmd()` computes Gini's mean difference for a numeric vector or for all numeric vectors in a data frame.

### Usage

```
gmd(x, select = NULL)
```

### Arguments

<code>x</code>	A vector or data frame.
<code>select</code>	Optional, names of variables as character vector that should be selected for further processing. Required, if <code>x</code> is a data frame (and no vector) and only selected variables from <code>x</code> should be processed.

### Value

For numeric vectors, Gini's mean difference. For non-numeric vectors or vectors of length  $< 2$ , returns NA.

### Note

Gini's mean difference is defined as the mean absolute difference between any two distinct elements of a vector. Missing values from `x` are silently removed.

### References

David HA. Gini's mean difference rediscovered. *Biometrika* 1968(55): 573-575

### Examples

```
data(efc)
gmd(efc$e17age)
gmd(efc, c("e17age", "c160age", "c12hour"))
```

---

`inequ_trend`*Compute trends in status inequalities*

---

### Description

This method computes the proportional change of absolute (rate differences) and relative (rate ratios) inequalities of prevalence rates for two different status groups, as proposed by Mackenbach et al. (2015).

### Usage

```
inequ_trend(data, prev.low, prev.hi)
```

### Arguments

<code>data</code>	A data frame that contains the variables with prevalence rates for both low and high status groups (see 'Examples').
<code>prev.low</code>	The name of the variable with the prevalence rates for the low status groups.
<code>prev.hi</code>	The name of the variable with the prevalence rates for the hi status groups.

### Details

Given the time trend of prevalence rates of an outcome for two status groups (e.g. the mortality rates for people with lower and higher socioeconomic status over 40 years), this function computes the proportional change of absolute and relative inequalities, expressed in changes in rate differences and rate ratios. The function implements the algorithm proposed by *Mackenbach et al. 2015*.

### Value

A data frame with the prevalence rates as well as the values for the proportional change in absolute (rd) and relative (rr) inequalities.

### References

Mackenbach JP, Martikainen P, Menvielle G, de Gelder R. 2015. The Arithmetic of Reducing Relative and Absolute Inequalities in Health: A Theoretical Analysis Illustrated with European Mortality Data. *Journal of Epidemiology and Community Health* 70(7): 730-36. doi:10.1136/jech-2015207018

### Examples

```
# This example reproduces Fig. 1 of Mackenbach et al. 2015, p.5

# 40 simulated time points, with an initial rate ratio of 2 and
# a rate difference of 100 (i.e. low status group starts with a
# prevalence rate of 200, the high status group with 100)

# annual decline of prevalence is 1% for the low, and 3% for the
```

```
# high status group

n <- 40
time <- seq(1, n, by = 1)
lo <- rep(200, times = n)
for (i in 2:n) lo[i] <- lo[i - 1] * .99

hi <- rep(100, times = n)
for (i in 2:n) hi[i] <- hi[i - 1] * .97

prev.data <- data.frame(lo, hi)

# print values
inequ_trend(prev.data, "lo", "hi")

# plot trends - here we see that the relative inequalities
# are increasing over time, while the absolute inequalities
# are first increasing as well, but later are decreasing
# (while rel. inequ. are still increasing)
plot(inequ_trend(prev.data, "lo", "hi"))
```

---

is\_prime

*Find prime numbers*

---

### Description

This functions checks whether a number is, or numbers in a vector are prime numbers.

### Usage

```
is_prime(x)
```

### Arguments

x                    An integer, or a vector of integers.

### Value

TRUE for each prime number in x, FALSE otherwise.

### Examples

```
is_prime(89)
is_prime(15)
is_prime(c(1, 2, 3, 4, 5, 6, 7, 8, 9, 10))
```

---

 kruskal\_wallis\_test *Kruskal-Wallis test*


---

### Description

This function performs a Kruskal-Wallis rank sum test, which is a non-parametric method to test the null hypothesis that the population median of all of the groups are equal. The alternative is that they differ in at least one. Unlike the underlying base R function `kruskal.test()`, this function allows for weighted tests.

### Usage

```
kruskal_wallis_test(data, select = NULL, by = NULL, weights = NULL)
```

### Arguments

data	A data frame.
select	Name(s) of the continuous variable(s) (as character vector) to be used as samples for the test. <code>select</code> can be one of the following: <ul style="list-style-type: none"> <li>• <code>select</code> can be used in combination with <code>by</code>, in which case <code>select</code> is the name of the continuous variable (and <code>by</code> indicates a grouping factor).</li> <li>• <code>select</code> can also be a character vector of length two or more (more than two names only apply to <code>kruskal_wallis_test()</code>), in which case the two continuous variables are treated as samples to be compared. <code>by</code> must be <code>NULL</code> in this case.</li> <li>• If <code>select</code> is of length <b>two</b> and <code>paired = TRUE</code>, the two samples are considered as <i>dependent</i> and a paired test is carried out.</li> <li>• If <code>select</code> specifies <b>one</b> variable and <code>by = NULL</code>, a one-sample test is carried out (only applicable for <code>t_test()</code> and <code>wilcoxon_test()</code>)</li> <li>• For <code>chi_squared_test()</code>, if <code>select</code> specifies <b>one</b> variable and both <code>by</code> and <code>probabilities</code> are <code>NULL</code>, a one-sample test against given probabilities is automatically conducted, with equal probabilities for each level of <code>select</code>.</li> </ul>
by	Name of the variable indicating the groups. Required if <code>select</code> specifies only one variable that contains all samples to be compared in the test. If <code>by</code> is not a factor, it will be coerced to a factor. For <code>chi_squared_test()</code> , if <code>probabilities</code> is provided, <code>by</code> must be <code>NULL</code> .
weights	Name of an (optional) weighting variable to be used for the test.

### Details

The function simply is a wrapper around `kruskal.test()`. The weighted version of the Kruskal-Wallis test is based on the **survey** package, using `survey::svyranktest()`.



**Value**

A data frame with test results.

**Which test to use**

The following table provides an overview of which test to use for different types of data. The choice of test depends on the scale of the outcome variable and the number of samples to compare.

<b>Samples</b>	<b>Scale of Outcome</b>	<b>Significance Test</b>
1	binary / nominal	<code>chi_squared_test()</code>
1	continuous, not normal	<code>wilcoxon_test()</code>
1	continuous, normal	<code>t_test()</code>
2, independent	binary / nominal	<code>chi_squared_test()</code>
2, independent	continuous, not normal	<code>mann_whitney_test()</code>
2, independent	continuous, normal	<code>t_test()</code>
2, dependent	binary (only 2x2)	<code>chi_squared_test(paired=TRUE)</code>
2, dependent	continuous, not normal	<code>wilcoxon_test()</code>
2, dependent	continuous, normal	<code>t_test(paired=TRUE)</code>
>2, independent	continuous, not normal	<code>kruskal_wallis_test()</code>
>2, independent	continuous, normal	<code>datawizard::means_by_group()</code>
>2, dependent	continuous, not normal	<i>not yet implemented</i> (1)
>2, dependent	continuous, normal	<i>not yet implemented</i> (2)

(1) More than two dependent samples are considered as *repeated measurements*. For ordinal or not-normally distributed outcomes, these samples are usually tested using a `friedman.test()`, which requires the samples in one variable, the groups to compare in another variable, and a third variable indicating the repeated measurements (subject IDs).

(2) More than two dependent samples are considered as *repeated measurements*. For normally distributed outcomes, these samples are usually tested using a ANOVA for repeated measurements. A more sophisticated approach would be using a linear mixed model.

**References**

- Bender, R., Lange, S., Ziegler, A. Wichtige Signifikanztests. Dtsch Med Wochenschr 2007; 132: e24–e25
- du Prel, J.B., Röhrig, B., Hommel, G., Blettner, M. Auswahl statistischer Testverfahren. Dtsch Arztebl Int 2010; 107(19): 343–8

**See Also**

- `t_test()` for parametric t-tests of dependent and independent samples.
- `mann_whitney_test()` for non-parametric tests of unpaired (independent) samples.
- `wilcoxon_test()` for Wilcoxon rank sum tests for non-parametric tests of paired (dependent) samples.
- `kruskal_wallis_test()` for non-parametric tests with more than two independent samples.
- `chi_squared_test()` for chi-squared tests (two categorical variables, dependent and independent).

**Examples**

```

data(efc)
# Kruskal-Wallis test for elder's age by education
kruskal_wallis_test(efc, "e17age", by = "c172code")

# when data is in wide-format, specify all relevant continuous
# variables in `select` and omit `by`
set.seed(123)
wide_data <- data.frame(
  scale1 = runif(20),
  scale2 = runif(20),
  scale3 = runif(20)
)
kruskal_wallis_test(wide_data, select = c("scale1", "scale2", "scale3"))

# same as if we had data in long format, with grouping variable
long_data <- data.frame(
  scales = c(wide_data$scale1, wide_data$scale2, wide_data$scale3),
  groups = rep(c("A", "B", "C"), each = 20)
)
kruskal_wallis_test(long_data, select = "scales", by = "groups")
# base R equivalent
kruskal.test(scales ~ groups, data = long_data)

```

---

mann\_whitney\_test

*Mann-Whitney test*


---

**Description**

This function performs a Mann-Whitney test (or Wilcoxon rank sum test for *unpaired* samples). Unlike the underlying base R function `wilcox.test()`, this function allows for weighted tests and automatically calculates effect sizes. For *paired* (dependent) samples, or for one-sample tests, please use the `wilcoxon_test()` function.

A Mann-Whitney test is a non-parametric test for the null hypothesis that two *independent* samples have identical continuous distributions. It can be used for ordinal scales or when the two continuous variables are not normally distributed. For large samples, or approximately normally distributed variables, the `t_test()` function can be used.

**Usage**

```

mann_whitney_test(
  data,
  select = NULL,
  by = NULL,
  weights = NULL,
  mu = 0,
  alternative = "two.sided",
  ...
)

```

**Arguments**

data	A data frame.
select	Name(s) of the continuous variable(s) (as character vector) to be used as samples for the test. <code>select</code> can be one of the following: <ul style="list-style-type: none"> <li>• <code>select</code> can be used in combination with <code>by</code>, in which case <code>select</code> is the name of the continuous variable (and <code>by</code> indicates a grouping factor).</li> <li>• <code>select</code> can also be a character vector of length two or more (more than two names only apply to <code>kruskal_wallis_test()</code>), in which case the two continuous variables are treated as samples to be compared. <code>by</code> must be <code>NULL</code> in this case.</li> <li>• If <code>select</code> is of length <b>two</b> and <code>paired = TRUE</code>, the two samples are considered as <i>dependent</i> and a paired test is carried out.</li> <li>• If <code>select</code> specifies <b>one</b> variable and <code>by = NULL</code>, a one-sample test is carried out (only applicable for <code>t_test()</code> and <code>wilcoxon_test()</code>)</li> <li>• For <code>chi_squared_test()</code>, if <code>select</code> specifies <b>one</b> variable and both <code>by</code> and <code>probabilities</code> are <code>NULL</code>, a one-sample test against given probabilities is automatically conducted, with equal probabilities for each level of <code>select</code>.</li> </ul>
by	Name of the variable indicating the groups. Required if <code>select</code> specifies only one variable that contains all samples to be compared in the test. If <code>by</code> is not a factor, it will be coerced to a factor. For <code>chi_squared_test()</code> , if <code>probabilities</code> is provided, <code>by</code> must be <code>NULL</code> .
weights	Name of an (optional) weighting variable to be used for the test.
mu	The hypothesized difference in means (for <code>t_test()</code> ) or location shift (for <code>wilcoxon_test()</code> and <code>mann_whitney_test()</code> ). The default is 0.
alternative	A character string specifying the alternative hypothesis, must be one of "two.sided" (default), "greater" or "less". See <code>?t.test</code> and <code>?wilcox.test</code> .
...	Additional arguments passed to <code>wilcox.test()</code> (for unweighted tests, i.e. when <code>weights = NULL</code> ).

**Details**

This function is based on `wilcox.test()` and `coin::wilcox_test()` (the latter to extract effect sizes). The weighted version of the test is based on `survey::svyranktest()`.

Interpretation of the effect size `r`, as a rule-of-thumb:

- small effect  $\geq 0.1$
- medium effect  $\geq 0.3$
- large effect  $\geq 0.5$

`r` is calculated as  $r = \frac{|Z|}{\sqrt{n_1+n_2}}$ .

**Value**

A data frame with test results. The function returns `p` and `Z`-values as well as effect size `r` and group-rank-means.

### Which test to use

The following table provides an overview of which test to use for different types of data. The choice of test depends on the scale of the outcome variable and the number of samples to compare.

Samples	Scale of Outcome	Significance Test
1	binary / nominal	chi_squared_test()
1	continuous, not normal	wilcoxon_test()
1	continuous, normal	t_test()
2, independent	binary / nominal	chi_squared_test()
2, independent	continuous, not normal	mann_whitney_test()
2, independent	continuous, normal	t_test()
2, dependent	binary (only 2x2)	chi_squared_test(paired=TRUE)
2, dependent	continuous, not normal	wilcoxon_test()
2, dependent	continuous, normal	t_test(paired=TRUE)
>2, independent	continuous, not normal	kruskal_wallis_test()
>2, independent	continuous, normal	datawizard::means_by_group()
>2, dependent	continuous, not normal	<i>not yet implemented</i> (1)
>2, dependent	continuous, normal	<i>not yet implemented</i> (2)

(1) More than two dependent samples are considered as *repeated measurements*. For ordinal or not-normally distributed outcomes, these samples are usually tested using a `friedman.test()`, which requires the samples in one variable, the groups to compare in another variable, and a third variable indicating the repeated measurements (subject IDs).

(2) More than two dependent samples are considered as *repeated measurements*. For normally distributed outcomes, these samples are usually tested using a ANOVA for repeated measurements. A more sophisticated approach would be using a linear mixed model.

### References

- Ben-Shachar, M.S., Patil, I., Thériault, R., Wiernik, B.M., Lüdtke, D. (2023). Phi, Fei, Fo, Fum: Effect Sizes for Categorical Data That Use the Chi-Squared Statistic. *Mathematics*, 11, 1982. doi:10.3390/math11091982
- Bender, R., Lange, S., Ziegler, A. Wichtige Signifikanztests. *Dtsch Med Wochenschr* 2007; 132: e24–e25
- du Prel, J.B., Röhrig, B., Hommel, G., Blettner, M. Auswahl statistischer Testverfahren. *Dtsch Arztebl Int* 2010; 107(19): 343–8

### See Also

- `t_test()` for parametric t-tests of dependent and independent samples.
- `mann_whitney_test()` for non-parametric tests of unpaired (independent) samples.
- `wilcoxon_test()` for Wilcoxon rank sum tests for non-parametric tests of paired (dependent) samples.
- `kruskal_wallis_test()` for non-parametric tests with more than two independent samples.
- `chi_squared_test()` for chi-squared tests (two categorical variables, dependent and independent).

**Examples**

```

data(efc)
# Mann-Whitney-U tests for elder's age by elder's sex.
mann_whitney_test(efc, "e17age", by = "e16sex")
# base R equivalent
wilcox.test(e17age ~ e16sex, data = efc)

# when data is in wide-format, specify all relevant continuous
# variables in `select` and omit `by`
set.seed(123)
wide_data <- data.frame(scale1 = runif(20), scale2 = runif(20))
mann_whitney_test(wide_data, select = c("scale1", "scale2"))
# base R equivalent
wilcox.test(wide_data$scale1, wide_data$scale2)
# same as if we had data in long format, with grouping variable
long_data <- data.frame(
  scales = c(wide_data$scale1, wide_data$scale2),
  groups = as.factor(rep(c("A", "B"), each = 20))
)
mann_whitney_test(long_data, select = "scales", by = "groups")
# base R equivalent
wilcox.test(scales ~ groups, long_data)

```

---

nhanes\_sample

*Sample dataset from the National Health and Nutrition Examination Survey*


---

**Description**

Selected variables from the National Health and Nutrition Examination Survey that are used in the example from Lumley (2010), Appendix E. See [svyglm.nb](#) for examples.

**References**

Lumley T (2010). *Complex Surveys: a guide to analysis using R*. Wiley

---

prop

*Proportions of values in a vector*


---

**Description**

`prop()` calculates the proportion of a value or category in a variable. `props()` does the same, but allows for multiple logical conditions in one statement. It is similar to `mean()` with logical predicates, however, both `prop()` and `props()` work with grouped data frames.

**Usage**

```
prop(data, ..., weights = NULL, na.rm = TRUE, digits = 4)
```

```
props(data, ..., na.rm = TRUE, digits = 4)
```

**Arguments**

<code>data</code>	A data frame. May also be a grouped data frame (see 'Examples').
<code>...</code>	One or more value pairs of comparisons (logical predicates). Put variable names the left-hand-side and values to match on the right hand side. Expressions may be quoted or unquoted. See 'Examples'.
<code>weights</code>	Vector of weights that will be applied to weight all observations. Must be a vector of same length as the input vector. Default is NULL, so no weights are used.
<code>na.rm</code>	Logical, whether to remove NA values from the vector when the proportion is calculated. <code>na.rm = FALSE</code> gives you the raw percentage of a value in a vector, <code>na.rm = TRUE</code> the valid percentage.
<code>digits</code>	Amount of digits for returned values.

**Details**

`prop()` only allows one logical statement per comparison, while `props()` allows multiple logical statements per comparison. However, `prop()` supports weighting of variables before calculating proportions, and comparisons may also be quoted. Hence, `prop()` also processes comparisons, which are passed as character vector (see 'Examples').

**Value**

For one condition, a numeric value with the proportion of the values inside a vector. For more than one condition, a data frame with one column of conditions and one column with proportions. For grouped data frames, returns a data frame with one column per group with grouping categories, followed by one column with proportions per condition.

**Examples**

```
data(efc)

# proportion of value 1 in e42dep
prop(efc, e42dep == 1)

# expression may also be completely quoted
prop(efc, "e42dep == 1")

# use "props()" for multiple logical statements
props(efc, e17age > 70 & e17age < 80)

# proportion of value 1 in e42dep, and all values greater
# than 2 in e42dep, including missing values. will return a data frame
prop(efc, e42dep == 1, e42dep > 2, na.rm = FALSE)
```

```
# for factors or character vectors, use quoted or unquoted values
library(datawizard)
# convert numeric to factor, using labels as factor levels
efc$e16sex <- to_factor(efc$e16sex)
efc$n4pstu <- to_factor(efc$n4pstu)

# get proportion of female older persons
prop(efc, e16sex == female)

# get proportion of male older persons
prop(efc, e16sex == "male")

# "props()" needs quotes around non-numeric factor levels
props(efc,
  e17age > 70 & e17age < 80,
  n4pstu == 'Care Level 1' | n4pstu == 'Care Level 3'
)
```

---

r2

*Deprecated functions*

---

## **Description**

A list of deprecated functions.

## **Usage**

r2(x)

cohens\_f(x, ...)

eta\_sq(x, ...)

epsilon\_sq(x, ...)

omega\_sq(x, ...)

scale\_weights(x, ...)

robust(x, ...)

icc(x)

p\_value(x, ...)

se(x, ...)

```
means_by_group(x, ...)
```

```
mean_n(x, ...)
```

### Arguments

x	An object.
...	Currently not used.

### Value

Nothing.

---

samplesize_mixed	<i>Sample size for linear mixed models</i>
------------------	--

---

### Description

Compute an approximated sample size for linear mixed models (two-level-designs), based on power-calculation for standard design and adjusted for design effect for 2-level-designs.

### Usage

```
samplesize_mixed(  
  eff.size,  
  df.n = NULL,  
  power = 0.8,  
  sig.level = 0.05,  
  k,  
  n,  
  icc = 0.05  
)
```

```
smpsize_lmm(  
  eff.size,  
  df.n = NULL,  
  power = 0.8,  
  sig.level = 0.05,  
  k,  
  n,  
  icc = 0.05  
)
```



## Arguments

eff.size	Effect size.
df.n	Optional argument for the degrees of freedom for numerator. See 'Details'.
power	Power of test (1 minus Type II error probability).
sig.level	Significance level (Type I error probability).
k	Number of cluster groups (level-2-unit) in multilevel-design.
n	Optional, number of observations per cluster groups (level-2-unit) in multilevel-design.
icc	Expected intraclass correlation coefficient for multilevel-model.

## Details

The sample size calculation is based on a power-calculation for the standard design. If `df.n` is not specified, a power-calculation for an unpaired two-sample t-test will be computed (using `pwr.t.test` of the `pwr`-package). If `df.n` is given, a power-calculation for general linear models will be computed (using `pwr.f2.test` of the `pwr`-package). The sample size of the standard design is then adjusted for the design effect of two-level-designs (see `design_effect`). Thus, the sample size calculation is appropriate in particular for two-level-designs (see *Snijders 2005*). Models that additionally include repeated measures (three-level-designs) may work as well, however, the computed sample size may be less accurate.

## Value

A list with two values: The number of subjects per cluster, and the total sample size for the linear mixed model.

## References

- Cohen J. 1988. Statistical power analysis for the behavioral sciences (2nd ed.). Hillsdale,NJ: Lawrence Erlbaum.
- Hsieh FY, Lavori PW, Cohen HJ, Feussner JR. 2003. An Overview of Variance Inflation Factors for Sample-Size Calculation. *Evaluation and the Health Professions* 26: 239-257.
- Snijders TAB. 2005. Power and Sample Size in Multilevel Linear Models. In: Everitt BS, Howell DC (Hrsg.). *Encyclopedia of Statistics in Behavioral Science*. Chichester, UK: John Wiley and Sons, Ltd.

## Examples

```
# Sample size for multilevel model with 30 cluster groups and a small to
# medium effect size (Cohen's d) of 0.3. 27 subjects per cluster and
# hence a total sample size of about 802 observations is needed.
samplesize_mixed(eff.size = .3, k = 30)

# Sample size for multilevel model with 20 cluster groups and a medium
# to large effect size for linear models of 0.2. Five subjects per cluster and
# hence a total sample size of about 107 observations is needed.
```

```
samplesize_mixed(eff.size = .2, df.n = 5, k = 20, power = .9)
```

---

se_ybar	<i>Standard error of sample mean for mixed models</i>
---------	---

---

### Description

Compute the standard error for the sample mean for mixed models, regarding the extent to which clustering affects the standard errors. May be used as part of the multilevel power calculation for cluster sampling (see *Gelman and Hill 2007, 447ff*).

### Usage

```
se_ybar(fit)
```

### Arguments

fit                      Fitted mixed effects model ([merMod](#)-class).

### Value

The standard error of the sample mean of fit.

### References

Gelman A, Hill J. 2007. Data analysis using regression and multilevel/hierarchical models. Cambridge, New York: Cambridge University Press

### Examples

```
fit <- lmer(Reaction ~ 1 + (1 | Subject), sleepstudy)
se_ybar(fit)
```

---

survey_median	<i>Weighted statistics for variables</i>
---------------	--

---

### Description

`weighted_se()` computes weighted standard errors of a variable or for all variables of a data frame.  
`survey_median()` computes the median for a variable in a survey-design (see `[survey::svydesign()]`).  
`weighted_correlation()` computes a weighted correlation for a two-sided alternative hypothesis.

**Usage**

```

survey_median(x, design)

weighted_correlation(data, ...)

## Default S3 method:
weighted_correlation(data, x, y, weights, ci.lvl = 0.95, ...)

## S3 method for class 'formula'
weighted_correlation(formula, data, ci.lvl = 0.95, ...)

weighted_se(x, weights = NULL)

```

**Arguments**

x	(Numeric) vector or a data frame. For <code>survey_median()</code> or <code>weighted_ttest()</code> , the bare (unquoted) variable name, or a character vector with the variable name.
design	An object of class <code>svydesign</code> , providing a specification of the survey design.
data	A data frame.
...	Currently not used.
y	Optional, bare (unquoted) variable name, or a character vector with the variable name.
weights	Bare (unquoted) variable name, or a character vector with the variable name of the numeric vector of weights. If <code>weights = NULL</code> , unweighted statistic is reported.
ci.lvl	Confidence level of the interval.
formula	A formula of the form <code>lhs ~ rhs1 + rhs2</code> where <code>lhs</code> is a numeric variable giving the data values and <code>rhs1</code> a factor with two levels giving the corresponding groups and <code>rhs2</code> a variable with weights.

**Value**

The weighted (test) statistic.

**Examples**

```

data(efc)
weighted_se(efc$c12hour, abs(runif(n = nrow(efc))))

# survey_median ----
# median for variables from weighted survey designs
data(nhanes_sample)

des <- survey::svydesign(
  id = ~SDMVPSU,
  strat = ~SDMVSTRA,
  weights = ~WTINT2YR,

```

```

    nest = TRUE,
    data = nhanes_sample
  )
  survey_median(total, des)
  survey_median("total", des)

```

---

 svyglm.nb

*Survey-weighted negative binomial generalised linear model*


---

## Description

svyglm.nb() is an extension to the **survey**-package to fit survey-weighted negative binomial models. It uses **svymle** to fit sampling-weighted maximum likelihood estimates, based on starting values provided by **glm.nb**, as proposed by Lumley (2010, pp249).

## Usage

```
svyglm.nb(formula, design, ...)
```

## Arguments

formula	An object of class formula, i.e. a symbolic description of the model to be fitted. See 'Details' in <b>glm</b> .
design	An object of class <b>svydesign</b> , providing a specification of the survey design.
...	Other arguments passed down to <b>glm.nb</b> .

## Details

For details on the computation method, see Lumley (2010), Appendix E (especially 254ff.)

**sjstats** implements following S3-methods for **svyglm.nb**-objects: `family()`, `model.frame()`, `formula()`, `print()`, `predict()` and `residuals()`. However, these functions have some limitations:

- `family()` simply returns the family-object from the underlying **glm.nb**-model.
- The `predict()`-method just re-fits the **svyglm.nb**-model with **glm.nb**, overwrites the `$coefficients` from this model-object with the coefficients from the returned **svymle**-object and finally calls `predict.glm` to compute the predicted values.
- `residuals()` re-fits the **svyglm.nb**-model with **glm.nb** and then computes the Pearson-residuals from the **glm.nb**-object.

## Value

An object of class **svymle** and **svyglm.nb**, with some additional information about the model.

## References

Lumley T (2010). Complex Surveys: a guide to analysis using R. Wiley

**Examples**

```

# -----
# This example reproduces the results from
# Lumley 2010, figure E.7 (Appendix E, p256)
# -----
if (require("survey")) {
  data(nhanes_sample)

  # create survey design
  des <- svydesign(
    id = ~SDMVPSU,
    strat = ~SDMVSTRA,
    weights = ~WTINT2YR,
    nest = TRUE,
    data = nhanes_sample
  )

  # fit negative binomial regression
  fit <- svyglm.nb(total ~ factor(RIAGENDR) * (log(age) + factor(RIDRETH1)), des)

  # print coefficients and standard errors
  fit
}

```

svyglm.zip

*Survey-weighted zero-inflated Poisson model***Description**

svyglm.zip() is an extension to the **survey**-package to fit survey-weighted zero-inflated Poisson models. It uses [svyml](#) to fit sampling-weighted maximum likelihood estimates, based on starting values provided by [zeroinfl](#).

**Usage**

```
svyglm.zip(formula, design, ...)
```

**Arguments**

formula	An object of class formula, i.e. a symbolic description of the model to be fitted. See 'Details' in <a href="#">zeroinfl</a> .
design	An object of class <a href="#">svydesign</a> , providing a specification of the survey design.
...	Other arguments passed down to <a href="#">zeroinfl</a> .

**Details**

Code modified from <https://notstatschat.rbind.io/2015/05/26/zero-inflated-poisson-from-complex-samples/>.

**Value**

An object of class `svymle` and `svyglm.zip`, with some additional information about the model.

**Examples**

```
if (require("survey")) {
  data(nhanes_sample)
  set.seed(123)
  nhanes_sample$malepartners <- rpois(nrow(nhanes_sample), 2)
  nhanes_sample$malepartners[sample(1:2992, 400)] <- 0

  # create survey design
  des <- svydesign(
    id = ~SDMVPSU,
    strat = ~SDMVSTRA,
    weights = ~WTINT2YR,
    nest = TRUE,
    data = nhanes_sample
  )

  # fit negative binomial regression
  fit <- svyglm.zip(
    malepartners ~ age + factor(RIDRETH1) | age + factor(RIDRETH1),
    des
  )

  # print coefficients and standard errors
  fit
}
```

---

table_values	<i>Expected and relative table values</i>
--------------	---

---

**Description**

This function calculates a table's cell, row and column percentages as well as expected values and returns all results as lists of tables.

**Usage**

```
table_values(tab, digits = 2)
```

**Arguments**

tab	Simple <code>table</code> or <code>fTable</code> of which cell, row and column percentages as well as expected values are calculated. Tables of class <code>xTable</code> and other will be coerced to <code>fTable</code> objects.
digits	Amount of digits for the table percentage values.

**Value**

(Invisibly) returns a list with four tables:

1. cell a table with cell percentages of tab
2. row a table with row percentages of tab
3. col a table with column percentages of tab
4. expected a table with expected values of tab

**Examples**

```
tab <- table(sample(1:2, 30, TRUE), sample(1:3, 30, TRUE))
# show expected values
table_values(tab)$expected
# show cell percentages
table_values(tab)$cell
```

---

t\_test

*Student's t test*

---

**Description**

This function performs a Student's t test for two independent samples, for paired samples, or for one sample. It's a parametric test for the null hypothesis that the means of two independent samples are equal, or that the mean of one sample is equal to a specified value. The hypothesis can be one- or two-sided.

Unlike the underlying base R function `t.test()`, this function allows for weighted tests and automatically calculates effect sizes. Cohen's *d* is returned for larger samples ( $n > 20$ ), while Hedges' *g* is returned for smaller samples.

**Usage**

```
t_test(
  data,
  select = NULL,
  by = NULL,
  weights = NULL,
  paired = FALSE,
  mu = 0,
  alternative = "two.sided"
)
```

## Arguments

data	A data frame.
select	Name(s) of the continuous variable(s) (as character vector) to be used as samples for the test. select can be one of the following: <ul style="list-style-type: none"> <li>• select can be used in combination with by, in which case select is the name of the continuous variable (and by indicates a grouping factor).</li> <li>• select can also be a character vector of length two or more (more than two names only apply to <code>kruskal_wallis_test()</code>), in which case the two continuous variables are treated as samples to be compared. by must be NULL in this case.</li> <li>• If select is of length <b>two</b> and <code>paired = TRUE</code>, the two samples are considered as <i>dependent</i> and a paired test is carried out.</li> <li>• If select specifies <b>one</b> variable and <code>by = NULL</code>, a one-sample test is carried out (only applicable for <code>t_test()</code> and <code>wilcoxon_test()</code>)</li> <li>• For <code>chi_squared_test()</code>, if select specifies <b>one</b> variable and both <code>by</code> and <code>probabilities</code> are NULL, a one-sample test against given probabilities is automatically conducted, with equal probabilities for each level of select.</li> </ul>
by	Name of the variable indicating the groups. Required if select specifies only one variable that contains all samples to be compared in the test. If by is not a factor, it will be coerced to a factor. For <code>chi_squared_test()</code> , if probabilities is provided, by must be NULL.
weights	Name of an (optional) weighting variable to be used for the test.
paired	Logical, whether to compute a paired t-test for dependent samples.
mu	The hypothesized difference in means (for <code>t_test()</code> ) or location shift (for <code>wilcoxon_test()</code> and <code>mann_whitney_test()</code> ). The default is 0.
alternative	A character string specifying the alternative hypothesis, must be one of "two.sided" (default), "greater" or "less". See <code>?t.test</code> and <code>?wilcox.test</code> .

## Details

Interpretation of effect sizes are based on rules described in `effectsize::interpret_cohens_d()` and `effectsize::interpret_hedges_g()`. Use these function directly to get other interpretations, by providing the returned effect size (*Cohen's d* or *Hedges's g* in this case) as argument, e.g. `interpret_cohens_d(0.35, rules = "sawilowsky2009")`.

## Value

A data frame with test results. Effectsize Cohen's *d* is returned for larger samples ( $n > 20$ ), while Hedges' *g* is returned for smaller samples.

## Which test to use

The following table provides an overview of which test to use for different types of data. The choice of test depends on the scale of the outcome variable and the number of samples to compare.



Samples	Scale of Outcome	Significance Test
1	binary / nominal	chi_squared_test()
1	continuous, not normal	wilcoxon_test()
1	continuous, normal	t_test()
2, independent	binary / nominal	chi_squared_test()
2, independent	continuous, not normal	mann_whitney_test()
2, independent	continuous, normal	t_test()
2, dependent	binary (only 2x2)	chi_squared_test(paired=TRUE)
2, dependent	continuous, not normal	wilcoxon_test()
2, dependent	continuous, normal	t_test(paired=TRUE)
>2, independent	continuous, not normal	kruskal_wallis_test()
>2, independent	continuous, normal	datawizard::means_by_group()
>2, dependent	continuous, not normal	not yet implemented (1)
>2, dependent	continuous, normal	not yet implemented (2)

(1) More than two dependent samples are considered as *repeated measurements*. For ordinal or not-normally distributed outcomes, these samples are usually tested using a `friedman.test()`, which requires the samples in one variable, the groups to compare in another variable, and a third variable indicating the repeated measurements (subject IDs).

(2) More than two dependent samples are considered as *repeated measurements*. For normally distributed outcomes, these samples are usually tested using a ANOVA for repeated measurements. A more sophisticated approach would be using a linear mixed model.

## References

- Bender, R., Lange, S., Ziegler, A. Wichtige Signifikanztests. Dtsch Med Wochenschr 2007; 132: e24–e25
- du Prel, J.B., Röhrig, B., Hommel, G., Blettner, M. Auswahl statistischer Testverfahren. Dtsch Arztebl Int 2010; 107(19): 343–8

## See Also

- `t_test()` for parametric t-tests of dependent and independent samples.
- `mann_whitney_test()` for non-parametric tests of unpaired (independent) samples.
- `wilcoxon_test()` for Wilcoxon rank sum tests for non-parametric tests of paired (dependent) samples.
- `kruskal_wallis_test()` for non-parametric tests with more than two independent samples.
- `chi_squared_test()` for chi-squared tests (two categorical variables, dependent and independent).

## Examples

```
data(sleep)
# one-sample t-test
t_test(sleep, "extra")
# base R equivalent
```

```
t.test(extra ~ 1, data = sleep)

# two-sample t-test, by group
t_test(mtcars, "mpg", by = "am")
# base R equivalent
t.test(mpg ~ am, data = mtcars)

# paired t-test
t_test(mtcars, c("mpg", "hp"), paired = TRUE)
# base R equivalent
t.test(mtcars$mpg, mtcars$hp, data = mtcars, paired = TRUE)
```

---

var\_pop

*Calculate population variance and standard deviation*

---

## Description

Calculate the population variance or standard deviation of a vector.

## Usage

```
var_pop(x)
```

```
sd_pop(x)
```

## Arguments

x (Numeric) vector.

## Details

Unlike `var`, which returns the sample variance, `var_pop()` returns the population variance. `sd_pop()` returns the standard deviation based on the population variance.

## Value

The population variance or standard deviation of x.

## Examples

```
data(efc)

# sampling variance
var(efc$c12hour, na.rm = TRUE)
# population variance
var_pop(efc$c12hour)

# sampling sd
```

```
sd(efc$c12hour, na.rm = TRUE)
# population sd
sd_pop(efc$c12hour)
```

---

weight

*Weight a variable*


---

### Description

These functions weight the variable `x` by a specific vector of weights.

### Usage

```
weight(x, weights, digits = 0)
weight2(x, weights)
```

### Arguments

<code>x</code>	(Unweighted) variable.
<code>weights</code>	Vector with same length as <code>x</code> , which contains weight factors. Each value of <code>x</code> has a specific assigned weight in <code>weights</code> .
<code>digits</code>	Numeric value indicating the number of decimal places to be used for rounding the weighted values. By default, this value is 0, i.e. the returned values are integer values.

### Details

`weight2()` sums up all `weights` values of the associated categories of `x`, whereas `weight()` uses a `xtabs` formula to weight cases. Thus, `weight()` may return a vector of different length than `x`.

### Value

The weighted `x`.

### Note

The values of the returned vector are in sorted order, whereas the values' order of the original `x` may be spread randomly. Hence, `x` can't be used, for instance, for further cross tabulation. In case you want to have weighted contingency tables or (grouped) box plots etc., use the `weightBy` argument of most functions.

**Examples**

```

v <- sample(1:4, 20, TRUE)
table(v)
w <- abs(rnorm(20))
table(weight(v, w))
table(weight2(v, w))

set.seed(1)
x <- sample(letters[1:5], size = 20, replace = TRUE)
w <- runif(n = 20)

table(x)
table(weight(x, w))

```

---

wilcoxon\_test

*Wilcoxon rank sum test*


---

**Description**

This function performs Wilcoxon rank sum tests for one sample or for two *paired* (dependent) samples. For *unpaired* (independent) samples, please use the `mann_whitney_test()` function.

A Wilcoxon rank sum test is a non-parametric test for the null hypothesis that two samples have identical continuous distributions. The implementation in `wilcoxon_test()` is only used for *paired*, i.e. *dependent* samples. For independent (unpaired) samples, use `mann_whitney_test()`.

`wilcoxon_test()` can be used for ordinal scales or when the continuous variables are not normally distributed. For large samples, or approximately normally distributed variables, the `t_test()` function can be used (with `paired = TRUE`).

**Usage**

```

wilcoxon_test(
  data,
  select = NULL,
  by = NULL,
  weights = NULL,
  mu = 0,
  alternative = "two.sided",
  ...
)

```

**Arguments**

`data` A data frame.

`select` Name(s) of the continuous variable(s) (as character vector) to be used as samples for the test. `select` can be one of the following:

- select can be used in combination with by, in which case select is the name of the continuous variable (and by indicates a grouping factor).
- select can also be a character vector of length two or more (more than two names only apply to kruskal\_wallis\_test()), in which case the two continuous variables are treated as samples to be compared. by must be NULL in this case.
- If select is of length **two** and paired = TRUE, the two samples are considered as *dependent* and a paired test is carried out.
- If select specifies **one** variable and by = NULL, a one-sample test is carried out (only applicable for t\_test() and wilcoxon\_test())
- For chi\_squared\_test(), if select specifies **one** variable and both by and probabilities are NULL, a one-sample test against given probabilities is automatically conducted, with equal probabilities for each level of select.

by	Name of the variable indicating the groups. Required if select specifies only one variable that contains all samples to be compared in the test. If by is not a factor, it will be coerced to a factor. For chi_squared_test(), if probabilities is provided, by must be NULL.
weights	Name of an (optional) weighting variable to be used for the test.
mu	The hypothesized difference in means (for t_test()) or location shift (for wilcoxon_test() and mann_whitney_test()). The default is 0.
alternative	A character string specifying the alternative hypothesis, must be one of "two.sided" (default), "greater" or "less". See ?t.test and ?wilcox.test.
...	Additional arguments passed to wilcox.test() (for unweighted tests, i.e. when weights = NULL).

### Value

A data frame with test results. The function returns p and Z-values as well as effect size r and group-rank-means.

### Which test to use

The following table provides an overview of which test to use for different types of data. The choice of test depends on the scale of the outcome variable and the number of samples to compare.

Samples	Scale of Outcome	Significance Test
1	binary / nominal	chi_squared_test()
1	continuous, not normal	wilcoxon_test()
1	continuous, normal	t_test()
2, independent	binary / nominal	chi_squared_test()
2, independent	continuous, not normal	mann_whitney_test()
2, independent	continuous, normal	t_test()
2, dependent	binary (only 2x2)	chi_squared_test(paired=TRUE)
2, dependent	continuous, not normal	wilcoxon_test()
2, dependent	continuous, normal	t_test(paired=TRUE)
>2, independent	continuous, not normal	kruskal_wallis_test()

>2, independent	continuous, normal	datawizard::means_by_group()
>2, dependent	continuous, not normal	not yet implemented (1)
>2, dependent	continuous, normal	not yet implemented (2)

(1) More than two dependent samples are considered as *repeated measurements*. For ordinal or not-normally distributed outcomes, these samples are usually tested using a `friedman.test()`, which requires the samples in one variable, the groups to compare in another variable, and a third variable indicating the repeated measurements (subject IDs).

(2) More than two dependent samples are considered as *repeated measurements*. For normally distributed outcomes, these samples are usually tested using a ANOVA for repeated measurements. A more sophisticated approach would be using a linear mixed model.

## References

- Bender, R., Lange, S., Ziegler, A. Wichtige Signifikanztests. Dtsch Med Wochenschr 2007; 132: e24–e25
- du Prel, J.B., Röhrig, B., Hommel, G., Blettner, M. Auswahl statistischer Testverfahren. Dtsch Arztebl Int 2010; 107(19): 343–8

## See Also

- `t_test()` for parametric t-tests of dependent and independent samples.
- `mann_whitney_test()` for non-parametric tests of unpaired (independent) samples.
- `wilcoxon_test()` for Wilcoxon rank sum tests for non-parametric tests of paired (dependent) samples.
- `kruskal_wallis_test()` for non-parametric tests with more than two independent samples.
- `chi_squared_test()` for chi-squared tests (two categorical variables, dependent and independent).

## Examples

```
data(mtcars)
# one-sample test
wilcoxon_test(mtcars, "mpg")
# base R equivalent, we set exact = FALSE to avoid a warning
wilcox.test(mtcars$mpg ~ 1, exact = FALSE)

# paired test
wilcoxon_test(mtcars, c("mpg", "hp"))
# base R equivalent, we set exact = FALSE to avoid a warning
wilcox.test(mtcars$mpg, mtcars$hp, paired = TRUE, exact = FALSE)

# when `by` is specified, each group must be of same length
data(iris)
d <- iris[iris$Species != "setosa", ]
wilcoxon_test(d, "Sepal.Width", by = "Species")
```

# Index

- \* **data**
  - efc, 18
  - nhanes\_sample, 29
- anova, 3
- anova\_stats, 2
- auto\_prior, 3
  
- boot\_ci, 6, 6
- boot\_est (boot\_ci), 6
- boot\_p (boot\_ci), 6
- boot\_se (boot\_ci), 6
- bootstrap, 5
  
- chi\_squared\_test, 10
- chi\_squared\_test(), 12, 25, 28, 41, 46
- chisq.test, 9
- chisq.test(), 11, 14
- chisq\_gof, 8
- cohens\_f (r2), 31
- coin::wilcox\_test(), 27
- cor.test(), 14
- cramer (cramers\_v), 13
- cramers\_v, 13
- crosstable\_statistics (cramers\_v), 13
- crossv\_kfold, 17
- cv, 16
- cv\_compare (cv\_error), 16
- cv\_error, 16
  
- design\_effect, 17, 33
  
- efc, 18
- effectsize::interpret\_cohens\_d(), 40
- effectsize::interpret\_cramers\_v(), 11
- effectsize::interpret\_fei(), 11
- effectsize::interpret\_hedges\_g(), 40
- effectsize::interpret\_phi(), 11
- epsilon\_sq (r2), 31
- eta\_sq (r2), 31
  
- find\_beta, 19
- find\_beta2 (find\_beta), 19
- find\_cauchy (find\_beta), 19
- find\_normal (find\_beta), 19
- fisher.test(), 11, 14
- friedman.test(), 12, 25, 28, 41, 46
- ftable, 38
- ftable(), 14
  
- glm, 36
- glm.nb, 36
- gmd, 21
  
- icc (r2), 31
- inequ\_trend, 22
- is\_prime, 23
  
- kruskal.test(), 24
- kruskal\_wallis\_test, 24
- kruskal\_wallis\_test(), 12, 25, 28, 41, 46
  
- mann\_whitney\_test, 26
- mann\_whitney\_test(), 12, 25, 28, 41, 46
- mcnemar.test(), 11
- mean\_n (r2), 31
- means\_by\_group (r2), 31
- merMod, 34
  
- nhanes\_sample, 29
  
- omega\_sq (r2), 31
  
- p\_value (r2), 31
- phi (cramers\_v), 13
- phi(), 14
- predict.glm, 36
- prop, 29
- props (prop), 29
- pwr.f2.test, 33
- pwr.t.test, 33

r2, 31  
rmse, 17  
robust(r2), 31

samplesize\_mixed, 32  
scale\_weights(r2), 31  
sd\_pop(var\_pop), 42  
se(r2), 31  
se\_ybar, 34  
set\_prior, 4  
smplsize\_lmm(samplesize\_mixed), 32  
survey::svyranktest(), 24, 27  
survey\_median, 34  
svydesign, 35–37  
svyglm.nb, 29, 36  
svyglm.zip, 37  
svymle, 36–38

t\_test, 39  
t\_test(), 12, 25, 28, 41, 46  
table, 38  
table(), 14  
table\_values, 38

var, 42  
var\_pop, 42

weight, 43  
weight2(weight), 43  
weighted\_correlation(survey\_median), 34  
weighted\_se(survey\_median), 34  
wilcox.test(), 27  
wilcoxon\_test, 44  
wilcoxon\_test(), 12, 25, 28, 41, 46

xtab\_statistics(cramers\_v), 13  
xtabs, 38, 43  
xtabs(), 11, 14

zeroinfl, 37